High-dimensional Bayesian inference of Graphical Models with Application to Brain Connectivity

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Bayesian Statistics in the Big Data Era



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Graphcial inference for brain connectivity

- Brain is complex system
- Treating neuro-deseases (e.g. Alzheimer) requires brain connectivity estimatoin
- Using graph theory to estimate brain connectivity
- Node is area of interest
- Link is connection between two area of brain





Brain connectivity and Alzheimer disease

- The most common neurodegenerative disease
- Can be assessed by various neuroimaging methods, MRI & PET



Dyrba, M., Mohammadi, A. & et al. (2017), Comparison of different hypotheses regarding the spread of Alzheimer's disease using Markov random fields and multimodal imaging, *Journal of Alzheimer Disease*

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Data and Aim

	control	stage 1	stage 2	AD
Sample size	179	269	134	85
Age	73.8	71.4	72.4	75.6
Education	16.6	15.9	16.5	15.7
Delayed recall	7.5	5.9	3.3	0.7



- 100 areas of brain
- 3 measurements

Aim

- Early detection of Alzheimer disease via graphical models
- Estimate brain connectivity for whole brain areas for each group

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Data collection



Graphcial inference for brain connectivity



Graphcial inference for brain connectivity



Software for graph estimation

• BDgraph & ssgraph: Multi-core R package





downloads 14K/month



Mohammadi, A. & Wit, E. (2017) BDgraph: An R Package for Bayesian Structure Learning in Graphical Models, *Journal of Statistical Software*, accepted

Mohammadi, A. & Dobra, A. (2017) "The R Package BDgraph for Bayesian Structure Learning in Graphical Models", *ISBA Bulletin*

Graphical models and conditional independence



Gaussian graphical model: Graph G = (V, E) as

 $X_1, ..., X_n \stackrel{iid}{\sim} \mathcal{N}_p(0, \Sigma), \quad K = \Sigma^{-1}$ is positive definite based on G

Graph, conditional independence and K:

$$X_{i}\perp X_{j} \mid X_{V\setminus\{i,j\}} \iff k_{ij} = 0$$

Bayesian structure learning in graphical models

$$Pr(G|data) = \frac{Pr(data|G)Pr(G)}{\sum_{G \in \mathcal{G}} Pr(data|G)Pr(G)}$$



Main problem

- Calculating posterior is not visible: graph-space $=2^{p(p-1)/2}$
- Key is to detect high posterior probability regimes

Solutions: Search algorithm

- Reversible-jump MCMC: Green (1995, 1999), Dobra et al (2011)
- Shotgun stochastic search: Jones et al (2005)
- Birth-death MCMC: Mohammadi and Wit (2015)

Our solution: Birth-death MCMC

- BD-MCMC is continuous time Markov process
- Based on special birth-death process (Preston, 1976)
- Stationary distribution is joint posterior distribution of graph and parameters



Advantages of BD-MCMC

- In BD-MCMC move between models are always accepted
- BD-MCMC converges more faster than RJ-MCMC



Mohammadi, A. & E. Wit (2015), Bayesian Structure Learning in Sparse Gaussian Graphical Models, *Bayesian Analysis*

BD-MCMC algorithm for GGMs

Theorem (Mohammadi and Wit, 2015)

Our birth-death MCMC algorithm has stationary distribution $Pr(G, K | \mathbf{x})$, if for each e = (i, j)

$$\beta_e(G)Pr(G,K|\mathbf{x}) = \delta_e(G^{-e})Pr(G^{-e},K^{-e}|\mathbf{x})$$





Preston (1976): Backward Kolmogorov

If balance conditions are hold, process converges to unique stationary distribution.

BD-MCMC search algorithm

- BD-MCMC is continuous time Markov birth-death process (Preston, 1976)
- Search in graph space: add/remove edge at birth/death event
- Stationary distribution is joint posterior of graph and parameters











BD-MCMC algorithm in GGMs

Step 1. Sampling from graph 1.1. Calculate birth and death rates

$$B_e(K) = \min\left\{\frac{P(G^{+e}, K^{+e}|x)}{P(G, K|x)}, 1\right\}$$

$$D_e(K) = \min\left\{\frac{P(G^{-e}, K^{-e}|x)}{P(G, K|x)}, 1\right\}$$

1.2. Calculate waiting time,
 1.3. Simulate type of jump, birth or death

 $Pr(\text{birth of edge } e) \propto B_e(K)$

 $Pr(\text{death of edge } e) \propto D_e(K)$

Step 2. Sampling from new precision matrix: K^{+e} or K^{-e}

Bayesian framework in GGMs

Likelihood

$$Pr(\mathbf{x}|K, G) \propto |K|^{n/2} \exp\left\{-\frac{1}{2}\mathrm{tr}(KS)\right\}, \quad \text{where } S = \mathbf{x}'\mathbf{x}$$

Prior for graph

- Discrete Uniform
- Truncated Poisson according to number of links

Prior for precision matrix

• G-Wishart: $W_G(b, D)$

$$p(K|G) \propto |K|^{(b-2)/2} \exp\left\{-\frac{1}{2}tr(DK)\right\}$$
$$I_G(b, D) = \int_{\mathbb{P}_G} |K|^{(b-2)/2} \exp\left\{-\frac{1}{2}tr(DK)\right\} dK$$

Computational challenge

$$D_{e}(K) = \frac{p(G^{-e}, K^{-e}|x)}{p(G, K|x)}$$

= $\left[\frac{I_{G}(b, D)}{I_{G^{-e}}(b, D)}\right] \left(\frac{|K^{-e}|}{|K|}\right)^{(b^{*}-2)/2} \exp\left\{-\frac{1}{2}\operatorname{tr}(D^{*}(K^{-e}-K))\right\}$

Normalizing constant of G-Wishart

$$\begin{array}{c} 1 & \hline & & \\ \hline & & & \\ \hline \hline & & & \\ \hline & & & \\ \hline \hline & & \\$$

Computational challenge

Exact form of Normalizing constant of G-Wishart

$$\begin{split} I_{G}(\delta,D) &= \pi^{|I|/2} \, \Gamma_{|B|} \left(\delta + \frac{1}{2} (|B|+1) \right) |D_{BB}|^{-(\delta + \frac{1}{2} (|B|+1))} \\ &\cdot \left(\det \partial_{I_{1},I_{2}}(D,T_{AA}) \right)^{-1/2} \\ &\cdot \sum_{\substack{0 \le j_{rs} < \infty \\ 1 \le r \le s \le |I|}} \left(\det \partial_{I_{1},I_{2}}(D,U_{AA}) \right)^{-j..} \\ &\cdot \left(\prod_{1 \le r \le s \le |I|} \frac{(1 + \delta_{rs})^{j_{rs}}}{j_{rs}!} \, D_{I_{1}(r),I_{2}(r)}^{j_{rs}} \, D_{I_{1}(s),I_{2}(s)}^{j_{rs}} \right) \\ &\cdot \left(\operatorname{Cof}_{rs} \partial_{I_{1},I_{2}}(D,U_{AA}) \right)^{j_{rs}} \right) \\ &\cdot I_{G_{A}}(\delta + \frac{1}{2} |I| + j_{..},U_{AA}) \, \bigg|_{U_{AA} = T_{AA}} \, \bigg|_{T_{AA} = D_{AA}} \end{split}$$

• Is it possible to implement it in model search algorithm?

Uhler, C., Lenkoski, A. & Richards, D. (2018) Exact formulas for the normalizing constants of Wishart distributions for graphical models, *Ann. Statist.*

Computational challenge

- Double Metropolis–Hasting (Murray & Ghahramani 2011)
- Our solution:

$$\frac{I_{G^{-e}}(b, I_p)}{I_G(b, I_p)} = \frac{1}{2\sqrt{\pi}} \frac{\Gamma(b/2)}{\Gamma((b+1)/2)} \frac{\mathbb{E}\left(\exp{-\frac{1}{2}\left(\sum_{(i,j)\in\bar{E}}\psi_{ij}^2 + \psi_e^2\right)}\right)}{\mathbb{E}(\exp{-\frac{1}{2}\sum_{(i,j)\in\bar{E}}\psi_{ij}^2})}.$$

$$\frac{I_{G^{-e}}(b, I_p)}{I_G(b, I_p)} \approx \frac{1}{2\sqrt{\pi}} \frac{\Gamma(\frac{b+d}{2})}{\Gamma(\frac{b+d+1}{2})}$$



Simulation study

ROC plot for 6 different graph Structures in GGMs with p=200, n=400



Mohammadi, A. & Wit, E. (2017) BDgraph: An R Package for Bayesian Structure Learning in Graphical Models, *Journal of Statistical Software*, accepted

Results for brain connectivity





Connections > 3000 How can we interpret connectivity network?

Dyrba, M., Mohammadi, A. & et al. (2018) "Assessing inter-modal and inter-regional dependencies in prodromal Alzheimer's disease using multimodal MRI/PET and Gaussian graphical models", arXiv:1804.00049

Results for brain connectivity

CN 0.6 EMCI Clustering coefficient LMCI AD 4,0 0.2 00 ю Characteristic path length 8 25 20 5 2 0.04 Small-world coefficient 0.03 0.02 0.01 0.00 BERE FACT John L VOI EMCI ANY ENCI 1980 LMCI TREADAD OLMO A HC

Comparison of graph statistics



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Extension to high-dimensional graphical models

Using Multivariate Birth-Death processes

• Instead of add/delete one link per time, add/delete multiple links

Using Pseudo-likelihood estimation

• replacing our likelihood with pseudo-likelihood:

$$Pr(G|x) = \frac{1}{Pr(x)} Pr(G) Pr(x|G)$$
$$Pr(x|G) = \int_{\theta_G} Pr(x, \theta_G|G)$$
$$Pr(x, \theta_G|G) \propto \sum_{i=1}^p Pr(x_i|x_V, \theta_G)$$



Dobra, A., & Mohammadi, R. (2018), Loglinear model selection and human mobility, *The Annals of Applied Statistics*

Wrap-up

Summary

- Develop Bayesian statistical methods for network analysis of high-dimensional data
- Compile methods into R-packages: BDgraph and ssgraph
- Applied our method to Alzheimer disease

Extension

• Applying to high-dimensional graphical models (Using Pseudo-likelihood estimation)

References



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