





# Introduction to data assimilation

Sebastian Reich (www.sfb1294.de)

Universität Potsdam/ University of Reading

CIRM2018, October 26, 2018

# Ensemble prediction I





Source: The quiet revolution of numerical weather prediction, Nature, 2015



#### **Ensemble prediction system** with *M* members:

 $\frac{\mathrm{d}}{\mathrm{d}t}Z_t^i=f(Z_t^i),\qquad Z_0^i\sim\pi_0\,,\qquad i=1,\ldots,M\,.$ 



#### Source: ECMWF





Source: The quiet revolution of numerical weather prediction, Nature, 2015

## Data assimilation II





Source: The quiet revolution of numerical weather prediction, Nature, 2015

## Data assimilation III





#### Model:

 $\mathrm{d} Z_t = f(Z_t) \mathrm{d} t + \gamma^{1/2} \mathrm{d} W_t$ 

**Data**:  $y_{t_n}$  at discrete times  $t_n$ 



#### Discrete-time observations:

$$y_{t_n} = h(Z_{t_n}) + R^{1/2} \Xi_{t_n}, \qquad n = 1, \dots, N.$$

#### Likelihood function:

$$L(z_{[0,T]}) := \exp\left(-\frac{1}{2}\sum_{n} \|y_{t_n} - h(z_{t_n})\|_R^2\right).$$

Bayes:

$$\frac{\mathrm{d}\widehat{\mathbb{P}}_{[0,T]}}{\mathrm{d}\mathbb{P}_{[0,T]}}(Z_{[0,T]}) := \frac{L(Z_{[0,T]})}{\mathbb{P}_{[0,T]}[L]} \,.$$

# Data assimilation V





## **prior** $\mathbb{P}_{[9:21]}$ in yellow;

posterior  $\widehat{\mathbb{P}}_{[9:21]}$  in green





- Daley, R., Atmospheric Data Analysis, Cambridge University Press, 1991.
- Bain, A. and Crisan, D., Fundamentals of Stochastic Filtering, Springer, 2009.
- SR, Cotter, J., Probabilistic Forecasting and Bayesian Data Assimilation, Cambridge University Press, 2015.
- Law, K. et al., Data Assimilation A Mathematical Introduction, Springer, 2015.
- P. Bauer et al, The quiet revolution of numerical weather prediction, Nature, 2015

# Sequential processing of data I





#### Notation:

Prediction/Forecast:
$$\pi(\widehat{z}_{n+1} | y_{1:n})$$
Filtering/Analysis: $\pi(z_{n+1} | y_{1:n+1})$ 

# Un<sup>iversita</sup>

## Classic **bootstrap particle filter** (sequential MC):

1) propagate *M* particles  $z_n^i$  under the model dynamics to yield **forecasts**  $\hat{z}_{n+1}^i$ , i = 1, ..., M.

2) data likelihood assigns **importance weights**  $w_{n+1}^i$  to each propagated particle  $\hat{z}_{n+1}^i$ 

3) **resample** to obtain equally weighted particles  $z_{n+1}^i$  from the filtering distribution at  $t_{n+1}$ .

# on<sup>iversita</sup>

# Classic **bootstrap particle filter** (sequential MC):

1) propagate *M* particles  $z_n^i$  under the model dynamics to yield **forecasts**  $\hat{z}_{n+1}^i$ , i = 1, ..., M.

2) data likelihood assigns **importance weights**  $w_{n+1}^i$  to each propagated particle  $\hat{z}_{n+1}^i$ 

3) **resample** to obtain equally weighted particles  $z_{n+1}^i$  from the filtering distribution at  $t_{n+1}$ .

This approach is **not applicable** to geophysical/meteorological applications because of:

- i) small sample sizes  $M \sim 10^2$
- ii) high dimensionality of models (discretised PDEs)  $D \sim 10^7$

iii) medium number of observations  $N \sim 10^5$ 



Remedies used in data assimilation:

1) Formulate assimilation of data as an **optimisation problem** (4DVar, 3DVar)



Remedies used in data assimilation:

- Formulate assimilation of data as an **optimisation problem** (4DVar, 3DVar)
- 2) Replace importance–resampling step by low variance **ensemble transformations**

$$z_{n+1}^{j} = \sum_{i=1}^{M} \widehat{z}_{n+1}^{i} t_{ij}^{*}$$



Remedies used in data assimilation:

- Formulate assimilation of data as an **optimisation problem** (4DVar, 3DVar)
- 2) Replace importance–resampling step by low variance **ensemble transformations**

$$z_{n+1}^{j} = \sum_{i=1}^{M} \widehat{z}_{n+1}^{i} t_{ij}^{*}$$

3) Localise assimilation of data





Forward observation model:

$$y_{n+1}^{obs} = h(\hat{z}_{n+1}) + R^{1/2} \Xi_{n+1}$$

**Best linear unbiased estimator** (BLUE):  $Z_{n+1} = Ay_{n+1}^{obs} + b$ 



Forward observation model:  $y_{n+1}^{obs} = h(\hat{z}_{n+1}) + R^{1/2} \Xi_{n+1}$ 

**Best linear unbiased estimator** (BLUE):  $Z_{n+1} = Ay_{n+1}^{obs} + b$ 

**Empirical BLUE/EnKF**:

$$A = PR^{-1}$$
  
$$b^{j} = \hat{z}_{n+1}^{j} - A\left(h(\hat{z}_{n+1}^{j}) + R^{1/2} \hat{\Xi}_{n+1}^{j}\right)$$

with

$$P = \frac{1}{M-1} \sum_{i=1}^{M} \widehat{z}_{n+1}^{i} (h(\widehat{z}_{n+1}^{i}) - \overline{h}_{n+1})^{T}, \qquad \widehat{\Xi}_{n+1}^{i} \sim N(0, I).$$



Forward observation model:  $y_{n+1}^{obs} = h(\hat{z}_{n+1}) + R^{1/2} \Xi_{n+1}$ 

**Best linear unbiased estimator** (BLUE):  $Z_{n+1} = Ay_{n+1}^{obs} + b$ 

**Empirical BLUE/EnKF**:

$$A = PR^{-1}$$
  

$$b^{j} = \hat{z}_{n+1}^{j} - A\left(h(\hat{z}_{n+1}^{j}) + R^{1/2} \hat{=}_{n+1}^{j}\right)$$

with

$$P = \frac{1}{M-1} \sum_{i=1}^{M} \widehat{z}_{n+1}^{i} (h(\widehat{z}_{n+1}^{i}) - \overline{h}_{n+1})^{T}, \qquad \widehat{\Xi}_{n+1}^{i} \sim N(0, I).$$

EnKF induced particle transformation:

$$z_{n+1}^{j} = \sum_{i=1}^{M} \widehat{z}_{n+1}^{i} t_{ij}^{*}$$
.



Importance weights:

$$w_{n+1}^{i} \propto \exp\left(-\frac{1}{2} \|h(\hat{z}_{n+1}^{i}) - y_{n+1})\|_{R}^{2}\right)$$



Importance weights:

$$w_{n+1}^{i} \propto \exp\left(-\frac{1}{2} \|h(\widehat{z}_{n+1}^{i}) - y_{n+1})\|_{R}^{2}\right)$$





















Lorenz-63 model, first component observed infrequently ( $\Delta t = 0.12$ ) and with large measurement noise (R = 8):



Figure: RMSEs for various second-order accurate LETFs compared to the ETPF, the ESRF, and the SIR PF as a function of the sample size, M.



Hybrid filter:  $\mathbf{P} := \mathbf{P}_{\mathsf{ESRF}}(\alpha) \mathbf{P}_{\mathsf{ETPF}}(1-\alpha)$ .



Figure: RMSEs for hybrid ESRF ( $\alpha = 0$ ) and 2nd-order corrected NETF/ETPF ( $\alpha = 1$ ) as a function of the sample size, *M*.



- Evensen, G., Data assimilation the ensemble Kalman filter, Springer, 2009.
- SR, A nonparametric ensemble transform method for Bayesian inference, SIAM J. Sci. Comput., 2013.
- SR, Cotter, J., Probabilistic Forecasting and Bayesian Data Assimilation, Cambridge University Press, 2015.
- Tödter J., Ahrens, B., A second-order exact ensemble square root filter for nonlinear data assimilation, Month. Weather Rev., 2015.
- Chustagulprom, N. et al., A hybrid ensemble transform particle filter for nonlinear and spatially extended systems, SIAM/ASA J. UQ, 2016.
- Acevedo, W., et al., Second-order accurate ensemble transform particle filters, SIAM J. Sci. Comput., 2017.
- Fearnhead, P., Künsch, H.R., Particle filters and data assimilation, Annual Review of Statistics and Its Application, 2018.



#### **Observation model**:

$$\mathrm{d} y_t = h(Z_t) \mathrm{d} t + R^{1/2} \mathrm{d} V_t$$



### **Observation model**:

$$\mathrm{d} y_t = h(Z_t) \mathrm{d} t + R^{1/2} \mathrm{d} V_t$$

## Linear SDEs and observations:

$$dZ_t = (AZ_t + b) dt + \gamma^{1/2} dW_t, \qquad dy_t = HZ_t dt + R^{1/2} dV_t.$$

## **Observation model**:

$$\mathrm{d} y_t = h(Z_t) \mathrm{d} t + R^{1/2} \mathrm{d} V_t$$

### Linear SDEs and observations:

$$\mathrm{d} Z_t = (A Z_t + b) \, \mathrm{d} t + \gamma^{1/2} \mathrm{d} W_t, \qquad \mathrm{d} y_t = H Z_t \mathrm{d} t + R^{1/2} \mathrm{d} V_t.$$

#### Kalman-Bucy equations for the mean

$$d\mu_t = (A\mu_t + b)dt - K_t dI_t, \qquad dI_t := (H\mu_t dt - dy_t)$$

and the covariance matrix

$$\frac{\mathrm{d}}{\mathrm{d}t}P_t = AP_t + P_tA^T + \gamma I - K_tHP_t.$$

with Kalman gain matrix  $K_t = P_t H^T R^{-1}$ .





Controlled mean-field SDE:

 $\mathrm{d} Z_t = (A Z_t + b) \mathrm{d} t + \gamma^{1/2} \mathrm{d} W_t + d u_t (Z_t), \qquad d u_t := -K_t \mathrm{d} \mathcal{I}_t$ 

with innovation  $\mathcal{I}_t$ 

$$\mathrm{d}\mathcal{I}_t = \frac{1}{2}(HZ_t + H\mu_t)\mathrm{d}t - \mathrm{d}y_t$$

or

$$\mathrm{d}\mathcal{I}_t = HZ_t \mathrm{d}t + R^{1/2} \mathrm{d}U_t - \mathrm{d}y_t$$

and Kalman gain  $K_t = P_t H^T R^{-1}$ .



Controlled mean-field SDE:

 $\mathrm{d} Z_t = (A Z_t + b) \mathrm{d} t + \gamma^{1/2} \mathrm{d} W_t + \mathrm{d} u_t(Z_t), \qquad \mathrm{d} u_t := -K_t \mathrm{d} \mathcal{I}_t$ 

with innovation  $\mathcal{I}_t$ 

$$\mathrm{d}\mathcal{I}_t = \frac{1}{2}(HZ_t + H\mu_t)\mathrm{d}t - \mathrm{d}y_t$$

or

$$\mathrm{d}\mathcal{I}_t = HZ_t \mathrm{d}t + R^{1/2} \mathrm{d}U_t - \mathrm{d}y_t$$

and Kalman gain  $K_t = P_t H^T R^{-1}$ .

Extension to nonlinear SDEs:

Ensemble Kalman-Bucy filter(EnKBF)

 $\mathrm{d} t Z_t = f(Z_t) \mathrm{d} t + \gamma^{1/2} \mathrm{d} W_t - K_t \mathrm{d} \mathcal{I}_t \,.$ 



# Feedback particle filter:

$$\mathrm{d} Z_t = f(Z_t) \mathrm{d} t + \gamma^{1/2} \mathrm{d} W_t - K_t \circ \mathrm{d} \mathcal{I}_t$$

with Kalman gain

$$K_t := \nabla_z \psi_t R^{-1}, \qquad \nabla_z \cdot (\pi_t \nabla_z \psi_t) = -\pi_t (h - \bar{h})$$

and innovations

$$\mathrm{d}\mathcal{I}_t = \frac{1}{2}(h(Z_t) + \bar{h}_t)\mathrm{d}t - \mathrm{d}y_t$$

or

$$\mathrm{d}\mathcal{I}_t = h(Z_t)\mathrm{d}t + R^{1/2}\mathrm{d}U_t - \mathrm{d}y_t\,.$$



Particle approximation to elliptic PDE:

$$\nabla_z \psi_t(z_t^j) \approx \sum_{i=1}^M z_t^i a_t^{ij}$$

with  $z_t^i \sim \pi_t$ , i = 1, ..., M and coefficients  $a_t^{ij} \in \mathbb{R}$  appropriately defined.



Particle approximation to elliptic PDE:

$$\nabla_z \psi_t(z_t^j) \approx \sum_{i=1}^M z_t^i a_t^{ij}$$

with  $z_t^i \sim \pi_t$ , i = 1, ..., M and coefficients  $a_t^{ij} \in \mathbb{R}$  appropriately defined.

Interacting particle system:

$$\mathrm{d} \boldsymbol{z}_t^j = f(\boldsymbol{z}_t^j) \mathrm{d} t + \gamma^{1/2} \mathrm{d} \boldsymbol{W}_t^j - \left(\sum_{i=1}^M \boldsymbol{z}_t^i \, \boldsymbol{a}_t^{ij}\right) \boldsymbol{R}^{-1} \circ \mathrm{d} \boldsymbol{\mathcal{I}}_t^j$$

with innovation

$$\mathrm{d}\mathcal{I}_t^j = h(z_t^j)\mathrm{d}t + R^{1/2}\mathrm{d}U_t^j - \mathrm{d}y_t\,.$$



#### Parameter-dependent model:

$$\mathrm{d} Z_t = f(Z_t, \, {oldsymbol heta}) \mathrm{d} t + \gamma^{1/2} \mathrm{d} W_t - K_t \circ \mathrm{d} \mathcal{I}_t$$

**Marginal likelihood** (evidence of the data  $\mathcal{Y}_t = y_{[0,t]}$  under  $\theta$ ):

 $\mathrm{d}\pi_t(\mathcal{Y}_t|\theta) = \pi_t(\mathcal{Y}_t|\theta)\,\bar{h}_t R^{-1}\mathrm{d}y_t\,.$ 



### Parameter-dependent model:

$$\mathrm{d} Z_t = f(Z_t, \, \theta) \mathrm{d} t + \gamma^{1/2} \mathrm{d} W_t - K_t \circ \mathrm{d} \mathcal{I}_t$$

**Marginal likelihood** (evidence of the data  $\mathcal{Y}_t = y_{[0,t]}$  under  $\theta$ ):

$$\mathrm{d}\pi_t(\mathcal{Y}_t|\boldsymbol{\theta}) = \pi_t(\mathcal{Y}_t|\boldsymbol{\theta})\,\bar{h}_t R^{-1} \mathrm{d}y_t\,.$$

#### Example.

SPDE model:

$$\dot{u}_t = \theta \Delta u_t + \dot{W}_t$$

 $\dot{W}_t$  space–time white noise.

Observation:

$$\mathrm{d} y_t = u_t(x_0)\,\mathrm{d} t + R^{1/2}\,\mathrm{d} V_t$$

for fixed spatial location  $x_0$ .

## Parameter estimation II



D = 80 grid points, state estimation with EnKBF, M = 20 particles, and localisation, observation interval [0, 20], measurement error R = 0.001





- Bergemann, K, SR, An ensemble Kalman–Bucy filter for continuous–time data assimilation, Meteorolog. Zeitschrift, 2012.
- Yang, T. et al., Feedback particle filter, IEEE Trans. Autom. Control, 2013.
- Taghvaei, A. et al., Kalman filter and its modern extensions for the continuous-time nonlinear filtering problem, J. Dyn. Sys., Meas., and Control, 2018.
- de Wiljes, et al., Long-time stability and accuracy of the ensemble Kalman–Bucy filter for fully observed processes and small measurement noise, SIAM J. Appl. Dyn. Sys., 2018.
- SR, Data assimilation, Acta Numerica, 2019, preliminary version on arXiv 1807.08351.



- Data assimilation has become a workhorse for day-to-day weather forecasting
- Key are a "smart" choice of biased and robust approximations to the sequential filtering problem
- Data assimilation has applications to many areas outside weather prediction such as medicine, geosciences, biology, etc.
- Current topics of research: combined state-parameter estimation for PDEs, theoretical understanding of approximations such as linear transformations & localisation
- Continuous-in-time DA naturally leads to an attractive unifying framework in the form of mean-field SDEs.





"Miss Peterson, may I go home? I can't assimilate any more data today."

(source: J.B. Handelsman, New Yorker, 05/31/1969)



- Walter Acevedo
- Kay Bergemann
- Yuan Cheng
- Nawinda Chustagulprom
- Colin Cotter
- Jana de Wiljes
- Prashant Mehta
- Wilhelm Stannat
- Amari Taghvaei