Likelihood-based estimation, model selection, and forecasting of integer-valued trawl processes

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This is joint work with



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Introduction

- Time series of counts appear in various applications: Medical science, epidemiology, meteorology, network modelling, actuarial science, econometrics and finance.
- Count data: Non-negative and integer-valued, and often over-dispersed (i.e. variance > mean).
- Recently the class of (integer-valued) trawl (IVT) processes has been introduced as a flexible model, see Barndorff-Nielsen et al. (2014) for the univariate and Veraart (2019) for the multivariate case.

Aim of the project

- Improve the estimation method for IVT processes (likelihood-based rather than moment-based);
- Tailor model selection tools to the IVT class;
- (Probabilistic) forecasting of IVT processes;

A very short and incomplete review of the literature

Recent surveys & some new developments:

Cameron & Trivedi (1998), Kedem & Fokianos (2002), Cui & Lund (2009); Davis et al. (1999); Davis & Wu (2009); Jung & Tremayne (2011); McKenzie (2003); Weiß (2008), Karlis (2016), Fokianos (2016).

- Literature on count data is spread across different disciplines.
- Overall, two predominant modelling approaches:
 - Discrete autoregressive moving-average (DARMA) models introduced by Jacobs & Lewis (1978a,b).
 - Models obtained from thinning operations going back to the influential work of Steutel & van Harn (1979), e.g. INAR(MA), see e.g. Pedeli et al. (2015).
- Further models: Regression type models (typically based on generalised linear models, see e.g. Fokianos (2016)), also Fokianos et al. (2020); state-space and Bayesian approaches.

Our approach:

- Use "trawling" for modelling counts.
- This is a continuous-time framework based on the idea of "thinning" points.

Introduction What is trawling...? A first "definition"

"Trawling is a method of fishing that involves pulling a fishing net through the water behind one or more boats. The net that is used for trawling is called a trawl." (Wikipedia)



Theoretical framework Definition of trawl process

We define a stationary integer-valued trawl (IVT) process $(X_t)_{t\geq 0}$ by

$$X_t = L(A_t) = \int_{\mathbb{R}\times\mathbb{R}} I_A(x, s-t) L(dx, ds).$$

L is the integer-valued, homogeneous Lévy basis on [0, 1] × R:

- ➡ $L(dx, ds) := \int_{-\infty}^{\infty} y N(dy, dx, ds), \quad (x, s) \in [0, 1] \times \mathbb{R}.$
- N is a homogeneous Poisson random measure on Z × [0, 1] × ℝ with compensator η ⊗ Leb ⊗ Leb, i.e. E(N(dy, dx, ds)) = η(dy)dxds, where η is a Lévy measure satisfying ∫[∞]_{-∞} min(1, |y|)η(dy) < ∞.</p>
- ▶ A Borel set $A_t = A + (0, t)$ with $A = A_0 \subseteq [0, 1] \times (-\infty, 0]$ and Leb $(A) < \infty$ is called the trawl.
 - Typically, we choose A to be of the form

 $A = \{(x, s) : s \le 0, 0 \le x \le d(s)\},\$

where $d: (-\infty, 0] \mapsto [0, 1]$ is continuous and $Leb(A) < \infty$. London

Example Poisson-Exponential trawl



Example Negative binomial-Exponential trawl



Some key properties of IVT processes Cumulants

- > The IVT process is stationary and infinitely divisible.
- > The IVT process is mixing \Rightarrow weakly mixing \Rightarrow ergodic.
- ► The cumulant (log-characteristic) function of a trawl process is, for $\theta \in \mathbb{R}$, given by

$$C_{X_t}(\theta) = C_{L(A_t)}(\theta) = \operatorname{Leb}(A)C_{L'}(\theta),$$

where the random variable L' (called the Lévy seed) associated with L satisfies

 $\mathbb{E}[\exp(i\theta L')] = \exp(\mathcal{C}_{L'}(\theta)), \quad \text{with} \quad \mathcal{C}_{L'}(\theta) = \int \left(e^{i\theta y} - 1\right) \eta(dy).$

I.e. to any infinitely divisible integer-valued law π , say, there exists a stationary integer-valued trawl process having π as its one-dimensional marginal law.

The autocorrelation function is given by

$$\rho(h) := \operatorname{Cor}(Y_t, Y_{t+h}) = \frac{\operatorname{Leb}(A \cap A_h)}{\operatorname{Leb}(A)}, \quad \text{for } h > 0.$$

Examples Modelling the marginal distribution

Example 1 (Poissonian Lévy seed)

Let $L' \sim \text{Poisson}(\nu)$. Then $X_t \sim \text{Poisson}(\nu \text{Leb}(A))$, i.e., for all $t \ge 0$,

 $P(X_t = k) = (\nu \text{Leb}(A))^k e^{-\nu \text{Leb}(A)} / k!, \quad k = 0, 1, 2, \dots$

Example 2 (Negative Binomial Lévy seed)

Let $L' \sim NB(m, p)$ for $m > 0, p \in [0, 1]$. Then $X_t \sim NB(mLeb(A), p)$, i.e., for all $t \ge 0$,

$$\mathsf{P}(X_t = k) = \frac{\Gamma(\operatorname{Leb}(A)m + k)}{k!\Gamma(\operatorname{Leb}(A)m)} (1 - p)^{\operatorname{Leb}(A)m} p^k, \quad k = 0, 1, 2, \dots,$$

where $\Gamma(z) = \int_0^\infty x^{z-1} e^{-x} dx$ for z > 0 is the Γ -function.

Recall the typical choice for the trawl:

 $A = A_0 = \{(x, s) : s \le 0, 0 \le x \le d(s)\}, \qquad A_t = A + (0, t).$

> Restrict attention to a class of *superposition trawls*:

$$oldsymbol{d}(oldsymbol{s}):=\int_0^\infty oldsymbol{e}^{\lambdaoldsymbol{s}}\pi(oldsymbol{d}\lambda),\quad oldsymbol{s}\leq 0,$$

where π is a probability measure on \mathbb{R}_+ .

For $h \ge 0$, the acf is given by

$$\rho(h) := \operatorname{Cor}(L(A_{t+h}), L(A_t)) = \frac{\operatorname{Leb}(A_h \cap A)}{\operatorname{Leb}(A)} = \frac{\int_h^\infty d(-s)ds}{\int_0^\infty d(-s)ds}$$

Examples Modelling the trawl function/correlation structure

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► Exponential trawl function: Let $\lambda > 0$ and $\pi(dx) = \delta_{\lambda}(dx)$, then $d(s) = e^{\lambda s}$ for $s \le 0$ and

 $\rho(h) = \operatorname{Cor}(X_{t+h}, X_t) = \exp(-\lambda h), \quad h \ge 0.$

Inverse Gaussian trawl function: Letting π be given by the inverse Gaussian distribution

$$\pi(dx) = \frac{(\gamma/\delta)^{1/2}}{2K_{1/2}(\delta\gamma)} x^{-1/2} \exp\left(-\frac{1}{2}(\delta^2 x^{-1} + \gamma^2 x)\right) dx,$$

where $K_{\nu}(\cdot)$ is the modified Bessel function of the third kind and $\gamma, \delta \ge 0$ with both not zero simultaneously. Then

$$d(s) = \left(1 - \frac{2s}{\gamma^2}\right)^{-1/2} \exp\left(\delta\gamma\left(1 - \sqrt{1 - \frac{2s}{\gamma^2}}\right)\right), \quad s \le 0,$$
$$\rho(h) = \operatorname{Cor}(X_{t+h}, X_t) = \exp\left(\delta\gamma(1 - \sqrt{1 + 2h/\gamma^2})\right), \quad h \ge 0.$$
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Solution Gamma trawl function: Let π have the $\Gamma(1 + H, \alpha)$ density,

$$\pi(dx) = \frac{1}{\Gamma(1+H)} \alpha^{1+H} \lambda^{H} e^{-\lambda \alpha} dx,$$

where $\alpha > 0$ and H > 0.

$$d(s) = \left(1 - \frac{s}{\alpha}\right)^{-(H+1)}, \quad s \leq 0,$$

and

$$\rho(h) = \operatorname{Cor}(X_{t+h}, X_t) = \frac{\operatorname{Leb}(A_h \cap A)}{\operatorname{Leb}(A)} = \left(1 + \frac{h}{\alpha}\right)^{-H}.$$

Note that in this case

$$\int_0^\infty \rho(h) dh = \begin{cases} \infty & \text{if } H \in (0, 1], \\ \frac{\alpha}{H-1} & \text{if } H > 1, \end{cases}$$

i.e. the trawl process has long memory for $H \in (0, 1]$.

Estimation From method of moments to composite likelihood

Suppose we have $n \in \mathbb{N}$ observations of the IVT process X, x_1, \ldots, x_n , on an equidistant grid of size $\Delta = T/n$.

Define

$$CL^{(h)}(\theta; x) := \prod_{i=1}^{n-h} f(x_{i+h}, x_i; \theta), \quad h \ge 1.$$



► Construct the composite likelihood function, for $\mathcal{H} \subseteq \{1, 2, ..., n-1\}$,

$$\mathcal{L}_{CL}^{\mathcal{H}}(\theta; \mathbf{x}) := \prod_{h \in \mathcal{H}} C \mathcal{L}^{(h)}(\theta; \mathbf{x}) = \prod_{h \in \mathcal{H}} \prod_{i=1}^{n-h} f(\mathbf{x}_{i+h}, \mathbf{x}_i; \theta).$$

> The maximum composite likelihood (MCL) estimator of θ is defined as

$$\hat{\theta}^{\textit{CL}} := \arg \max_{\theta \in \Theta} I^{\mathcal{H}}_{\textit{CL}}(\theta; \mathbf{X}),$$

 $\begin{array}{l} \text{Imperial College} & \int_{CL}^{\mathcal{H}}(\theta; x) := \log L_{CL}^{\mathcal{H}}(\theta; x) \text{ is the log composite likelihood function.} \\ \text{London} & \begin{array}{c} 14/25 \end{array} \end{array}$

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> The joint probability mass function of two observations x_{i+h} and x_i is

$$\begin{aligned} (x_{i+h}, x_i; \theta) &:= \mathcal{P}_{\theta} \left(X_{(i+h)\Delta} = x_{i+h}, X_{i\Delta} = x_i \right) \\ &= \sum_{c=-\infty}^{\infty} \mathcal{P}_{\theta} \left(L(\mathcal{A}_{(i+h)\Delta} \setminus \mathcal{A}_{i\Delta}) = x_{i+h} - c \right) \\ &\cdot \mathcal{P}_{\theta} \left(L(\mathcal{A}_{i\Delta} \setminus \mathcal{A}_{(i+h)\Delta}) = x_i - c \right) \\ &\cdot \mathcal{P}_{\theta} \left(L(\mathcal{A}_{(i+h)\Delta} \cap \mathcal{A}_{i\Delta}) = c \right). \end{aligned}$$

> Suppose the Lévy basis L is positive, i.e. $\eta(y) = 0$ for y < 0. Then we can replace $\sum_{c=-\infty}^{\infty}$ by $\sum_{c=0}^{\min\{x_{i+h}, x_i\}}$ in the above formula.

▶ Let $t, s \ge 0$, choose $C \in \mathbb{N}$ and let $c^{(j)} \sim L(A_t \cap A_s), j = 1, 2, ..., C$, be an iid sample. A simulation based unbiased estimator of $f(x_t, x_s; \theta)$ is

$$\hat{f}(x_t, x_s; \theta) = \frac{1}{C} \sum_{j=1}^{C} P_{\theta}(L(A_t \setminus A_s) = x_t - c^{(j)}) P_{\theta}(L(A_s \setminus A_t) = x_s - c^{(j)}).$$
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MCL outperforms GMM for IVTs



Model selection for IVTs

 Following Takeuchi (1976), Varin & Vidoni (2005) we apply the composite likelihood information criterion (CLAIC)

$$CLAIC = I_{LC}(\hat{\theta}^{CL}; x) + \operatorname{tr}\left\{\hat{V}(\hat{\theta}^{CL})\hat{H}(\hat{\theta}^{CL})^{-1}\right\}$$

as a basis for model selection.

- ► Note that $G(\theta)^{-1} = H(\theta)^{-1} V(\theta) H(\theta)^{-1}$ is the asymptotic covariance matrix of the MCE. We use the straight-forward estimator $\hat{H}(\hat{\theta}^{CL}) = -n^{-1} \frac{\partial}{\partial \theta \partial \theta'} I_{CL}(\hat{\theta}^{CL}; x)$ which is consistent for $H(\theta)$ due to the stationarity and ergodicity of the IVT process, and estimate $\hat{V}(\hat{\theta}^{CL})$ by parametric bootstrap.
- We also apply the BIC adapated to the composite likelihood case, see Gao & Song (2010)

$$CLBIC = I_{CL}(\hat{\theta}^{CL}; x) + \frac{\log(n)}{2} \operatorname{tr} \left\{ \hat{V}(\hat{\theta}^{CL}) \hat{H}(\hat{\theta}^{CL})^{-1} \right\},$$

Imperial College *n* is the number of observations of the data series *x*.

Simulation study of model selection procedure



The numbers plotted are average selection rates of the models given on the *x*-axis, using a given criteria over M = 100 Monte Carlo simulations. For each Monte Carlo replication, n = 4000 observations of the true DGP are simulated on a grid with step size $\Delta = 0.1$.

Probabilistic forecasting of IVTs

► Let $\mathcal{F}_t = \sigma((X_s)_{s \le t})$, let h > 0 be a forecast horizon.

- > Goal: Forecast the future value X_{t+h} (and its distribution).
- > Note that $\tilde{X}_{t+h|t} = \mathbb{E}[X_{t+h}|\mathcal{F}_t]$ is not data coherent.
- > Consider instead a probabilistic forecasting approach, where the interest is in the distribution of $X_{t+h}|\mathcal{F}_t$ and generate data coherent point forecasts, e.g. using the median or mode of the distribution.
- However, since the IVT process X_t is in general non-Markovian, the distribution of X_{t+h}|F_t is highly intractable.
- > The probabilistic forecast of $X_{t+h}|X_t$ gives promising results.

Probabilistic forecasting of IVTs

Proposition 1

Suppose the Lévy basis L is positive, i.e. $\eta(y) = 0$ for y < 0. Now

$$P(X_{t+h} = x_{t+h} | X_t = x_t) = \sum_{c=0}^{\min(x_t, x_{t+h})} P(L(A_{t+h} \setminus A_t) = x_{t+h} - c)$$
$$\cdot P(L(A_t \cap A_{t+h}) = c | X_t = x_t),$$

where

$$P(L(A_t \cap A_{t+h}) = c | X_t = x_t) = \frac{P(L(A_t \setminus A_{t+h}) = x_t - c)P(L(A_t \cap A_{t+h}) = c)}{P(X_t = x_t)}.$$

Empirical study: Goal: Forecasting the bid-ask spread

- Study of high frequency data of bid-ask spreads of equity prices.
- Spread data of Agilent Technologies Inc. stock (ticker: A, NYSE) (measured in U.S. dollar cents); single day, May 4, 2020 [used data from 10:30AM to 4PM, i.e. discarded the first 60 minutes]
- The data were cleaned using the approach proposed in Barndorff-Nielsen et al. (2009), then sampled equidistantly (5s) using the previous-tick method resulting in n = 3 961 observations.
- Let s_t be the spread level at time t. Since the minimum spread level in the data is one tick (i.e. one cent), we work on $x_t = s_t 1$.
- Model selection: The NB-Gamma model is the preferred model on all three criteria, while the second-best model is the NB-IG model. Note, however, that these two models appear to provide an almost identical fit.

Empirical study: Model selection



Empirical study: Forecasting the bid-ask spread



Summary

- Integer-valued trawl processes provide a continuous-time framework for modelling stationary, serially correlated count data.
- They consist of two key components:
 - Integer-valued, homogeneous Lévy basis: Generates random point pattern and determines marginal distribution.
 - **Trawl**: Thins the point pattern and determines the autocorrelation structure.
- We showed that the pairwise likelihood for IVTs is tractable and MCL outperforms previously used GMM.
- The pairwise likelihood can be used for model selection criteria (CLBIC slightly preferred)
- Method for probabilistic forecasting of IVTs using the pairwise likelihood principle.
- Application to forecasting equity spread data: Superior performance of IVT compared to INAR(1) benchmark model.

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