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# Sparse multiple testing: can one estimate the null distribution?

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Etienne Roquain<sup>1</sup>

*Joint work with A. Carpentier<sup>2</sup>, S. Delattre<sup>3</sup>, N. Verzelen<sup>4</sup>,*

<sup>1</sup>LPSM, Sorbonne Université, France

<sup>2</sup>Otto-von-Guericke-Universität Magdeburg, Allemagne

<sup>3</sup>LPSM, Université de Paris, France

<sup>4</sup>INRAE, Montpellier, France

MMMS2 Luminy, 02/06/2020



Arxiv 1912.03109. "On using empirical null distributions in Benjamini-Hochberg procedure"



To appear in AoS. "Estimating minimum effect with outlier selection "

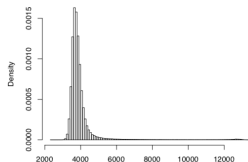
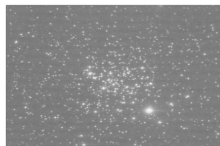
ANR "Sanssouci", ANR "BASICS", GDR ISIS "TASTY"

- 1 Introduction
- 2 Upper bound
- 3 Lower bound
- 4 Additional results
- 5 One-sided alternatives

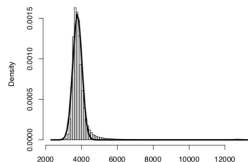
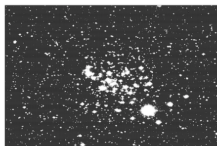
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M67 photography, Package `photutils`

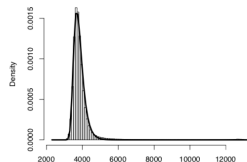
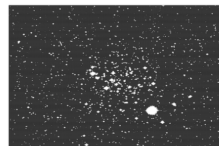
Original



Gaussian fitting



Gumbel fitting

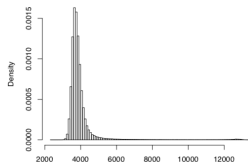
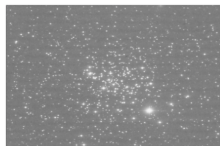


- ▶ Naive null distribution fitting
- ▶ Impact on the risk?

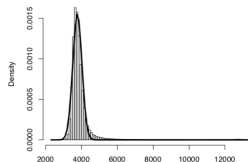
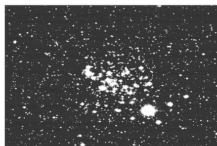
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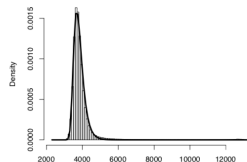
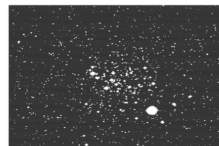
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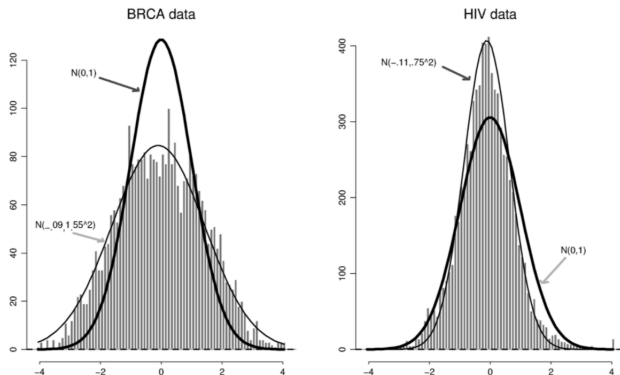
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## Motivation 2: null distribution wrong

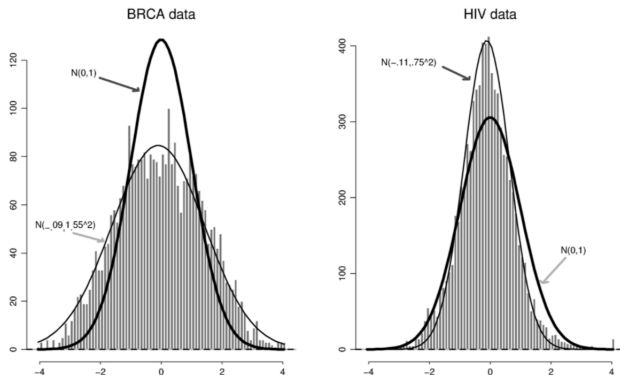
Figure 4 in [Efron (2008)]



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# Existing work (selection)

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## Estimation of the null:

- ▶ Series of work [Efron (2004,2007,2008,2009)]
- ▶ Minimax rate with Fourier analysis: [Jin and Cai (2007)]; [Cai and Jin (2010)]
- ▶ Two group mixture model: [Efron et al. (2001)]; [Sun and Cai (2009)]; [Cai and Sun (2009)]; [Padilla and Bickel (2012)]; [Nguyen and Matias (2014)]; [Heller and Yekutieli (2014)]; [Zablocki et al. (2017)]; [Amar et al. (2017)]; [Cai et al. (2019)]; [Rebafka et al. (2019)]
- ▶ Estimation in factor model: [Efron (2007a)]; [Leek and Storey (2008)]; [Friguet et al. (2009)]; [Fan et al. (2012)]; [Fan and Han (2017)]

## Impact on the risk:

- ▶ FDR control in symmetric, centered, one-sided case: [Barber and Candès (2015)]; [Arias-Castro and Chen (2017)]

## Lower bounds in multiple testing:

- ▶ [Arias-Castro and Chen (2017)]; [Rabinovich et al. (2017)]; [Castillo and R. (2020).]

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# Setting

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## Observations

$$Y = (Y_i)_{1 \leq i \leq n} \text{ indep}, \quad Y_i \sim P_i, \quad \text{parameter } P = (P_i)_{1 \leq i \leq n} \in \mathcal{P}$$

## Gaussian null assumption:

Most of the  $P_i$ 's equal  $\mathcal{N}(\theta, \sigma^2)$ , for some unknown  $\theta, \sigma$

## Example:

$$P = (P_1, \mathcal{N}(\theta, \sigma^2), P_3, \mathcal{N}(\theta, \sigma^2), \mathcal{N}(\theta, \sigma^2), \mathcal{N}(\theta, \sigma^2), P_7, \mathcal{N}(\theta, \sigma^2))$$

- ▶ Ensures  $\theta = \theta(P)$  and  $\sigma = \sigma(P)$  uniquely defined
- ▶ Test  $H_{0,i} : "P_i = \mathcal{N}(\theta(P), \sigma^2(P))"$  against  $H_{1,i} : "P_i \neq \mathcal{N}(\theta(P), \sigma^2(P))"$   
"item  $i$  comes from the background"      "item  $i$  comes from signal"

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# Criteria

- ▶ True null set  $\mathcal{H}_0(P) = \{i : P \text{ satisfies } H_{0,i}\}$ ,  $n_0(P) = |\mathcal{H}_0(P)|$
- ▶ False null set  $\mathcal{H}_1(P) = \mathcal{H}_0(P)^c$ ,  $n_1(P) = |\mathcal{H}_1(P)|$
- ▶ for a procedure  $R(Y) \subset \{1, \dots, n\}$

$$\text{FDP}(P, R(Y)) = \frac{|R(Y) \cap \mathcal{H}_0(P)|}{|R(Y)| \vee 1} \quad \text{'false discovery proportion'}$$

$$\mathbf{E}_P[\text{FDP}(P, R(Y))] = \text{FDR}(P, R) \quad \text{'false discovery rate'}$$

$$\text{TDP}(P, R(Y)) = \frac{|R(Y) \cap \mathcal{H}_1(P)|}{n_1(P) \vee 1} \quad \text{'true discovery proportion'}$$

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$$n_1(P) \leq k_n \text{ with } k_n \text{ 'small'}$$

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# Oracle procedure $BH_{\alpha}^*$

► Rescaled observation  $Z_i = (Y_i - \theta(P))/\sigma(P)$

► Apply the standard BH procedure to the  $Z_i$ 's:

- Sorting  $|Z|_{(1)} \geq |Z|_{(2)} \geq \dots \geq |Z|_{(n)}$

- Quantiles

$$t_k = \bar{\Phi}^{-1}(\alpha k / (2n))$$

- Rejection number

$$\hat{k} = \max\{k : |Z|_{(k)} \geq t_k\}$$

- Select the  $Z_i$ 's corresponding to  $|Z|_{(1)}, |Z|_{(2)}, \dots, |Z|_{(\hat{k})}$ .

Theorem [Benjamini and Hochberg (1995), Benjamini and Yekutieli (2001)]

$$\forall P \in \mathcal{P}, \quad \text{FDR}(P, BH_{\alpha}^*) = \alpha n_0(P)/n \quad \simeq \alpha \text{ under sparsity}$$

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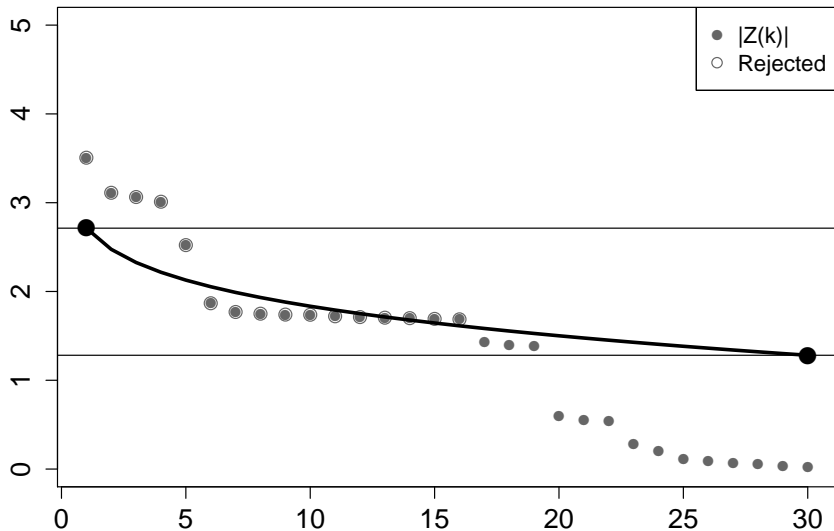
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# Optimality under a sparsity range

Procedure  $R$  optimal:  $R \approx \text{BH}_\alpha^*$  **both** for FDP and TDP

## Definition

Procedure  $R$  optimal for a sparsity  $k_n$ : there exists  $\eta_n \rightarrow 0$ , s.t.

$$(I) \limsup_n \sup_{\alpha \in (1/n, 1/2)} \left\{ \sup_{\substack{P \in \mathcal{P} \\ n_1(P) \leq k_n}} \{ \text{FDR}(P, R) \} - \alpha \right\} \leq 0$$

$$(II) \lim_n \sup_{\alpha \in (1/n, 1/2)} \left\{ \sup_{\substack{P \in \mathcal{P} \\ n_1(P) \leq k_n}} \left\{ \mathbf{P}_{Y \sim P} \left( \text{TDP}(P, R) < \text{TDP}(P, \text{BH}_{\alpha(1-\eta_n)}^*) \right) \right\} \right\} = 0$$

► Robust criteria: alternatives arbitrary

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1 Introduction

**2 Upper bound**

3 Lower bound

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# Plugged BH procedure

- ▶ Estimation: robust minimax estimator of  $\theta(P)$ ,  $\sigma(P)$ :

$$\tilde{\theta} = Y_{(\lceil n/2 \rceil)}; \quad \tilde{\sigma} = U_{(\lceil n/2 \rceil)} / \bar{\Phi}^{-1}(1/4), \quad U_i = |Y_i - Y_{(\lceil n/2 \rceil)}|$$

of  $L^1$  max risk  $\asymp (k_n/n) \vee n^{-1/2}$  for sparsity  $k_n$  [Huber, 1964], [Chen et al. (2018)]

- ▶ Plugged BH procedure  $BH(\tilde{\theta}, \tilde{\sigma})$

- Rescaled observation  $Z'_i = (Y_i - \tilde{\theta}) / \tilde{\sigma}$
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# Plugged BH procedure

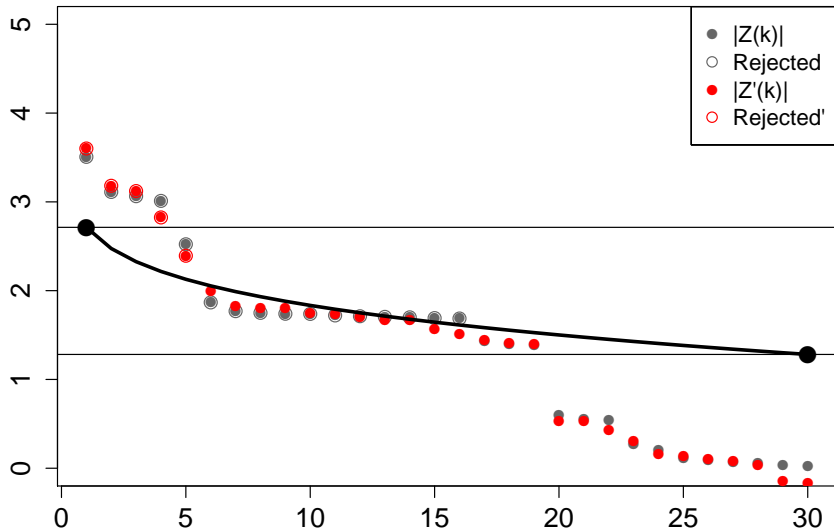
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  - Rescaled observation  $Z_i' = (Y_i - \tilde{\theta}) / \tilde{\sigma}$
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# Plugged BH procedure



# Upper bound

Heuristic:  $\text{BH}_\alpha(\tilde{\theta}, \tilde{\sigma}) \approx \text{BH}_\alpha^*$  if

$$|\tilde{\theta} - \theta(P)| \ll \min_k \left\{ \bar{\Phi}^{-1}(\alpha k/n) - \bar{\Phi}^{-1}(\alpha(k+1)/n) \right\} \approx 1/\sqrt{\log n}$$

$$|\tilde{\sigma} - \sigma(P)| \ll \min_k \left\{ (\bar{\Phi}^{-1}(\alpha k/n) - \bar{\Phi}^{-1}(\alpha(k+1)/n)) / \bar{\Phi}^{-1}(\alpha k/n) \right\} \approx 1/\log n$$

Suggest  $\text{BH}_\alpha(\tilde{\theta}, \tilde{\sigma}) \approx \text{BH}_\alpha^*$  for  $k_n/n \ll 1/\log(n)$ .

Proposition 1 [R. and Verzelen (2020)]

$\text{BH}_\alpha(\tilde{\theta}, \tilde{\sigma})$  is optimal for any sparsity sequence  $k_n \ll n/\log(n)$ .

Proof: rescaling of  $p$ -value process; combining BH procedure and  $(\tilde{\theta}, \tilde{\sigma})$  leave-one-out properties

$$\{p_i(\tilde{\theta}, \tilde{\sigma}) \leq T_\alpha(Y; \tilde{\theta}, \tilde{\sigma})\} \subset \{p_i(\tilde{\theta}^{(i)}, \tilde{\sigma}^{(i)}) \leq T_\alpha(Y^{(i)}; \tilde{\theta}^{(i)}, \tilde{\sigma}^{(i)})\}.$$

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# Upper bound

Heuristic:  $\text{BH}_\alpha(\tilde{\theta}, \tilde{\sigma}) \approx \text{BH}_\alpha^*$  if

$$|\tilde{\theta} - \theta(P)| \ll \min_k \left\{ \bar{\Phi}^{-1}(\alpha k/n) - \bar{\Phi}^{-1}(\alpha(k+1)/n) \right\} \approx 1/\sqrt{\log n}$$

$$|\tilde{\sigma} - \sigma(P)| \ll \min_k \left\{ (\bar{\Phi}^{-1}(\alpha k/n) - \bar{\Phi}^{-1}(\alpha(k+1)/n)) / \bar{\Phi}^{-1}(\alpha k/n) \right\} \approx 1/\log n$$

Suggest  $\text{BH}_\alpha(\tilde{\theta}, \tilde{\sigma}) \approx \text{BH}_\alpha^*$  for  $k_n/n \ll 1/\log(n)$ .

## Proposition 1 [R. and Verzelen (2020)]

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1 Introduction

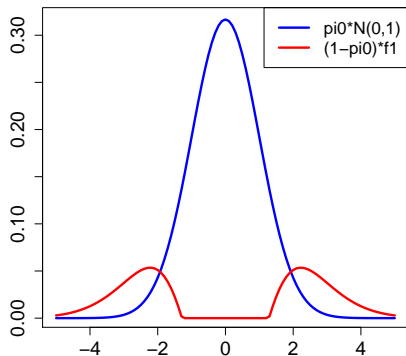
2 Upper bound

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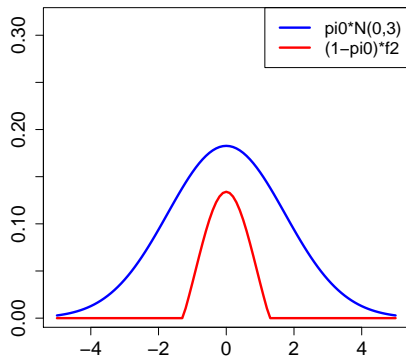
4 Additional results

5 One-sided alternatives

## Procedure BH\*



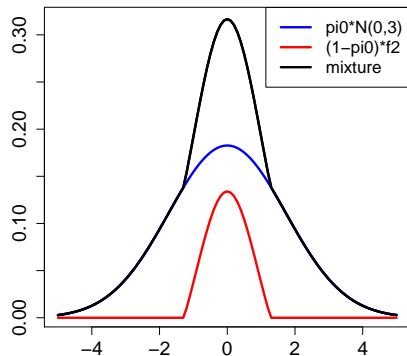
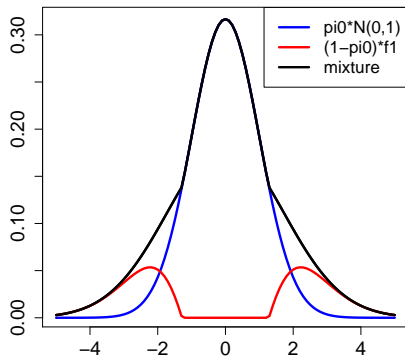
► rejects something



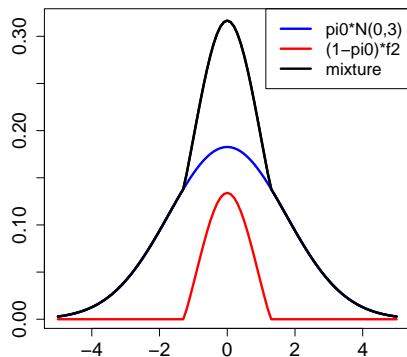
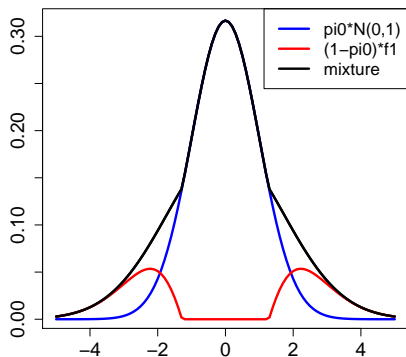
► does not reject anything



Any procedure  $R = R(Y)$



Does not distinguish between the two!

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**Not able to mimic  $BH^*$**

## Proposition 2 [R. and Verzelen (2020)]

For a sparsity  $k_n \gg n / \log(n)$ , there exists no optimal procedure.

Proof : Le Cam's two-point reduction scheme with the above configuration.

- ▶ for all  $n \geq c_1$ , any  $\alpha \in (0, 1)$ , any  $k$  with  $c_2 \frac{n \log(2/\alpha)}{\log(n)} \leq k < n/2$
- ▶ For any multiple testing procedure  $R$  such that

$$\text{FDR}(P, R) \leq c_3, \text{ for any } P \in \mathcal{P} \text{ with } n_1(P) \leq k,$$

- ▶ Then there exists some  $P \in \mathcal{P}$  with  $n_1(P) \leq k$  such that we have

$$|R(Y) \cap \mathcal{H}_1(P)| = 0 \text{ with } P\text{-proba} \geq 2/5$$

$$|\text{BH}_{\alpha/2}^* \cap \mathcal{H}_1(P)| \geq c_4 \alpha^{-1} n^{1/2} / \log^{1/2} n \text{ with } P\text{-proba} \geq 4/5.$$

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# Main result

## Theorem 1 [R. and Verzelen (2020)]

- (i) for a sparsity  $k_n \gg n/\log(n)$ , there exists no optimal procedure (of any kind);
- (ii) for a sparsity  $k_n \ll n/\log(n)$ ,  $\text{BH}_\alpha(\tilde{\theta}, \tilde{\sigma})$  is optimal with  $(\tilde{\theta}, \tilde{\sigma})$  above.

Procedure  $R$  optimal for a sparsity  $k_n$ : there exists  $\eta_n \rightarrow 0$ , s.t.

$$\text{(I)} \quad \limsup_n \sup_{\alpha \in (1/n, 1/2)} \left\{ \sup_{\substack{P \in \mathcal{P} \\ n_1(P) \leq k_n}} \{ \text{FDR}(P, R) \} - \alpha \right\} \leq 0$$

$$\text{(II)} \quad \lim_n \sup_{\alpha \in (1/n, 1/2)} \left\{ \sup_{\substack{P \in \mathcal{P} \\ n_1(P) \leq k_n}} \left\{ \mathbf{P}_{Y \sim P} \left( \text{TDP}(P, R) < \text{TDP}(P, \text{BH}_{\alpha(1-\eta_n)}^*) \right) \right\} \right\} = 0$$

1 Introduction

2 Upper bound

3 Lower bound

**4 Additional results**

5 One-sided alternatives

# No adaptation across boundary

Remark: always possible to achieve **(I)** by rejecting no null

Reformulation Theorem 1:

- (i) if  $k_n \gg n/\log(n)$ , possible to achieve **(I)** but not with **(II)**;
- (ii) if  $k_n \ll n/\log(n)$ , possible to achieve optimality (both **(I)** and **(II)**).

Procedure achieving (i) and (ii)?

NO !

Theorem 2 [R. and Verzelen (2020)]

- ▶ Any procedure achieving **(I)** for a sparsity  $k_n \gg n/\log(n)$  will fail to achieve optimality for a sparsity  $k_n \ll n/\log(n)$ .
- ▶ Any procedure achieving optimality for a sparsity  $k_n \ll n/\log(n)$  will fail to achieve **(I)** for some regime  $k_n \gg n/\log(n)$ .

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# Location model

## Case where $\sigma(P)$ is known

- ▶ only estimating  $\theta(P)$
- ▶ the sparsity boundary becomes  $n/\log^{1/2}(n)$

## Extension to non-Gaussian null $g(\cdot - \theta)$

- ▶  $g$  known, **symmetric**, continuous and non-increasing on  $\mathbb{R}_+$
- ▶ lower-bound and upper-bound matching up to some term
- ▶ Subbotin case:  $g(x) = L_\zeta^{-1} e^{-|x|^\zeta/\zeta}$ ,  $\zeta > 1$   
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- 1 Introduction
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# One-sided setting

One sided assumption:

- ▶ the  $P_i$ 's under the alternative are assumed  $\succeq \mathcal{N}(\theta, \sigma^2)$
- ▶ easier problem

**Proposition** [Carpentier, Dellatre, R., Verzelen (2020)]

Estimation of  $\theta$ :

- ▶ Identifiable as soon as  $k \leq n - 1$
- ▶ Minimax rate  $\frac{k/n}{\log^{1/2}(e\nu(k^2/n))}$  for sparsity  $1 \leq k \leq 0.9n$
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# Upper bound in one-sided case

Plugged  $\text{BH}_\alpha(\tilde{\theta}, \tilde{\sigma})$  (one-sided version) with new estimators:

$$\begin{cases} \tilde{\theta} = Y_{(q_n)} + \tilde{\sigma} \overline{\Phi}^{-1} \left( \frac{q_n}{n} \right) ; \\ \tilde{\sigma} = \frac{Y_{(q_n)} - Y_{(q'_n)}}{\overline{\Phi}^{-1}(q'_n/(n-\ell_0)) - \overline{\Phi}^{-1}(q_n/n)} , \end{cases}$$

for  $\ell_0 \leq \lfloor 0.9n \rfloor$ ,  $q_n = \lfloor n^{3/4} \rfloor$  and  $q'_n = \lfloor n^{1/4} \rfloor$ .

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# Outlook

## Take home message

- ▶ Challenging and useful direction of research
- ▶ First results on the feasibility of using empirical null in BH procedure
- ▶ Good news: weak sparsity  $k_n \ll n / \log(n)$  enough to mimic the oracle
- ▶ Bad news: it is needed

## Comments

- ▶ Robust minimax angle, so quite 'pessimistic'
- ▶ One-sided structure on the alternatives makes the problem easier

## Future work

- ▶ More structured alternatives
- ▶ Less structured nulls

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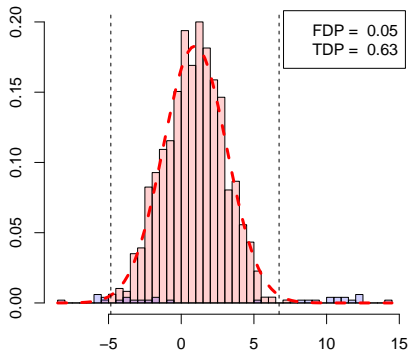
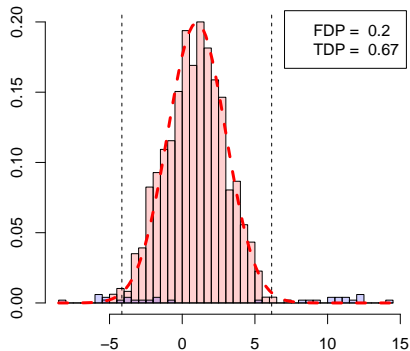
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# Illustration - Gaussian alternative

$$k = n^{1/2}$$

Oracle

Estimated

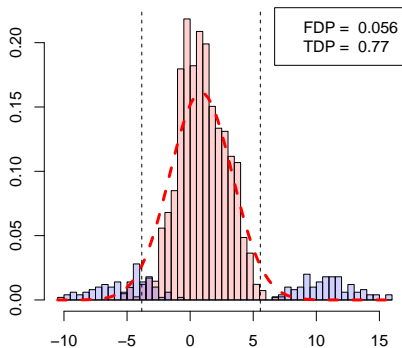
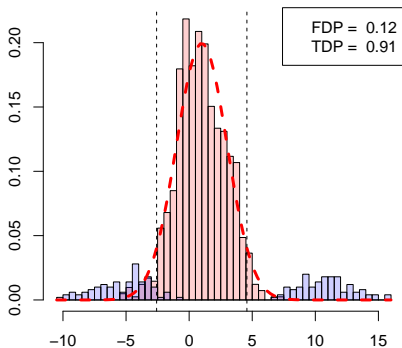


# Illustration - Gaussian alternative

$$k = n^{3/4}$$

Oracle

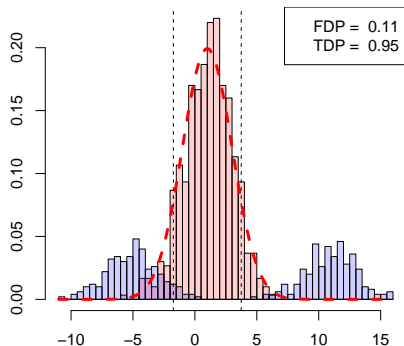
Estimated



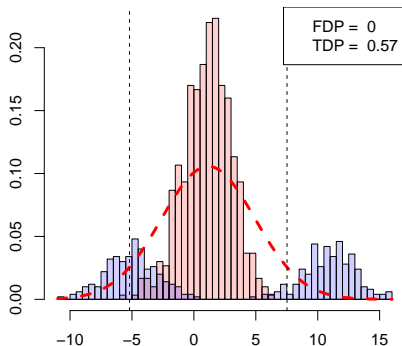
# Illustration - Gaussian alternative

$$k = 0.4n$$

Oracle



Estimated

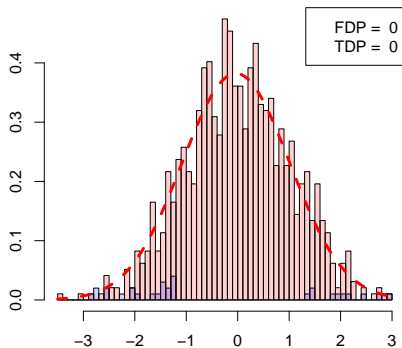
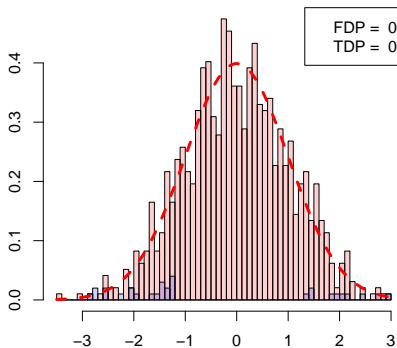


# Illustration - $f_1$ alternative

$$k = n^{1/2}$$

Oracle

Estimated



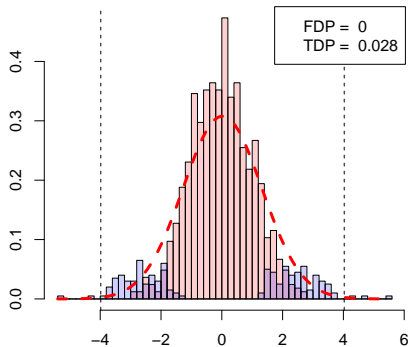
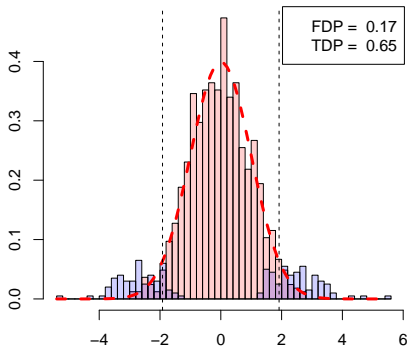


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Oracle

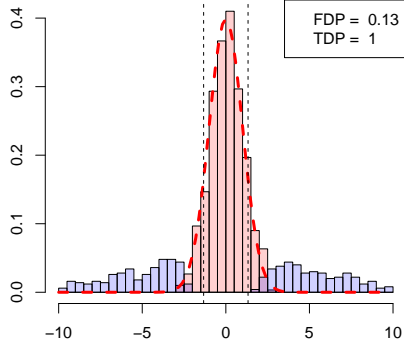
Estimated



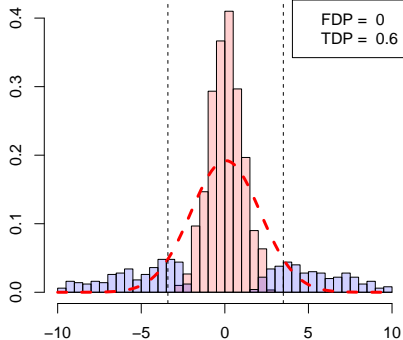
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Oracle



Estimated

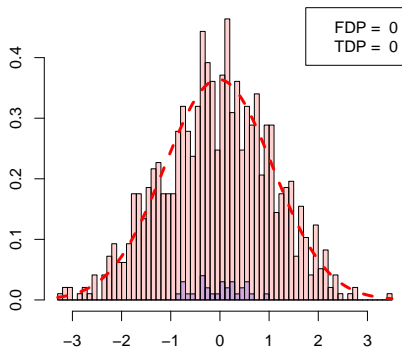
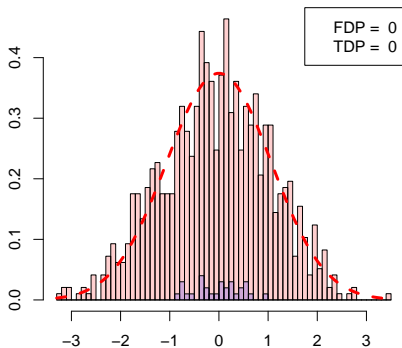


# Illustration - $f_2$ alternative

$$k = n^{1/2}$$

Oracle

Estimated

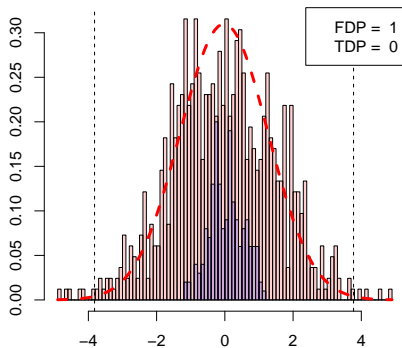
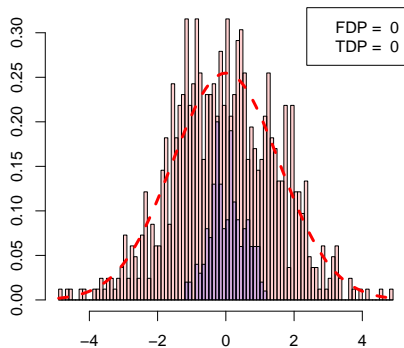


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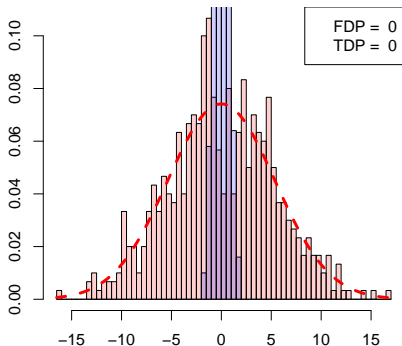
Estimated



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Oracle



Estimated

