

MULTIPLICATIVE ERGODICITY OF LAPLACE TRANSFORMS
FOR ADDITIVE FUNCTIONAL OF MARKOV CHAINS WITH
APPLICATION TO AGE-DEPENDENT BRANCHING
PROCESS.

Sana Louhichi

University Grenoble Alpes

Based on common works with *Bernard Ycart (2015)*
and *Loïc Hervé & Françoise Pène (2017)*

October 30, 2018

Introduction: classical model

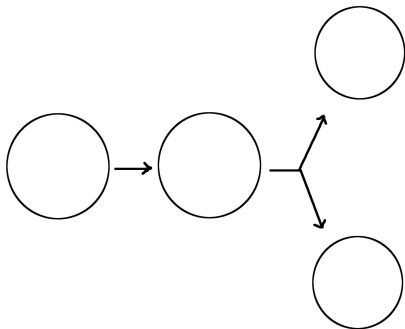
- 1 An object born at time 0 has a random life-length.

Introduction: classical model

- ① An object born at time 0 has a random life-length.
- ② At the end of its life it is replaced by some number of similar objects

Introduction: classical model

- 1 An object born at time 0 has a random life-length.
- 2 At the end of its life it is replaced by some number of similar objects
- 3 The time of birth of a child coincides with the time of death of the parent.



The process continue.

N_t : the number of objects present at time t .

$(N_t)_{t \geq 0}$: age-dependent branching process.

Age-dependent branching processes are mathematical models for

- 1 **Cell division:** mitosis
- 2 **Microbiology (growth of bacteria):** Certain phases of the multiplication of colonies of bacteria
- 3 **Models for the reproduction**

Purpose: asymptotic behavior of the process $(N_t)_{t \geq 0}$

We summarize the model, as follows,

- (A) to each cell v , is associated a parameter $x_v \in \mathbb{X}$, called its characteristics, (with $(\mathbb{X}, \mathcal{X})$ a measurable space) which determines its lifetime $\xi(x_v)$ and the number of new cells $\kappa(x_v)$ in which the cell splits at the end of its lifetime (where ξ and κ are two measurable functions with values in $[0, +\infty)$ and in \mathbb{Z}_+ respectively);
- (B) there exists a process $(X_n)_n$ with values in \mathbb{X} such that, for each line $(v_n)_{n \geq 0}$ of cells, the characteristics along this line is given by a copy of $(X_n)_{n \geq 0}$ (these copies are not assumed to be mutually independent);
- (C) $\kappa(x) \geq 2$ for any x , i.e. each cell gives birth to more than two children.

Classical assumptions and results about (N_t)

- **Classical assumptions:** Bellman and Harris(1952). **when the lifetimes are modeled by a sequence of i.i.d. random variables independent of the random numbers of the news cells which are also assumed to be i.i.d.**
- **Results:** N_t is a.s. asymptotically exponential:

$$N_t \sim C e^{\nu_0 t} \mathbf{W} \text{ as } t \text{ tends to infinity.}$$

ν_0 which determines the exponential rate of growth called *Malthusian parameter*.

Thomas Robert **Malthus** (1766-1834): a British economist. *In 1798 published the "Essay on the Principle of Population," which argued that population multiplies exponentially or geometrically and food arithmetically. Therefore, the population will outstrip the food supply.* Contributions : Malthusian growth model

- The growth rate ν_0 (the Malthusian parameter) was defined, in this context, as the positive root of the equation,

$$\mathbb{E}[\kappa(X_1)] \mathbb{E} \left[e^{-\nu_0 \xi(X_1)} \right] = 1, \quad (1)$$

as soon as the distribution of $\xi(X_1)$ is not lattice.

- The constant C equals to

$$C = \frac{\mathbb{E}[\kappa(X_1)] - 1}{\nu_0 (\mathbb{E}[\kappa(X_1)])^2 \mathbb{E} [\xi(X_1) e^{-\nu_0 \xi(X_1)}]}$$

- W is a positive random variable with finite second moment and $\mathbb{E}(W) = 1$. (Harris (1963))

Explicit calculations for the Malthusian parameter in the iid case

- if $\kappa(x) = 2$ for any x , then

$$E(e^{-\nu_0 \xi(X_1)}) = \frac{1}{2}, \quad C = \frac{1}{4\nu_0 E(\xi(X_1)e^{-\nu_0 \xi(X_1)})}.$$

- if $\kappa(x) = 2$ for any x and in the exponential case:
 $\xi(X_1) \sim \mathcal{E}(\lambda)$, then

$$\nu_0 = \lambda, \quad C = 1$$

Extension of Harris's results.

Louhichi and Ycart (2015) extend some results of Harris to the case where the lifetimes are a sequence of dependent random variables and when each cell is divided, after a random lifetime, into two cells: $(X_n)_n$ is a stationary process and $\kappa(x) = 2$ for any x : ν_1 is expressed in terms of the Laplace transform of S_n

$$S_n := \sum_{k=0}^n \xi(\mathbf{X}_k) \quad (2)$$

which models the birth date of the $(n + 1)$ -th individual of a same line. More precisely,

$$\nu_1 = \inf \left\{ \gamma > 0, \sum_{n \geq 0} 2^n \mathbb{E} \left[e^{-\gamma S_n} \right] < \infty \right\}. \quad (3)$$

$$\lim_{\gamma \rightarrow 0} \frac{\gamma}{\gamma + \nu_1} \sum_{n=1}^{\infty} 2^{n-1} \mathbb{E}(e^{-(\gamma + \nu_1) S_n}) =: \mathbf{C}_{\nu_1} < \infty.$$

Extension to random dependent variables $(\kappa(X_n))_n$

(I)- Calculation of $\mathbb{E}(N_t)$ in a general setting

We suppose $\kappa(x) \geq 2$ for any x and only the stationary assumption: for each line of cells $(0, v_1, \dots, v_{n-1})$ (n -sequence of the form parent-child),

$$(X_0, X_{v_1}, \dots, X_{v_{n-1}}) \sim (X_0, X_1, \dots, X_{n-1})$$

Extension to random dependent variables $(\kappa(X_n))_n$

(I)- Calculation of $\mathbb{E}(N_t)$ in a general setting

We suppose $\kappa(x) \geq 2$ for any x and only the stationary assumption: for each line of cells $(0, v_1, \dots, v_{n-1})$ (n -sequence of the form parent-child),

$$(X_0, X_{v_1}, \dots, X_{v_{n-1}}) \sim (X_0, X_1, \dots, X_{n-1})$$

Proposition (Hervé, Louhichi & Pène (2017))

Let $t > 0$ be fixed. If $\sum_{n \geq 0} \mathbb{E} \left[\left(\prod_{j=0}^n \kappa(X_j) \right) \mathbf{1}_{\{S_n \leq t\}} \right] < \infty$, then $\mathbb{E}[N_t] < \infty$ and

$$\mathbb{E}[N_t] = 1 + \sum_{n \geq 0} \mathbb{E} \left[\left(\prod_{j=0}^{n-1} \kappa(X_j) \right) (\kappa(X_n) - 1) \mathbf{1}_{\{S_n \leq t\}} \right] \quad (4)$$

(with the usual convention $\prod_{j=0}^{-1} \kappa(X_j) = 1$).

we obtain the following exponential behavior in mean of $\mathbb{E}[N_t]$ in a very general setting of dependence with the use of the function G given by

$$G(\gamma) := \sum_{n \geq 0} g_n(\gamma), \quad (5)$$

$$g_n(\gamma) = \mathbb{E} \left[\left(\prod_{j=0}^{n-1} \kappa(X_j) \right) (\kappa(X_n) - 1) e^{-\gamma S_n} \right].$$

$$\nu = \inf \{ \gamma > 0, \mathbf{G}(\gamma) < \infty \}.$$

$$\mathbf{C}_\nu := \lim_{\gamma \rightarrow 0} \frac{\gamma}{\gamma + \nu} \mathbf{G}(\nu + \gamma).$$

Corollary

Assume the stationary assumption and that $\nu < \infty$ and that the following limit exists

$$C_\nu := \lim_{\gamma \rightarrow 0} \frac{\gamma}{\gamma + \nu} G(\nu + \gamma). \quad (6)$$

Then,
$$\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t e^{-\nu s} \mathbb{E}[N_s] ds = C_\nu. \quad (7)$$

Proof. Let \tilde{A}_ν be the Laplace transform of A_ν . As a particular case of Feller (1971):

$$\lim_{\gamma \searrow 0} \gamma \tilde{A}_\nu(\gamma) = C \iff \lim_{t \rightarrow +\infty} \frac{1}{t} \int_0^t A_\nu(s) ds = C.$$

Here $A_\nu(s) = e^{-\nu s} \mathbb{E}[N_s]$.

Outline of the proof of the calculation on $\mathbb{E}(N_t)$

Harris's approach in the iid case. The result follows from [the renewal equation](#), denoting by $m(t) = \mathbb{E}(N_t)$,

$$m(t) = \mathbb{P}(\xi(X_0) > t) + \mathbb{E}(\kappa(X_0)) \int_0^t m(t-u) dF_{\xi(X_0)}(u)$$

Outline of the proof of the calculation on $\mathbb{E}(N_t)$

Harris's approach in the iid case. The result follows from the [renewal equation](#), denoting by $m(t) = \mathbb{E}(N_t)$,

$$m(t) = \mathbb{P}(\xi(X_0) > t) + \mathbb{E}(\kappa(X_0)) \int_0^t m(t-u) dF_{\xi(X_0)}(u)$$

Proof. Based on a control of the generating function $F(s, t) = \mathbb{E}(s^{N_t})$, $T_0 = \xi(X_0)$,

$$N_t \mathbb{I}_{T_0 < t} = \sum_{i=1}^{\kappa(X_0)} N_i(t) \mathbb{I}_{T_0 < t}, \quad (N_i(t) \mathbb{I}_{T_0 < t})_i \text{ iid} \sim N_{t-T_0}.$$

Outline of the proof of the calculation on $\mathbb{E}(N_t)$

Harris's approach in the iid case. The result follows from the [renewal equation](#), denoting by $m(t) = \mathbb{E}(N_t)$,

$$m(t) = \mathbb{P}(\xi(X_0) > t) + \mathbb{E}(\kappa(X_0)) \int_0^t m(t-u) dF_{\xi(X_0)}(u)$$

Proof. Based on a control of the generating function $F(s, t) = \mathbb{E}(s^{N_t})$, $T_0 = \xi(X_0)$,

$$N_t \mathbb{I}_{T_0 < t} = \sum_{i=1}^{\kappa(X_0)} N_i(t) \mathbb{I}_{T_0 < t}, \quad (N_i(t) \mathbb{I}_{T_0 < t})_i \text{ iid } \sim N_{t-T_0}.$$

$$\mathbb{E}(s^{N_t} \mathbb{I}_{T_0 < t}) = \mathbb{E} \left[\left(\mathbb{E}(s^{N_{t-T_0}} \mathbb{I}_{T_0 < t}) \right)^{\kappa(X_0)} \right].$$

$$F(s, t) = s\mathbb{P}(T_0 > t) + \int_0^t \mathbb{E}[(F(s, t-u))^{\kappa(X_0)}] dF_{T_0}(u), \quad |s| \leq 1, \quad t \geq 0.$$

Outline of the proof of the calculation on $\mathbb{E}(N_t)$

Direct calculations. For every $n \geq 0$, we write $\Sigma_n(t)$ for the number of cells of generation n alive at time t . Observe that $\mathbb{E}[\Sigma_0(t)] = \mathbb{P}(\xi(X_0) > t)$ and that, for every $n \geq 1$ (with the convention $k_0 = 0$),

$$\begin{aligned}\mathbb{E}[\Sigma_n(t)] &= \mathbb{E} \left[\sum_{k_1=1}^{D_0} \sum_{k_2=1}^{D_{0,k_1}} \cdots \sum_{k_n=1}^{D_{0,k_1,\dots,k_{n-1}}} \right. \\ &\quad \left. \mathbf{1}_{\{T_0 + T_{0,k_1} + \cdots + T_{0,k_1,\dots,k_{n-1}} \leq t < T_0 + T_{0,k_1} + \cdots + T_{0,k_1,\dots,k_n}\}} \right] \\ &= \mathbb{E} \left[D_0 D_{0,1} \cdots D_{0,1^{n-1}} \mathbf{1}_{\{T_0 + T_{0,1} + \cdots + T_{0,1^{n-1}} \leq t < T_0 + T_{0,1} + \cdots + T_{0,1^n}\}} \right] \\ &= \mathbb{E} \left[\left(\prod_{j=0}^{n-1} \kappa(X_j) \right) (\mathbf{1}_{\{S_{n-1} \leq t\}} - \mathbf{1}_{\{S_n \leq t\}}) \right].\end{aligned}$$

Multiplicative ergodicity, application to Markov chains

In order to study the function $G(\cdot)$, and so ν and C_ν , we adapt the notion of "multiplicative ergodicity", as introduced in Kontoyiannis & Meyn (2003-2005) to our context.

Definition

Let $\gamma_1 > 0$. We say that $(S_n, \kappa(X_n))_n$ is **multiplicatively ergodic** on $J = [0, \gamma_1)$ if there exist two continuous maps A and ρ from J to $(0, +\infty)$ such that, for every compact subset K of $(0, \gamma_1)$, there exist $M_K > 0$ and $\theta_K \in (0, 1)$ such that, for every $n \geq 1$,

$$\forall \gamma \in K, \quad |g_n(\gamma) - A(\gamma)(\rho(\gamma))^n| \leq M_K(\rho(\gamma)\theta_K)^n. \quad (8)$$

When $\kappa(\cdot)$ is constant, we will simply say that $(S_n)_n$ is multiplicatively ergodic on J .

Remark

Assume that $(S_n, \kappa(X_n))_n$ is multiplicatively ergodic on $J = [0, \gamma_1)$.
Then

- For every $\gamma \in J$ we have:

$$G(\gamma) = \sum_{n \geq 0} g_n(\gamma) < \infty \iff \rho(\gamma) < 1.$$

- For every compact subset K of J , we obtain that

$$\forall \gamma \in K \cap (\nu, +\infty), \quad \left| G(\gamma) - \frac{A(\gamma)}{1 - \rho(\gamma)} \right| \leq \frac{M_K}{1 - \rho(\gamma)\theta_K}.$$

- $\nu < \gamma_1$ means that

$$\nu = \inf\{\gamma \in \mathbf{J} : \rho(\gamma) < \mathbf{1}\} < \gamma_1.$$

- If moreover ρ is differentiable at ν with $\rho(\nu) = \mathbf{1}$ and $\rho'(\nu) \neq 0$, then (6) follows with

$$\mathbf{C}_\nu = -\frac{\mathbf{A}(\nu)}{\nu \rho'(\nu)}.$$

Multiplicative ergodicity property and additive functional of Markov chains

The multiplicative ergodicity property is specially adapted for additive functional of Markov chains, that is: $X = (X_n)_n$ is a Markov chain on $(\mathbb{X}, \mathcal{X})$ with Markov kernel $P(x, dy)$, invariant probability π , and initial distribution μ (i.e. μ is the distribution of X_0).

- *The Laplace kernel associated with (P, ξ, κ) .* We assume that, for every $n \geq 1$, the random variable $\prod_{j=0}^n \kappa(X_j)$ is integrable. We set $h_{\kappa, \gamma} := (\kappa - 1) e^{-\gamma \xi}$. Let $\gamma \in (0, +\infty)$. For $n \geq 1$,

$$\begin{aligned} g_n(\gamma) &= \mathbb{E} \left[\left(\prod_{j=0}^{n-1} \kappa(X_j) e^{-\gamma \xi(X_j)} \right) h_{\kappa, \gamma}(X_n) \right] \\ &= \mathbb{E} \left[\left(\prod_{j=0}^{n-1} \kappa(X_j) e^{-\gamma \xi(X_j)} \right) (Ph_{\kappa, \gamma})(X_{n-1}) \right], \end{aligned}$$

with $(Ph)(x) := \int_{\mathbb{X}} h(y) P(x, dy)$.

If $n \geq 2$, we continue and obtain

$$g_n(\gamma) = \mathbb{E} \left[\left(\prod_{j=0}^{n-2} \kappa(X_j) e^{-\gamma \xi(X_j)} \right) (P_\gamma(Ph_{\kappa, \gamma}))(X_{n-2}) \right],$$

with $P_\gamma h := P(h\kappa e^{-\gamma \xi})$. An easy induction gives

$$\forall n \geq 1, \quad g_n(\gamma) = \mu \left(\kappa e^{-\gamma \xi} P_\gamma^{n-1} (Ph_{\kappa, \gamma}) \right). \quad (9)$$

The multiplicative ergodicity property can be proved in the case when the Laplace kernels P_γ satisfy some nice spectral properties on a suitable Banach space \mathcal{B} :

$$\|P_\gamma^n f - r(\gamma)^n \Pi_\gamma f\|_{\mathcal{B}} \leq M_K (\theta_K r(\gamma))^n \|f\|_{\mathcal{B}}.$$

The multiplicative ergodicity property can be proved in the case when the Laplace kernels P_γ satisfy some nice spectral properties on a suitable Banach space \mathcal{B} :

$$\|P_\gamma^n f - r(\gamma)^n \Pi_\gamma f\|_{\mathcal{B}} \leq M_K (\theta_K r(\gamma))^n \|f\|_{\mathcal{B}}.$$

Then ν is proved to be finite and given by

$$\nu = \inf\{\gamma > \mathbf{0}, \mathbf{r}(\gamma) < \mathbf{1}\},$$

where $r(\gamma)$ denotes the spectral radius of P_γ on \mathcal{B} .

More precisely, the following assertions hold:

- (i) If the functions $\gamma \mapsto r(\gamma)$ and $\gamma \mapsto B(\gamma) := \mu(\kappa e^{-\gamma\xi} \Pi_\gamma(Ph_{\kappa,\gamma}))$ are continuous from J_0 to $(0, +\infty)$, then $(S_n, \kappa(X_n))_n$ is multiplicatively ergodic on J_0 with $A(\gamma) := \frac{B(\gamma)}{r(\gamma)}$ and $\rho(\gamma) = r(\gamma)$.
- (ii) If moreover $\inf_{\gamma \in J_0} r(\gamma) < 1 < \sup_{\gamma \in J_0} r(\gamma)$, then ν is finite and

$$\nu = \inf\{\gamma > 0 : r(\gamma) < 1\}. \quad (10)$$

- (iii) If furthermore the functions $r(\cdot)$ and $B(\cdot)$ are C^1 -smooth on J_0 , and if $r'(\nu) \neq 0$, then the constant C_ν is well defined and finite.

Theorem (Linear autoregressive model)

Let $\mathbb{X} := \mathbb{R}$ and $X_n = \alpha X_{n-1} + \vartheta_n$ for $n \geq 1$, where X_0 is a real-valued random variable, $\alpha \in (-1, 1)$, and $(\vartheta_n)_{n \geq 1}$ is a sequence of i.i.d. real-valued random variables independent of X_0 . Let $r_0 > 0$. We assume that ϑ_1 has a continuous Lebesgue probability density function $p > 0$ on \mathbb{X} satisfying the following condition: for all $x_0 \in \mathbb{R}$, there exist a neighbourhood V_{x_0} of x_0 and a non-negative function $q_{x_0}(\cdot)$ such that $y \mapsto (1 + |y|)^{r_0} q_{x_0}(y)$ is Lebesgue-integrable and such that

$$\forall y \in \mathbb{R}, \forall v \in V_{x_0}, p(y + v) \leq q_{x_0}(y). \quad (11)$$

Assume that the initial distribution μ is either the stationary probability measure π or δ_x for some $x \in \mathbb{R}$. Let N_0 be a positive integer. Assume that κ is bounded, that $\lim_{|x| \rightarrow +\infty} \xi(x) = +\infty$, that the Lebesgue measure of the set $[\xi = 0]$ is zero, and that $\sup_{x \in \mathbb{R}} \frac{\xi(x)}{(1+|x|)^{r_0}} < \infty$.

Then $(S_n, \kappa(X_n))_n$ is multiplicatively ergodic on $J = [0, +\infty)$ with $\lim_{\gamma} \rho(\gamma) \geq 2$ and $\lim_{\gamma \rightarrow +\infty} \rho(\gamma) = 0$.

Thus ν is well defined (and is independent of the choice of the initial distribution μ).

If moreover there exists $\tau > 0$ such that $\sup_{x \in \mathbb{R}} \frac{\xi(x)^{1+\tau}}{(1+|x|)^{\tau_0}} < \infty$, then the constant C_ν is well defined in $(0, +\infty)$.

(II) Second moment: behaviour of $\mathbb{E}[N_t N_{t+\tau}]$ and the a.s. convergence of $(e^{-\nu t} N_t)_{t \geq 0}$

we need an additional assumption involving the characteristics for lines of cells coinciding up to the k -th generation.

Hypothesis (Second assumption of stationarity)

The first stationary assumption holds true. Moreover, for each $k \in \mathbb{N}$, there exists a process $X^{(k)} = (X_n^{(k)})_{n \geq 0}$ such that

$$\begin{cases} (X_n^{(k)})_{0 \leq n \leq k} = (X_n)_{0 \leq n \leq k} & \text{a.s.} \\ (X_n^{(k)})_{n \geq 0} = (X_n)_{n \geq 0} & \text{in law,} \end{cases} \quad (12)$$

and such that, for every couple of sequences of positive integers $(m_i)_{i \geq 1}$ and $(\ell_i)_{i \geq 1}$ such that $m_1 = \ell_1, \dots, m_k = \ell_k$ and $\ell_{k+1} \neq m_{k+1}$, $((X_{0,m_1,\dots,m_n})_n, (X_{0,\ell_1,\dots,\ell_n})_n)$ has the same distribution as $(X, X^{(k)})$.

Now define, for any integers $n \geq 1$, $m \geq 1$ and $\min(n, m) - 1 \geq k \geq 0$ the random variables $A_{n,m,k}$ as follows:

$$A_{n,m,k} = \left(\prod_{i=0}^{n-2} \kappa(X_i) \right) \left(\prod_{j=\min(k+1, n-1)}^{m-2} \kappa(X_j^{(k)}) \right) \\ \left(\prod_{j \in \{k\} \setminus \{n-1, m-1\}} (\kappa(X_j) - 1) \right) (\kappa(X_{n-1}) - 1) (\kappa(X_{m-1}^{(k)}) - 1),$$

with the usual convention $\prod_{i=k+1}^{\ell} \dots = 1$ if $\ell \leq k$. Define also $S_n^{(k)} := \sum_{j=0}^n \xi(X_j^{(k)})$. The main result of this section is the following proposition.

Proposition (2)

Assume that the second assumption of stationarity holds. Let $t > 0$ and $\tau \geq 0$ be fixed. If

$\sum_{n \geq 0} \mathbb{E} \left[\left(\prod_{j=0}^n \kappa(X_j) \right) \mathbf{1}_{\{S_n \leq t + \tau\}} \right] < \infty$, then

$$\begin{aligned} \mathbb{E}[N_t N_{t+\tau}] &= \mathbb{E}[N_t] + \mathbb{E}[N_{t+\tau}] - 1 \\ &+ \sum_{n=1}^{\infty} \sum_{m=1}^{\infty} \sum_{k=0}^{\min(n,m)-1} \mathbb{E} \left[A_{n,m,k} \mathbf{1}_{\{S_{n-1} \leq t, S_{m-1}^{(k)} \leq t + \tau\}} \right]. \end{aligned}$$

Corollary

Assume that the assumptions of Proposition 2 are satisfied, that $\nu < \infty$, that $\limsup_{t \rightarrow \infty} e^{-\nu t} \mathbb{E}[N_t] < \infty$ and that there exists $K > 0$ such that

$$\limsup_{t \rightarrow \infty} \sup_{\tau \geq 0}$$

$$\left| e^{-\nu(2t+\tau)} \sum_{n=1}^{\infty} \sum_{m=1}^{\infty} \sum_{k=0}^{\min(n,m)-1} \mathbb{E} \left[A_{n,m,k} \mathbf{1}_{\{S_{n-1} \leq t, S_{m-1}^{(k)} \leq t+\tau\}} \right] - K \right| = 0$$

Then there exists a square integrable random variable W such that $e^{-\nu t} N_t$ converges in quadratic mean to W as t tends to infinity.

Corollary

Assume that the assumptions of Proposition 2 are satisfied, that $\nu < \infty$, that $\limsup_{t \rightarrow \infty} e^{-\nu t} \mathbb{E}[N_t] < \infty$ and that there exists $K > 0$ such that

$$\limsup_{t \rightarrow \infty} \sup_{\tau \geq 0}$$

$$\left| e^{-\nu(2t+\tau)} \sum_{n=1}^{\infty} \sum_{m=1}^{\infty} \sum_{k=0}^{\min(n,m)-1} \mathbb{E} \left[A_{n,m,k} \mathbf{1}_{\{S_{n-1} \leq t, S_{m-1}^{(k)} \leq t+\tau\}} \right] - K \right| = 0$$

Then there exists a square integrable random variable W such that $e^{-\nu t} N_t$ converges in quadratic mean to W as t tends to infinity. If moreover the above convergence is exponentially fast and if $W > 0$ then $e^{-\nu t} N_t$ converges almost surely to W as t tends to infinity.

Proof of Corollary 2.

$$\begin{aligned} & \mathbb{E} \left[\left(e^{-\nu t} N_t - e^{-\nu(t+\tau)} N_{t+\tau} \right)^2 \right] \\ &= e^{-2\nu t} \mathbb{E} [N_t^2] + e^{-2\nu(t+\tau)} \mathbb{E} [N_{t+\tau}^2] - 2e^{-2\nu t - \nu\tau} \mathbb{E} [N_t N_{t+\tau}]. \end{aligned}$$

Now Proposition 2 gives,

$$\begin{aligned} & e^{-2\nu t - \nu\tau} \mathbb{E} [N_t N_{t+\tau}] \\ &= e^{-2\nu t - \nu\tau} \mathbb{E} [N_t] + e^{-2\nu t - \nu\tau} (\mathbb{E} [N_{t+\tau}] - 1) \\ &+ e^{-2\nu t - \nu\tau} \sum_{n=1}^{\infty} \sum_{m=1}^{\infty} \sum_{k=0}^{\min(n,m)-1} \mathbb{E} \left[A_{n,m,k} \mathbf{1}_{\{S_{n-1} \leq t, S_{m-1}^{(k)} \leq t+\tau\}} \right]. \end{aligned}$$

Thanks to the assumptions of Corollary 2, the two first terms of the right hand side of the last equality tends to 0 as t tends to infinity. While the third term tends to K . Those three limits hold for any $\tau \geq 0$ and uniformly in τ . □

$$\lim_{t \rightarrow \infty} \mathbb{E} \left[\left(e^{-\nu t} N_t - e^{-\nu(t+\tau)} N_{t+\tau} \right)^2 \right] = K + K - 2K = 0,$$

for any $\tau \geq 0$, uniformly in τ . The Cauchy criterion ensures then the convergence in quadratic mean of $e^{-\nu t} N_t$ as t tends to infinity to a random variable W with finite second moment.

$$\lim_{t \rightarrow \infty} \mathbb{E} \left[\left(e^{-\nu t} N_t - e^{-\nu(t+\tau)} N_{t+\tau} \right)^2 \right] = K + K - 2K = 0,$$

for any $\tau \geq 0$, uniformly in τ . The Cauchy criterion ensures then the convergence in quadratic mean of $e^{-\nu t} N_t$ as t tends to infinity to a random variable W with finite second moment. For the last point, we deduce from Proposition 2 that

$\int_0^\infty \mathbb{E} \left[\left(e^{-\nu t} N_t - W \right)^2 \right] dt < \infty$. This yields (arguing as for the proof of Theorem 21.1 in [2]) the almost sure convergence, as t tends to infinity, of $e^{-\nu t} N_t$ to W .

(III)-Some extensions of Harris' results

For further results, we will make the following stronger assumption involving some independence assumptions.

Hypothesis

The sequence of "Children number" is a sequence of i.i.d. square integrable random variables of expectation κ_1 , and is independent of the sequence of life-time length. Moreover, for all $k \in \mathbb{N}$, $(X_n^{(k)})_{n \geq k+1}$ and $(X_n)_{n \geq k+1}$ are independent given X_k . Finally the number ν satisfies

$$\forall x \in \mathbb{X}, \quad \nu = \inf \left\{ \gamma > 0, \sum_{n \geq 0} \kappa_1^n \mathbb{E} \left[e^{-\gamma S_{n+1}} | X_0 = x \right] < \infty \right\} < \infty \quad (13)$$

We set $\kappa_2 := \mathbb{E}[\kappa(X_1)(\kappa(X_1) - 1)]$.

Remark

Observe that under this Hypothesis,

$$\mathbb{E} \left[\left(\prod_{j=0}^n \kappa(X_j) \right) \mathbf{1}_{\{S_n \leq t\}} \mid X_0 \right] = \kappa_1^n \kappa(X_0) \mathbb{E} \left[\mathbf{1}_{\{S_n \leq t\}} \mid X_0 \right] \leq \kappa_1^{n+1} \kappa(X_0) \mathbb{E} \left[e^{-\gamma(S_n - t)} \mid X_0 \right].$$

Hence, Proposition 1 applies and (4) can be rewritten

$$\mathbb{E}[\mathbf{N}_t] = \mathbf{1} + \sum_{n \geq 0} \kappa_1^n (\kappa_1 - \mathbf{1}) \mathbb{P}(S_n \leq t).$$

Technical lemma (Harris, 1963)

Let f be a function and \tilde{f} its Laplace transform. Suppose that there exist two positive reals δ and ϵ such that:

① \tilde{f} is analytic in $\{z = x + iy, |x| < \delta + \epsilon\} \setminus \{0\}$,

② \tilde{f} has a simple pole at 0, with residue C ,

③

$$\int_{-\infty}^{\infty} |\tilde{f}(\delta + iy)| dy < \infty,$$

④

$$\lim_{y \rightarrow \pm\infty} \tilde{f}(x + iy) = 0,$$

uniformly in $x \in [-\delta, \delta]$,

⑤

$$\psi := \int_{-\infty}^{\infty} |\tilde{f}(-\delta + iy)| dy < \infty.$$

Technical lemma (Harris, 1963)

Let f be a function and \tilde{f} its Laplace transform. Suppose that there exist two positive reals δ and ϵ such that:

① \tilde{f} is analytic in $\{z = x + iy, |x| < \delta + \epsilon\} \setminus \{0\}$,

② \tilde{f} has a simple pole at 0, with residue C ,

③

$$\int_{-\infty}^{\infty} |\tilde{f}(\delta + iy)| dy < \infty,$$

④

$$\lim_{y \rightarrow \pm\infty} \tilde{f}(x + iy) = 0,$$

uniformly in $x \in [-\delta, \delta]$,

⑤

$$\psi := \int_{-\infty}^{\infty} |\tilde{f}(-\delta + iy)| dy < \infty.$$

Then, for all $t > 0$,

$$|f(t) - C| \leq \frac{\psi}{2\pi} e^{-\delta t}.$$

Define $f_{x,0}(t) = (\kappa_1 - 1)e^{-\nu t} \sum_{n \geq 0} \kappa_1^n \mathbb{P}(S_{n+1} - S_0 \leq t | X_0 = x)$.
 We will make the following assumption involving the Laplace transform $\tilde{f}_{x,0}$ of $f_{x,0}$: $\forall \gamma > 0$,

$$\begin{aligned} \tilde{f}_{x,0}(\gamma) &= \int_0^\infty e^{-\gamma t} f_{x,0}(t) dt \\ &= \frac{\kappa_1 - 1}{\gamma + \nu} \sum_{n \geq 0} \kappa_1^n \mathbb{E} \left[e^{-(\gamma + \nu)(S_{n+1} - S_0)} | X_0 = x \right]. \end{aligned}$$

Hypothesis

Suppose that there exist two positive reals $\delta < \nu$ and ϵ such that, for any x , the Laplace transform $\tilde{f}_{x,0}$, extended on the complex plane, satisfies the following conditions:

- 1 $\tilde{f}_{x,0}$ has a simple pole at 0, with residue $\tilde{C}_0(x)$,
- 2 $\Psi_0(x) := \int_{-\infty}^{+\infty} |\tilde{f}_{x,0}(-\delta + iy)| dy < \infty$.

Lemma

Assume Hypothesis of the technical lemma. Then, for any $t > 0$,

$$\left| e^{-\nu t} \sum_{n \geq 0} \kappa_1^n (\kappa_1 - 1) \mathbb{P}(S_{1,n+1} \leq t | X_0 = x) - \tilde{C}_0(x) \right| \leq \frac{\Psi_0(x)}{2\pi} e^{-\delta t}.$$

Proposition

If, moreover, $\mathbb{E}[e^{-\nu\xi(X_0)} \tilde{C}_0(X_0)] < \infty$ and if $\mathbb{E}(e^{-(\nu-\delta)\xi(X_0)} \Psi_0(X_0)) < \infty$, then $\mathbb{E}[N_t] < \infty$ and there exists $\epsilon_1 > 0$ such that

$$\mathbb{E}[N_t] = e^{\nu t \kappa_1} \mathbb{E}[e^{-\nu\xi(X_0)} \tilde{C}_0(X_0)] (1 + O(e^{-\epsilon_1 t})), \text{ as } t \rightarrow \infty.$$

Proposition

If, moreover, $\mathbb{E}[e^{-\nu\xi(X_0)} \tilde{C}_0(X_0)] < \infty$ and if $\mathbb{E}(e^{-(\nu-\delta)\xi(X_0)} \Psi_0(X_0)) < \infty$, then $\mathbb{E}[N_t] < \infty$ and there exists $\epsilon_1 > 0$ such that

$$\mathbb{E}[N_t] = e^{\nu t} \kappa_1 \mathbb{E}[e^{-\nu\xi(X_0)} \tilde{C}_0(X_0)] (1 + O(e^{-\epsilon_1 t})), \text{ as } t \rightarrow \infty.$$

Proposition

Suppose, moreover, that $(X_n)_n$ a Markov process and also that $\sum_{k=0}^{\infty} \kappa_1^k \mathbb{E}[\tilde{C}_0^2(X_k) e^{-2\nu S_k}] < \infty$, then, for any $t > 0$, $\tau \geq 0$, , $\mathbb{E}[N_t N_{t+\tau}] < \infty$ and as $t \rightarrow \infty$

$$\mathbb{E}[N_t N_{t+\tau}] = e^{\nu(2t+\tau)} \kappa_2 \sum_{k=0}^{\infty} \kappa_1^k \mathbb{E} \left[\tilde{C}_0^2(X_k) e^{-2\nu S_k} \right] (1 + a e^{-\epsilon_1 t}),$$

where a and ϵ_1 are positive constants independent of t and τ .

Theorem

Assume Hypotheses of the two previous propositions (with $(X_n)_n$ a Markov process) then there exists a square integrable random variable W such that $e^{-\nu t} N_t$ converges in quadratic mean to W as t tends to infinity, with

$$\begin{aligned}\mathbb{E}[W] &= \kappa_1 \mathbb{E} \left[e^{-\nu \xi(X_0)} \tilde{C}_0(X_0) \right] \\ \text{Var}(W) &= \kappa_2 \sum_{k=0}^{\infty} \kappa_1^k \mathbb{E} \left[\tilde{C}_0^2(X_k) e^{-2\nu S_k} \right] \\ &\quad - \kappa_1^2 \left(\mathbb{E} \left[e^{-\nu \xi(X_0)} \tilde{C}_0(X_0) \right] \right)^2.\end{aligned}$$

If, moreover, $W > 0$ almost surely then $e^{-\nu t} N_t$ converges almost surely to W .

Remark.

Recall that, $S_{1,n+1} = \sum_{i=1}^{n+1} \xi(X_i)$. Suppose that

$$\mathbb{E} \left[e^{-\gamma S_{1,n+1}} | X_0 = x \right] = \alpha(\gamma, x) L^{n+1}(\gamma) + r_{n+1}(\gamma, x) \quad (14)$$

for suitable non-negative functions α , L and $(r_n)_n$ satisfying

- (a) $\forall x \in \mathbb{X}$, $\alpha(\cdot, x)$, L and $r_n(\cdot, x)$ can be extended to analytic functions in $\{z = u + iy, |u| \leq \delta + \epsilon < \nu, y \in \mathbb{R}\}$
- (b) L is positive and non-increasing on \mathbb{R}_+^* . The equation $\kappa_1 L(z) = 1$, has a unique positive solution in \mathbb{C} , denoted by ν .
- (c) The series $\sum_{n>0} \kappa_1^n r_n(\gamma, x)$ converges uniformly in γ in a neighborhood of ν uniformly in x .

Remark.

Recall that, $S_{1,n+1} = \sum_{i=1}^{n+1} \xi(X_i)$. Suppose that

$$\mathbb{E} \left[e^{-\gamma S_{1,n+1}} | X_0 = x \right] = \alpha(\gamma, x) L^{n+1}(\gamma) + r_{n+1}(\gamma, x) \quad (14)$$





for suitable non-negative functions α , L and $(r_n)_n$ satisfying





- (a) $\forall x \in \mathbb{X}$, $\alpha(\cdot, x)$, L and $r_n(\cdot, x)$ can be extended to analytic functions in $\{z = u + iy, |u| \leq \delta + \epsilon < \nu, y \in \mathbb{R}\}$
- (b) L is positive and non-increasing on \mathbb{R}_+^* . The equation $\kappa_1 L(z) = 1$, has a unique positive solution in \mathbb{C} , denoted by ν .
- (c) The series $\sum_{n>0} \kappa_1^n r_n(\gamma, x)$ converges uniformly in γ in a neighborhood of ν uniformly in x .

Lemma

The assumptions of the technical lemma are satisfied under Conditions (a)-(b)-(c) with

$$\tilde{C}_0(x) = -\frac{(\kappa_1 - 1)}{\kappa_1^2 \nu L'(\nu)} \alpha(\nu, x)$$

-  V. Baladi. *Positive transfer operators and decay of correlations*. World Scientific Publishing Co. Inc. (2000), River Edge, NJ.
-  T. E. HARRIS (1963). *The theory of branching processes*. Springer-Verlag, Berlin.
-  H. Hennion. *Sur un théorème spectral et son application aux noyaux lipchitziens*. Proc. Amer. Math. Soc. 118 (1993), 627-634.
-  L. Hervé, S. Louhichi, F. Pène (2017). Exponential growth of branching processes in a general context of lifetimes and birthtimes dependence.
<https://arxiv.org/abs/1701.00718>

-  G. Keller and C. Liverani. *Stability of the Spectrum for Transfer Operators*. Annali della Scuola Normale Superiore di Pisa - Classe di Scienze Sér. 4, 28 (1999), 141-152.
-  I. Kontoyiannis and S.P. Meyn. *Spectral Theory and Limit Theorems for Geometrically Ergodic Markov Processes*. Annals of Applied Probability 13 (2003), 304-362.
-  I. Kontoyiannis and S.P. Meyn. *Large deviations asymptotics and the spectral theory of multiplicatively regular Markov processes*. Electronic Journal of Probability 10(3) (2005), 61-123.
-  S. Louhichi and B. Ycart (2015). *Exponential growth of bifurcating processes with ancestral dependence*. Adv. Appl. Probab, 47(2), 545-564.