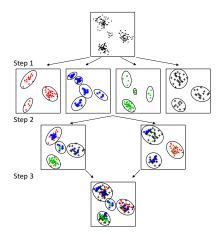
Scalable Bayesian Nonparametric Clustering and Classification

 P. MÜLLER, S. WILLIAMSON & M. DIESENDRUCK, UT Austin, D. A. ZUANETTI, UF Sao Carlos, Y. NI, TX A&M, and Y. JI, U Chicago



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Data: n = 85,021 patients

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Outcome: diabetes

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Outcome: diabetes

Goal: clustering, classification & prediction

Zodiac Data

Data: gene-gene interactions of n = 19,304 genes with all other genes. Data on gene-gene interactions from Zodiac (Zhu et al., 2015) with TCGA data.



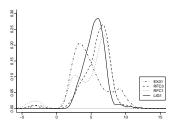
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Zodiac Data

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Goal: cluster genes by distribution of interactions (with all other n - 1 genes).

Clustering large data

Problem: clustering large (not "Big") data, $y_i, i = 1, ..., n$ Random partition: exchangeable partition of $[n] \equiv \{1, ..., n\} \Leftrightarrow$ ties under sampling from discrete random prob measure (Kingman, 1978)

$$heta_i \sim G$$
 and $G \sim H(G)$

cluster membership $s_i = j \Leftrightarrow \theta_i = \theta_j^*$ for *j*th unique value θ_j^* ; BNP prior H(G), e.g. DP

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Sampling model: together with $y_i \sim f(y_i \mid \theta_i)$, e.g., normal kernel, BNP mixture:

$$y_i \sim \int f(y_i \mid heta) \ dG(heta)$$
 and $G \sim \mathsf{DP},$

or any other BNP mixture model.

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- Predictive recursion: Newton, Quintana & Zhang (1998) use approximate predictive recursion, to approximate $p(y_{n+1} | y)$ under DP mixture model.

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Similar idea in Wang & Dunson (2011, JCGS) who sequentially build up clusters by assigning (i + 1) to a cluster in a partition of [i] (SUGS)

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$$g_i(heta)=(1-w_i)g_{i-1}(heta)+w_irac{f(y_i\mid heta)g_{i-1}(heta)}{c(y_i,g_{i-1})},$$

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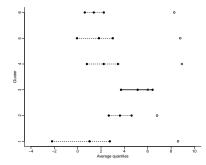
$$g_i(heta) = (1-w_i)g_{i-1}(heta) + w_i rac{f(y_i \mid heta)g_{i-1}(heta)}{c(y_i, g_{i-1})},$$

exact for i = 1 (and $w_i = 1/(1 + \alpha)$), and approx beyond.

Clustering: g_i builds up as a mixture model, which implicitely defines a random partition (with some computational simplifications, like dropping terms with very small weight, etc.)

GE-GE Interactions

Summarize each gene histogram by Jacobi polynomials \rightarrow clustering of $y_i \in \Re^8$;



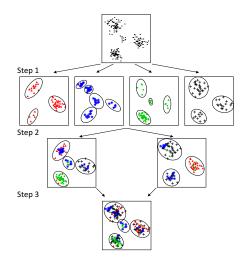
PRC clusters. Top (more than 1% of the genes) 6 clusters: average 25%, 50% and 75% quantiles (solid bullets) and average $10f_{0i}$ (empty circle).

Cluster 3 are the genes of interest.

1. split data into "shards";

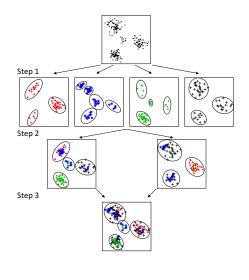
SIGN algorithm for BNP clustering

- 1. split data into "shards";
- 2. subset posterior on shards;



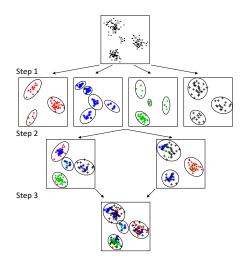
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- 1. split data into "shards";
- 2. subset posterior on shards;
- split clusters into shards → Step 2 with clusters as the new units;
- 4. stop when only one shard is left.



In Step 2, 3, ...: clustering of clusters; similar notion in Argiento et al. (2014) and Malsiner-Walli, Frühwirth-Schnatter & Grün (2017) (for mixture of non-Gaussian dists).

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Notation: let $\tilde{s}_i = j$ if *i*th cluster (of original units) joins the *j*th cluster of clusters.

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- Prior prob: exchangeable prior $\Leftrightarrow p(\rho) = f(n_1, \dots, n_C)$ for C clusters with cardinalities $n_c \Rightarrow$

$$p(\tilde{s}_i = c \mid \tilde{s}^{-i}) \propto \frac{p(\rho^{+c})}{p(\rho^{-i})}$$

for any BNP prior - easy;

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Transdim MCMC: Neal's (2000) Algorithm 8 for new singleton clusters

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Partition: $s_i = k$ if *i*th unit in *k*th cluster; alternatively use indicators $\delta_{ij} = I(s_i = s_j)$.

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Summary: judge approximation by

$$F_{.1} = \%$$
 pairs with $|E_q(\delta_{ij}) - E(\delta_{ij} \mid \mathbf{y})| > 0.1$

SIGN algorithm for BNP clustering

Simulations

Simulation I: simple mixture of $C_0 = 5$ (truth) normals; comparison with PY mixture (full MCMC) and DBSCAN (Ester et al., 1996)

	SIGN	SIGN-VI	PYM	DBSCAN			
С	4.94 (0.31)	4.90 (0.30)	5.08 (0.27)	3.64 (1.63)			
MISC	0.08 (0.03)	0.09 (0.03)	0.04 (0.01)	0.41 (0.11)			
MSE	0.01 (0.01)	0.01 (0.01)	0.01 (0.00)	0.49 (0.18)			
MISC=misclassification rate; MSE=estimation of							
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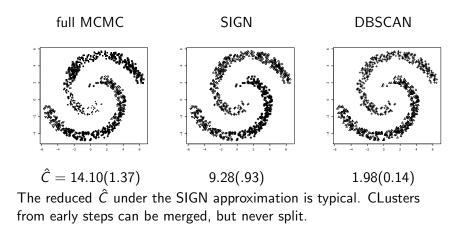
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MISC=misclassification rate; MSE=estimation of cluster-specific means

Estimated partition at the end of each step: need estimated clusters after each step; use Dahl (2006) summary. Alternatively, variation of information loss (SIGN-VI)

Simulation II: two spirals



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Prediction

Density estimation: BNP clustering,

 $p(\rho) p(\theta \mid \rho) p(-y_i \mid \rho, \theta),$

implies density estimation $p(y_{n+1} | \mathbf{y})$.

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Classification

Prediction

Density estimation: BNP clustering,

$$p(\rho) p(\theta \mid \rho) p(-y_i \mid \rho, \theta),$$

implies density estimation $p(y_{n+1} | \mathbf{y})$.

Regression & prediction: to be useful for regression, need conditioning in $p(\rho \mid \mathbf{x})$, on covariates x_i ; $p(y_{n+1} \mid x_{n+1}, \mathbf{y}, \mathbf{x})$ defindes desired regression.

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Predictive $p(x_{n+1}, y_{n+1} | \mathbf{x}, \mathbf{y})$ implies regression; conditional regression or density regression (Park & Dunson, 2010; M & al, 1996).

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PPMx: define mroe general $p(\rho \mid x)$, avoiding explicit modeling of a covariate distribution

Simulations

Classification: AUC. Comparison with full MCMC ("PPMx"), BART, random forest (RF), logistic regression (LR) and SVM

	Simulation III	Simulation IV
SIGN	0.808 (0.067)	0.838 (0.067)
PPMx	0.824 (0.060)	0.841 (0.063)
BART	0.755 (0.062)	0.866 (0.050)
RF	0.793 (0.059)	0.838 (0.067)
LR	0.600 (0.091)	0.524 (0.073)
SVM	0.622 (0.077)	0.585 (0.077)

Results

Response y_i : indicator for diabetes

- Covariates x_i: white blood cell count (WBC), red blood cell count (RBC), hemoglobin (HGB), platelets (PLT), fasting blood glucose (FBG), low density lipoproteins (LDL), total cholesterol (TC), triglycerides (Trig), triketopurine (Trik), high density lipoproteins (HDL), serum creatinine (SCr), serum glutamic oxaloacetic transaminase (SGOT), and total bilirubin (TB); sex, height, weight, blood pressure, and waist.
- Data: n = 85,021 patients
- SIGN: $M_1 = 250$ shards $\rightarrow 1351$ local clusters;
 - $M_2 = 5$ shards $\rightarrow 25$ regional clusters;

Algorithm stops at step K = 3

EHR - Results

AUC for classification by diabetes:

	EHR	Bank
SIGN	0.880	0.825
PPMx	-	-
BART	0.867	0.792
RF	0.869	0.786
LR	0.856	0.781
SVM	0.856	0.761

("Bank" is another data set, on success of telemarketing)

GAN: Chinese policy requires "China first" publication; We use a "Generative Adversarial Network" (GAN) (Goodfellow et al. 2014) to generate a hypothetical repeat

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Iterate until discrimination is impossible.

We comply with Chinese law, but statistical inference is identical

Simplification: replace clustering of clusters (and beyond) by deterministic match and merge.

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Clustering: Split data into shards with common overlap

Variations

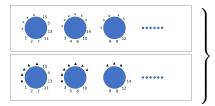
Simplification: replace clustering of clusters (and beyond) by deterministic match and merge.

Clustering: Split data into shards with common overlap

Consensus: Merge clusters C_1 , C_2 with m_{12} common members if $\min \left\{ \frac{m_{12}}{|C_1|}, \frac{m_{12}}{|C_1|} \right\} > \lambda$

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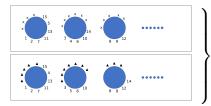
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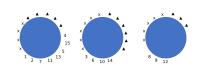


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Random subsets (feature allocation)

Feature allocation: Merge features F_1, F_2 if feature-specific parameters are close, $d(F_1, F_2) < \lambda$.

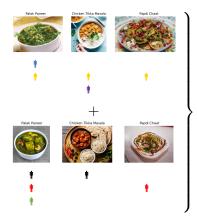
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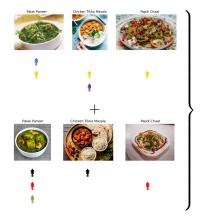
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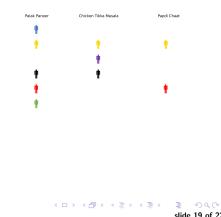
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Random subsets (feature allocation)

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Variations

Example: Tumor heterogeneity

• experimental unites = mutations i = 1, ..., n;

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Variations

Example: Tumor heterogeneity

- experimental unites = mutations i = 1, ..., n;
- Features = homogeneous subclones F_j ⊆ [n], subsets of mutations;

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Example: Tumor heterogeneity

- experimental unites = mutations i = 1, ..., n;
- Features = homogeneous subclones F_j ⊆ [n], subsets of mutations;
- Each subclone is linked with a set of weights, w_j, for observed tissue samples,

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use $d(w_j, w_\ell)$ to decide merging

Double feature allocation

Double feature allocation: two sets of experimental units,

i = 1, ..., n (e.g., patients) and s = 1, ..., S (e.g., symptoms); each feature $F_j \subseteq [n]$ (e.g., disease) is associated with a subset $S_j \subseteq [S]$.

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Double feature allocation

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Clustering of clusters: Same – merge features F_1, F_2 if feature-specific parameters S_1, S_2 are close, $d(S_1, S_2) < \lambda$.

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Clustering of clusters: Same – merge features F_1, F_2 if feature-specific parameters S_1, S_2 are close, $d(S_1, S_2) < \lambda$. Example: EHR, features = "disease", $F_j \subseteq [n]$, subsets of patients; Each disease is linked to a set S_j of symptoms, $d(S_j, S_\ell) = |S_j \cap S_\ell| / |S_j \cup S_\ell|$.

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Summary

- Model-based clustering (feature allocation, double FA) is more flexible than purely algorithmic methods, but computationally challenging for large n (also for large p)
- Several algorithms, using predictive recursion, approximation, parallelization, subset posteriors (consensus MC)
- Approxiamte posterior uncertainties important for decision problems (e.g., phenotype discovery in EHR data)