

# Introduction to Bayesian Statistical Modelling and Analysis

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# Getting Started: Screening for cancer

- Medical diagnostic tests are very rigorous:
  - The chance of a positive test given cancer
$$\Pr(+|CA) = 0.90$$
  - The chance of a negative test given no cancer
$$\Pr(-|Not) = 0.95$$

*But what is the question?*

# What is the question?

- From the perspective of the test:

*Given a person has cancer, what is the probability that the test is positive?*

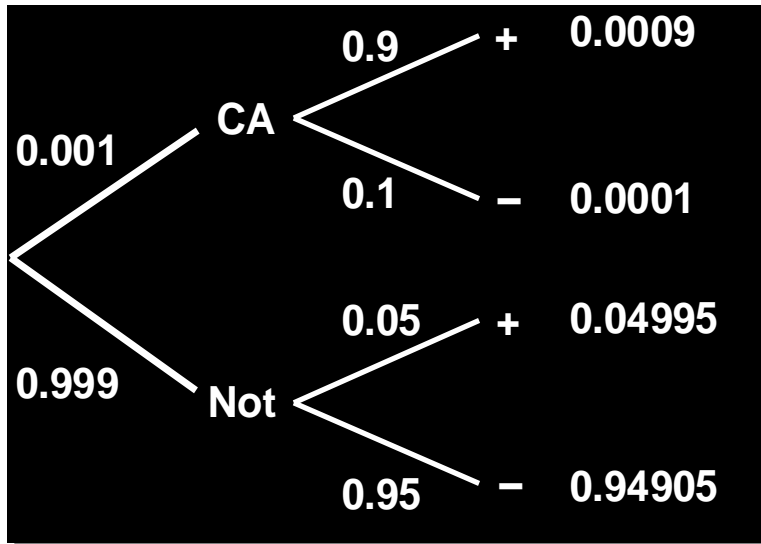
$$\Pr(+|CA) = 0.90$$

- From the perspective of the person:

*Given that the test is positive, what is the probability that a person has cancer?*

$$\Pr(CA|+) = ?$$

# What is the answer? via Bayes' Theorem



$$\Pr(\text{CA}|+) = \Pr(\text{CA and } +) / \Pr(+)$$

$$\begin{aligned}\Pr(\text{CA and } +) &= \Pr(\text{CA}) \Pr(+|\text{CA}) \\ &= 0.0009\end{aligned}$$

$$\Pr(+) = 0.0009 + 0.04995 = 0.05085$$

$$\Pr(\text{CA}|+) = 0.0009 / 0.05085 = 0.02$$

$$\Pr(\text{CA}|+) = \Pr(+|\text{CA}) \Pr(\text{CA}) / \Pr(+)$$

## Your turn

Suppose the diagnostic test has the same sensitivity and specificity but the cancer is more common: 10% of the population has a chance of getting this cancer.

What is the probability that a person has cancer, given that the test is positive?

# Solution

$$\Pr(+|CA) = 0.9$$

$$\Pr(-|CA) = 0.95$$

$$\Pr(CA) = 0.1$$

$$\begin{aligned}\Pr(+) &= \Pr(+|CA) \Pr(CA) + \Pr(+|Not) \Pr(Not) \\ &= 0.9*0.1 + 0.05*0.9 = 0.135\end{aligned}$$

$$\begin{aligned}\Pr(CA|+) &= \Pr(+|CA) \Pr(CA) / \Pr(+) \\ &= 0.9*0.1 / 0.135 = 0.67\end{aligned}$$

# Bayesian Modelling

Overall Aim: To learn about unknowns  $\theta$ , given information  $y$

$$p(\theta | y) = p(y | \theta) p(\theta) / p(y)$$

$$p(\theta | y) = p(\theta) p(y | \theta) / p(y)$$

# Frequentist approach to modelling

We have some data  $Y$ , and want to know about  $\theta$

$\theta$  can be unknown parameters, missing data, latent variables, etc.

Frequentist: estimate  $\theta$  through the likelihood:  $p(Y|\theta)$

How likely is  $Y$  for given values of  $\theta$  ?

Use moment estimators or maximum likelihood.

But we *really* want to know about  $p(\theta|Y)$

→ **Bayes**

# Why Bayes?

Bayesian methods allow us to:

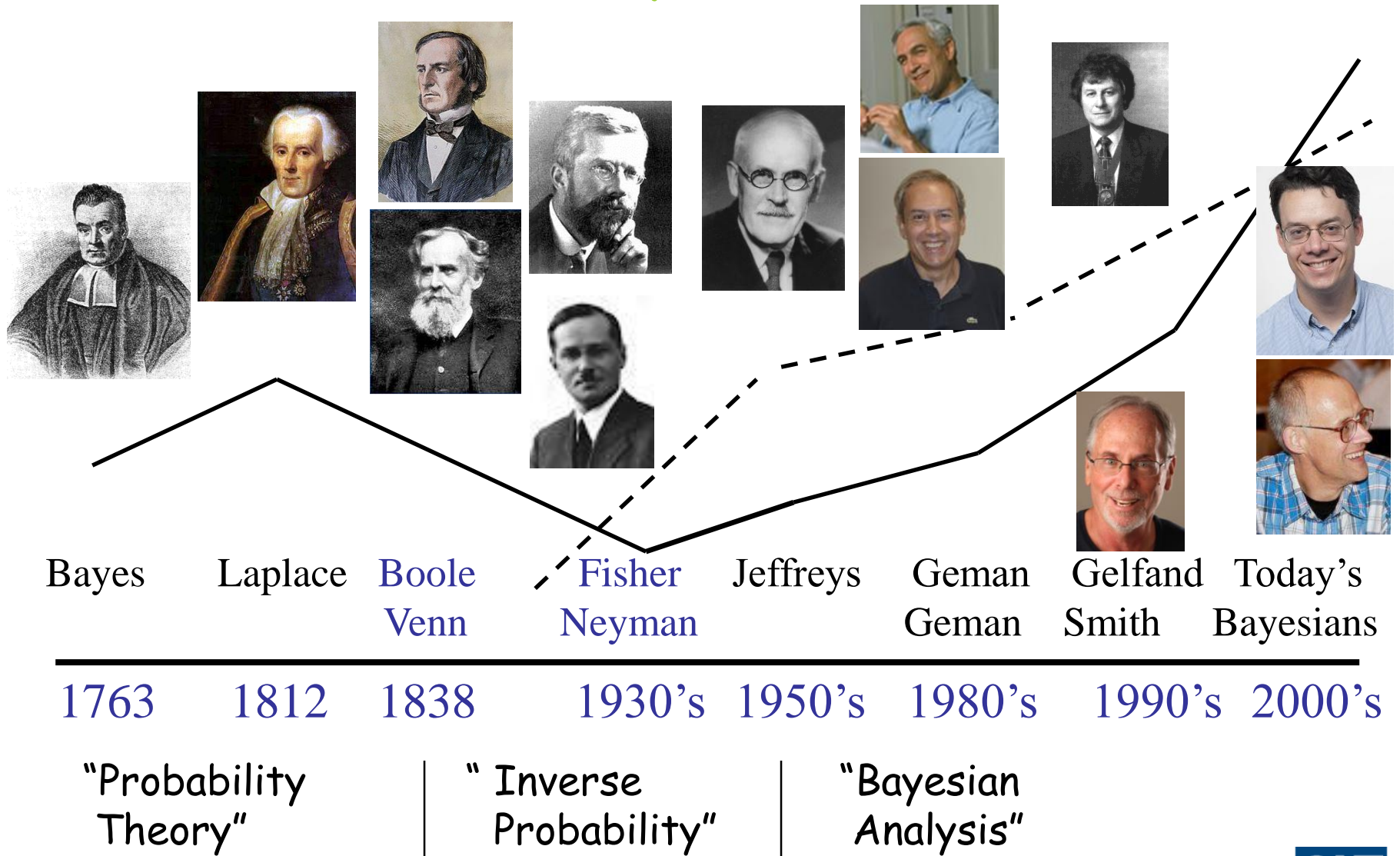
- Think differently about estimating and interpreting unknown parameters  
“what are possible values of this parameter?”
- Combine prior information with the data  
“what else do I know about this parameter and model?”
- Describe many sources of uncertainty in the model  
“how sure am I about the inputs to my model?”
- Describe complex systems using hierarchical or multi-level models

# Why Bayes?

Bayesian computational methods (such as MCMC) allow us to:

- Use non-standard distributions
- Fit very complex models
- Obtain a very wide variety of estimates
- Make a very wide range of inferences, based directly on posterior probabilities

# Modern Bayesian Statistics



# Binomial example

*y = number of successes from n trials*

Unknown: probability of success  $\theta$

$$y|\theta \sim \text{Bin}(n, \theta)$$

$$p(y|\theta) \propto \theta^y (1 - \theta)^{n-y}$$

# Prior

We could use a point prior:

“the population probability can only be 0.1 or 0.3, with 70% and 30% chance respectively”.

That is:  $p(\theta=0.1) = 0.7$  and  $p(\theta=0.3) = 0.3$

We might not be this certain:

“ $\theta$  could be any value [between 0 and 1] but based on previous studies it is more likely to be around 0.1, and very unlikely to be larger than 0.8.”

*What is an appropriate prior distribution for  $\theta$ ?*

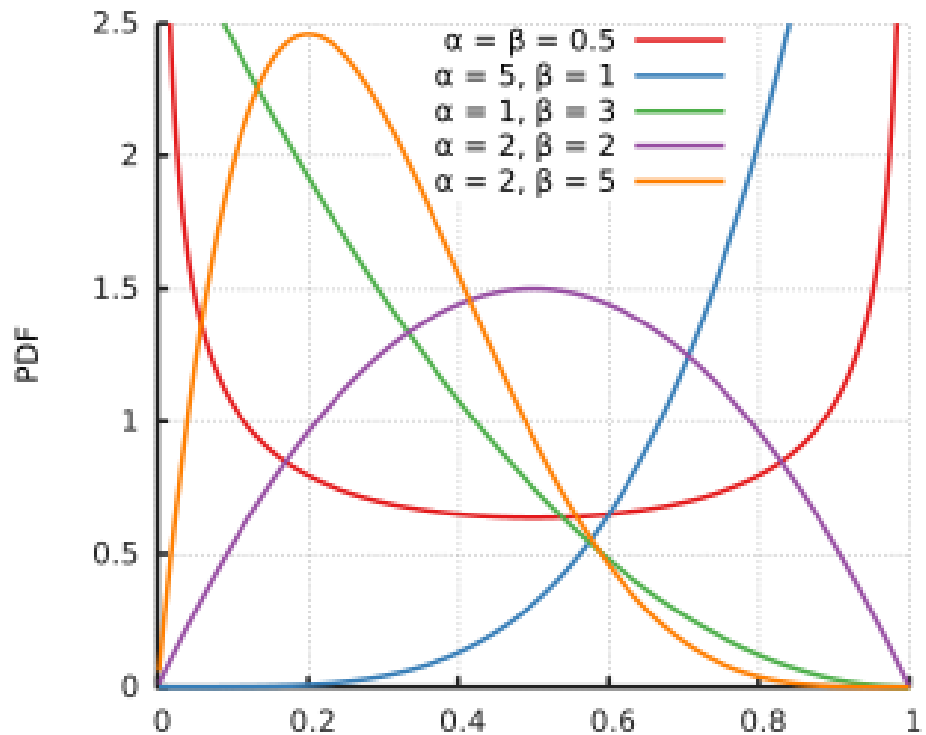
# Possible Prior

$$\theta \sim \text{Beta}(\alpha, \beta)$$

$$p(\theta) \propto \theta^{\alpha-1} (1-\theta)^{\beta-1}$$

$$E(\theta) = \frac{\alpha}{\alpha+\beta}$$

$$\text{Var}(\theta) = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$$

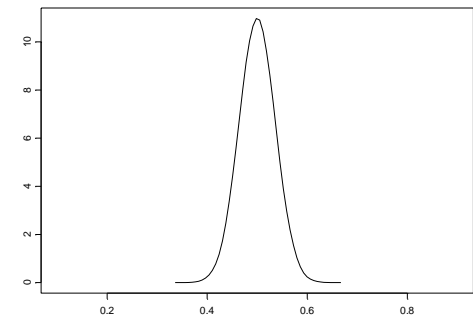
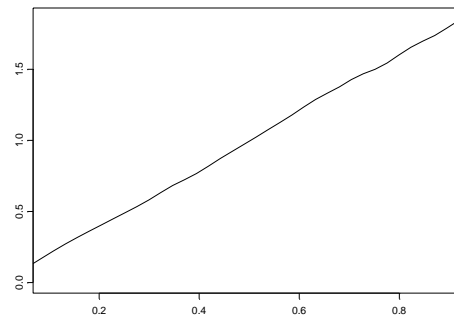
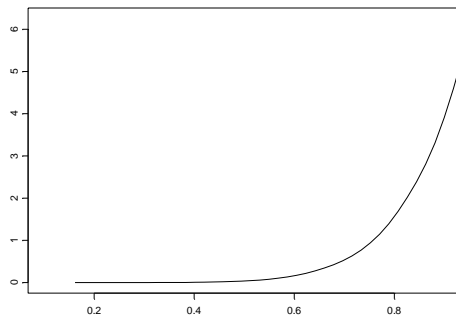
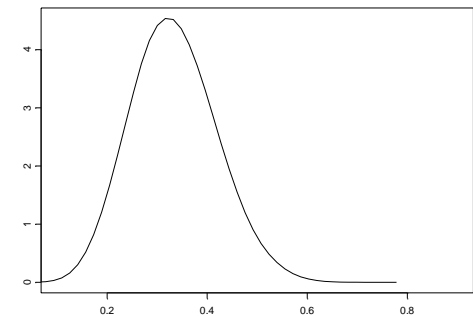
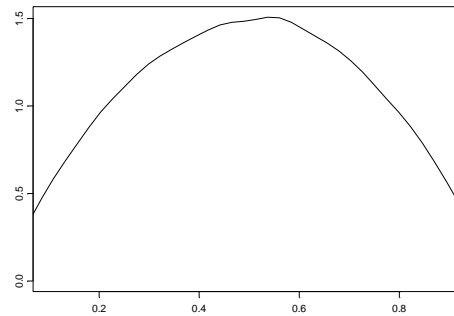
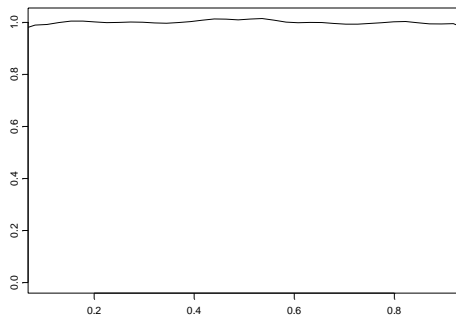


Wikipedia

# Beta Distribution

Match the plots to the distributions.

Beta(1,1), Beta(2,2), Beta(100,100),  
Beta(2,1), Beta(10,20), Beta(9,1)



# Posterior

$$p(\theta | y) \propto \textit{likelihood} \times \textit{prior}$$

$$= \theta^y (1 - \theta)^y \theta^{a-1} (1 - \theta)^{b-1}$$

$$= \theta^{y+a-1} (1 - \theta)^{n-y+b-1}$$

$$\theta|y \sim \text{Beta}(y + a, n - y + b)$$

# Your turn!

Binomial example with 22 successes, 7 failures:

Consider the following priors for  $\theta$ :

Beta(1,1)

Beta(9,1)

Beta(100,100)

*Choose one of these:*

- 1. What is the prior mean for  $\theta$ ?*
- 2. What is the posterior distribution for  $\theta$ ?*
- 3. What is the posterior mean for  $\theta$ ?*
- 4. What general conclusions can you make about the influence of priors and sample size?*

## Answers:

Sample proportion =  $22/29 = 0.76$

***Beta(1,1):***

Prior mean =  $(1)/(1+1) = 0.5$

Posterior mean =  $(22+1)/(22+1+7+1) = 0.74$

***Beta(9,1):***

Prior mean =  $(9)/(9+1) = 0.90$

Posterior mean =  $(22+9)/(22+9+7+1) = 0.79$

***Beta(100,100):***

Prior mean =  $(100)/(100+100) = 0.5$

Posterior mean =  $(22+100)/(22+100+7+100) = 0.53$

# Dynamic Updating

If we obtain more data, we do not have to redo all of the analysis: our posterior from the first analysis simply becomes our prior for this next analysis.

Binomial example:

Stage 0. Prior  $p(\theta) \sim \text{Beta}(1,1)$ ; ie  $E(\theta)=0.5$ .

Stage 1. Observe  $y=22$  successes from  $n=29$  trials.

Likelihood:  $p(y|\theta) \sim \text{Bin}(n=29, \theta)$ ;

Posterior:  $p(\theta|y) \sim \text{Beta}(23,8)$ ; ie  $E(\theta|y) = 0.74$

Stage 2: Observe 5 successes from 10 new trials.

Likelihood:  $p(y|\theta) \sim \text{Bin}(n=10, \theta)$ ;

Prior  $p(\theta) \sim \text{Beta}(23,8)$ ;

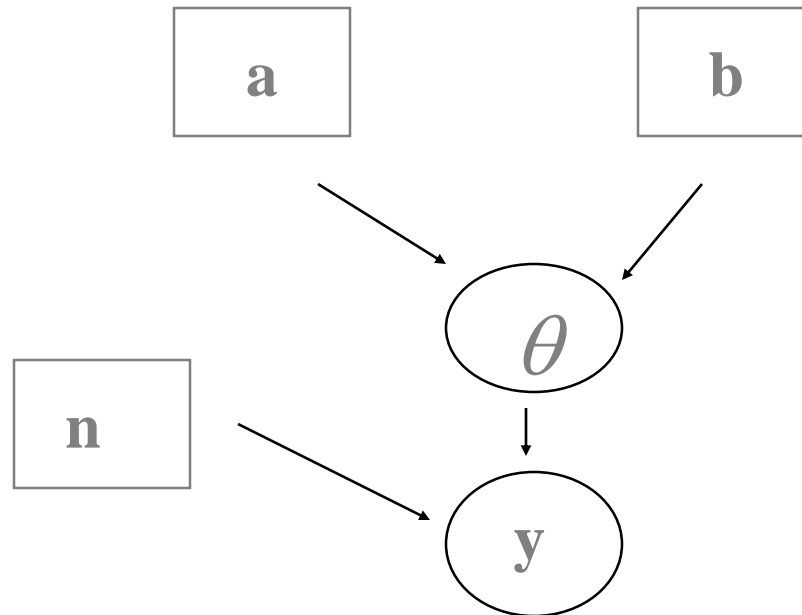
Posterior  $p(\theta|y) \sim \text{Beta}(28,13)$ ; ie  $E(\theta|y) = 0.68$ .

# DAG: Binomial model

## Model

$y \sim \text{Binomial}(\theta, n)$

$\theta \sim \text{Beta}(a, b)$



# Normal Model

Data  $D = (y_1, \dots, y_n)$ ;  $\theta = \mu$  unknown;  $\sigma_y^2$  known

- Likelihood:  $D|\mu, \sigma_y^2 \sim N(\mu, \sigma_y^2)$

- Conjugate prior:  $\mu \sim N(\mu_0, \sigma_0^2)$

- Posterior:  $\mu|D \sim N(\mu_n, \sigma_n^2)$

$$\mu_n = \left( \frac{\mu_0}{\sigma_0^2} + \frac{n\bar{y}}{\sigma_y^2} \right) / \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma_y^2} \right)$$

$$\frac{1}{\sigma_n^2} = \frac{1}{\sigma_0^2} + \frac{n}{\sigma_y^2}$$

# Normal Model cont.

Data  $D = (y_1, \dots, y_n)$ ;  $\theta = \mu$  unknown;  $\sigma_y^2$  known

- Prior predictive:
$$p(y') = \int p(y'|\mu, \sigma_y^2) p(\mu) d\mu$$
$$= \int N(y'|\mu, \sigma_y^2) N(\mu|\mu_0, \sigma_0^2) d\mu$$
$$y' \sim N(\mu_0, \sigma_0^2 + \sigma_y^2)$$
- Posterior predictive:  $p(y'|D) = \int p(y'|\mu, \sigma_y^2, D) p(\mu|D) d\mu$ 
$$= \int N(y'|\mu, \sigma_y^2) N(\mu|\mu_n, \sigma_n^2) d\mu$$
$$y' \sim N(\mu_n, \sigma_n^2 + \sigma_y^2)$$

Posterior predictive variance is the uncertainty due to the observation noise  $\sigma^2$  plus the uncertainty due to the parameters,  $\sigma_n^2$

# Your turn!

1. Suppose that we observe  $y = 2$  and wish to estimate the population mean  $\mu$ .
2. Assume  $p(y|\mu) \sim N(\mu, \sigma^2=3)$   
and our prior is  $p(\mu) \sim N(\mu_0=0, \sigma_0^2=1)$ .
3. What is the posterior distribution for  $\mu$ ?
4. What if the prior is  $N(2,1)$ ?  $N(0,10)$ ?
5. What happens to the comparative weight of the likelihood and prior as the same size increases? As the prior variance increases?

# Answers

- Observe  $y = 2$  ;  $p(y|\mu) \sim N(\mu, \sigma^2=3)$
- If prior  $p(\mu) \sim N(\mu_0=0, \sigma_0^2=1)$   
then the posterior is  $p(\mu|y) \sim N(\mu_1, \sigma_1^2)$   
posterior mean:  $\mu_1 = (0/1 + 2/3) / (1/1 + 1/3) = 0.50$   
posterior variance:  $1/\sigma_1^2 = 1/1 + 1/3 = 1.333$  so  $\sigma_1^2 = 0.75$
- If prior  $p(\mu) \sim N(\mu_0=2, \sigma_0^2=1)$   
then the posterior is  $p(\mu|y) \sim N(\mu_1, \sigma_1^2)$   
 $\mu_1 = (2/1 + 2/3) / (1/1 + 1/3) = 2$   
 $1/\sigma_1^2 = 1/1 + 1/3 = 1.333$  so  $\sigma_1^2 = 0.75$
- If prior  $p(\mu) \sim N(\mu_0=0, \sigma_0^2=10)$   
then the posterior is  $p(\mu|y) \sim N(\mu_1, \sigma_1^2)$   
 $\mu_1 = (0/10 + 2/3) / (1/10 + 1/3) = 1.54$   
 $1/\sigma_1^2 = 1/10 + 1/3 = 0.4433$  so  $\sigma_1^2 = 2.31$

# Dynamic updating – normal model

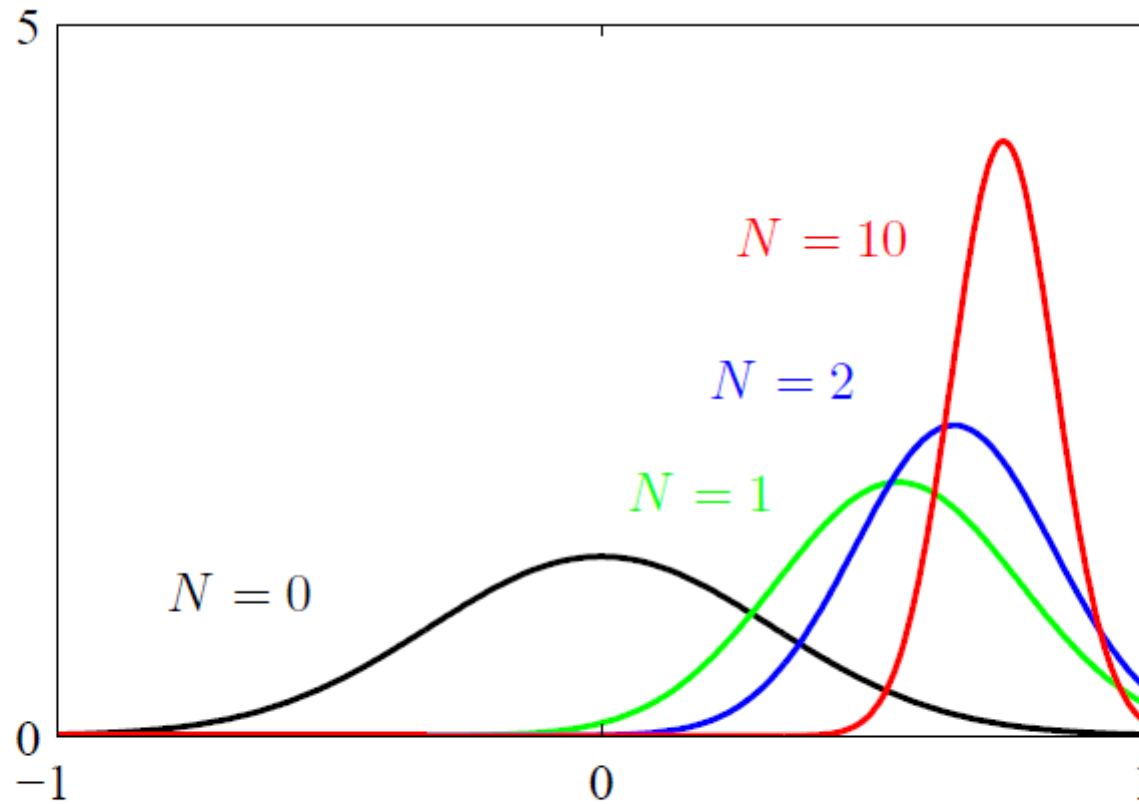


Figure 1: Sequentially updating a Gaussian mean starting with a prior centered on  $\mu_0 = 0$ . The true parameters are  $\mu^* = 0.8$  (unknown),  $(\sigma^2)^* = 0.1$  (known). Notice how the data quickly overwhelms the prior, and how the posterior becomes narrower. Source: Figure 2.12 [Bis06].

**Your turn!**

Draw a DAG for the Normal model

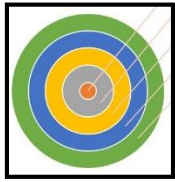
# Hierarchical models

- Useful when observations or parameters have a natural structure.
- Can be used to describe more complex priors.
- Can simplify computational strategies.

$$y_j | \theta_j, \phi \sim P(y_j | \theta_j, \phi)$$

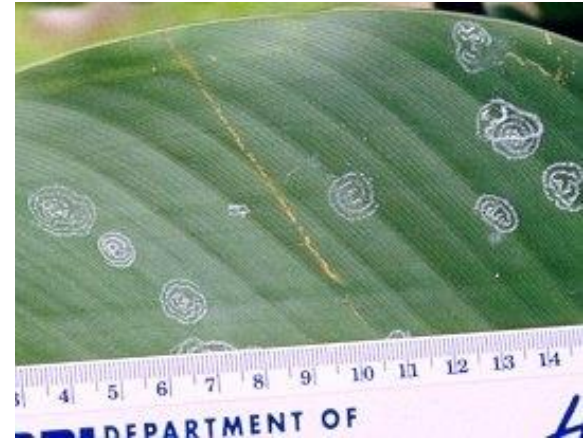
$$\theta_j | \phi \sim P(\theta_j | \phi)$$

$$\phi \sim P(\phi)$$



# Hierarchical models in practice: Spiralling whitefly

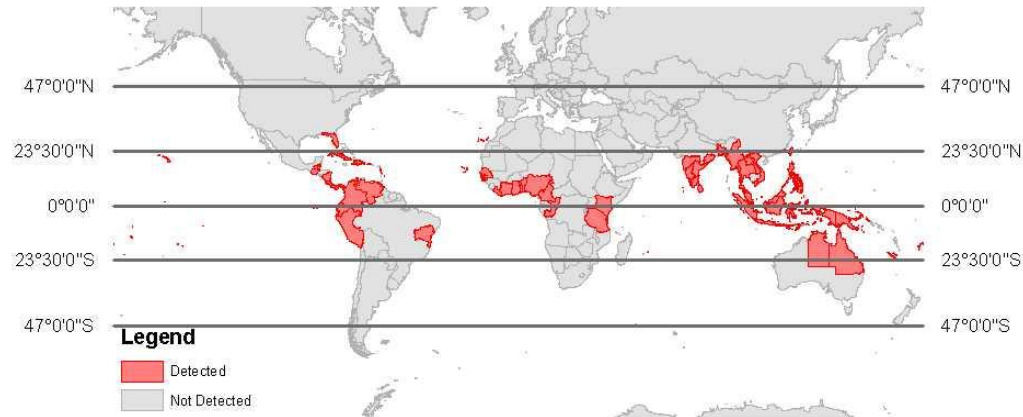
Stanaway et al.



- The Problem
  - Major tropical plant pest
  - Lives on 100 hosts +
  - Restricts market access to other states

- Information
  - Literature: Characteristics, growth, spread
  - Detectability (inspectors)
  - Surveillance data (> 30 000 records)

- Scope of modelling
  - Local, district and statewide



Countries where spiralling whitefly has been detected. Administrative regions within some countries are shown when documented. Source (CABI 2004, Monteiro et al. 2005, CABI 2006). Personal communications (J.H. Martin, 2008, B.M. Waterhouse, 2008)

# Hierarchical Bayesian model

- **Data Model:**  $\Pr(\text{data} \mid \text{incursion process and data parameters})$ 
  - How data is observed given underlying pest extent
- **Process Model:**  $\Pr(\text{incursion process} \mid \text{process parameters})$ 
  - Potential extent given epidemiology / ecology
- **Parameter Model:**  $\Pr(\text{data and process parameters})$ 
  - Prior distribution to describe uncertainty in detectability, exposure, growth ...
- **The posterior distribution of the incursion process (and parameters) is related to the prior distribution and data by:**

$$\Pr(\text{process, parameters} \mid \text{data}) \propto$$

$$\Pr(\text{data} \mid \text{process, parameters}) \Pr(\text{process} \mid \text{parameters}) \Pr(\text{parameters})$$

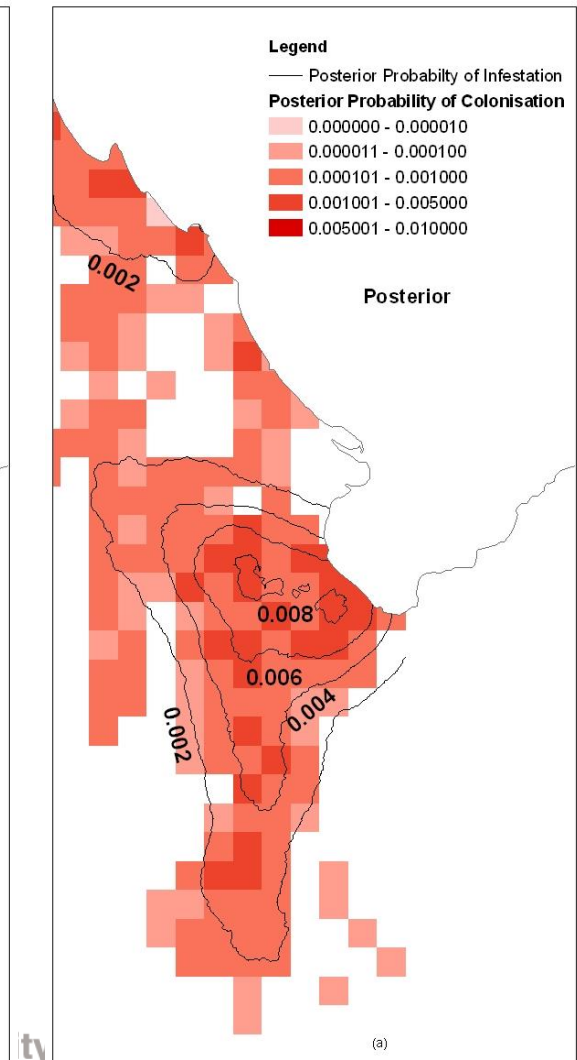
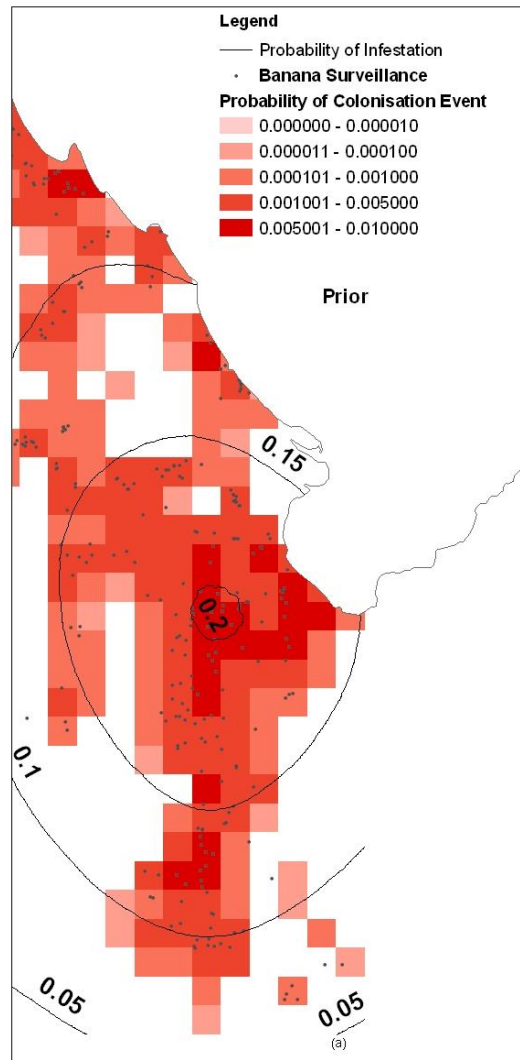
# Early Warning Surveillance

## Priors

## Surveillance data

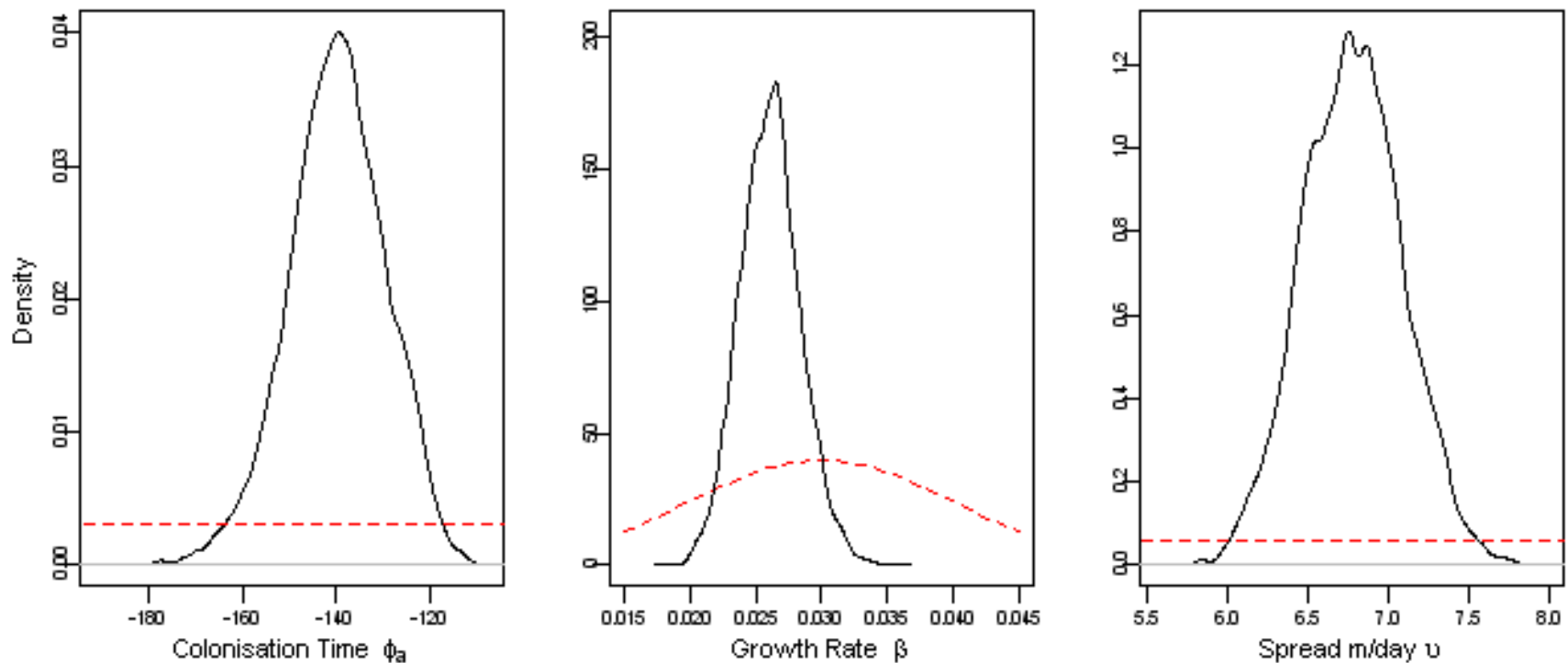
## Posterior learning

- modest reduction in area freedom
- large reduction in estimated extent
- residual “risk” maps to target surveillance



# Invasion Parameter Estimates

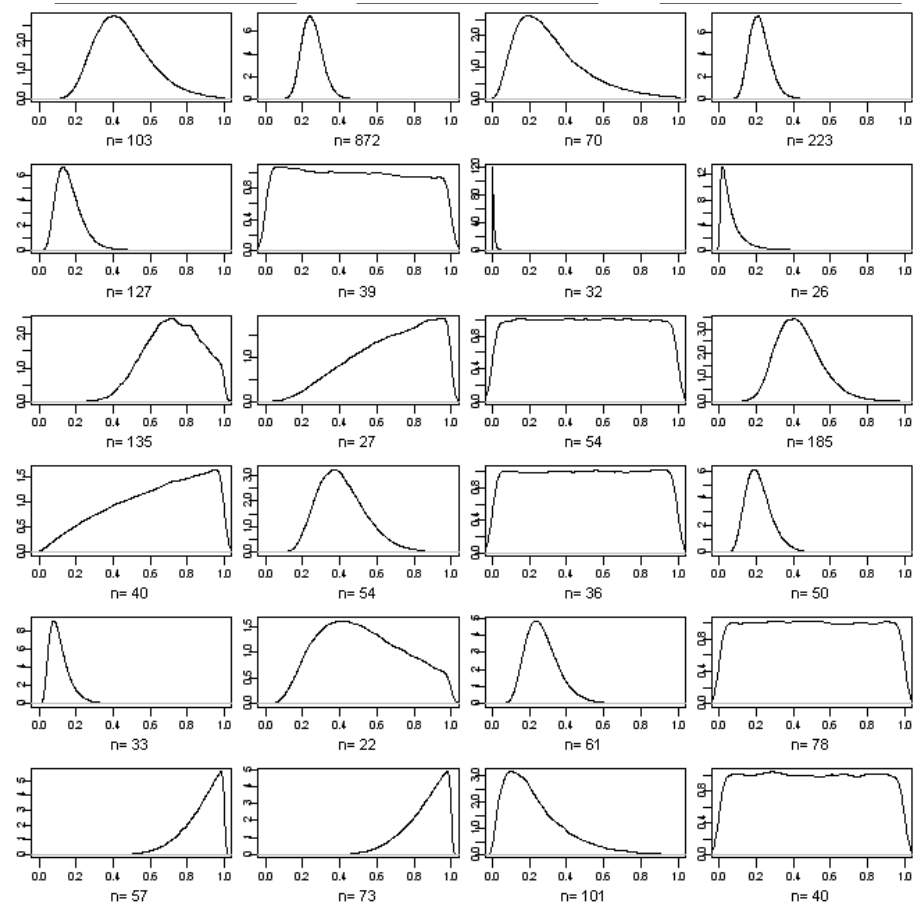
Useful for local management



# Observation parameter estimates

Also learn about:

- Host suitability
- Inspector efficiency



# Let's talk about Priors

- Conjugate priors
- Uninformative (objective, default)
- Weakly informative (vague)
- Informative

# Conjugate priors

<u>Family</u>	<u>Conjugate Prior</u>
Binomial( $N, \theta$ )	$\theta \sim \text{beta}(\alpha, \lambda)$
Poisson( $\theta$ )	$\theta \sim \text{gamma}(\delta_0, \gamma_0)$
$N(\mu, \sigma^2)$ , $\sigma^2$ known	$\mu \sim N(\mu_0, \sigma_0^2)$
$N(\mu, \sigma^2)$ , $\mu$ known, $\tau = 1/\sigma^2$	$\tau \sim \text{gamma}(\delta_0, \gamma_0)$
$\text{gamma}(\alpha, \lambda)$ , $\alpha$ known	$\lambda \sim \text{gamma}(\delta_0, \gamma_0)$
Beta( $\alpha, \lambda$ ), $\lambda$ known	$\alpha \sim \text{gamma}(\delta_0, \gamma_0)$

# What did Bayes say?

**Bayes:** “... it is plain, that in the case of such an event as I there call M, from the number of times it happens and fails in a certain number of trials, without knowing any thing more concerning it, one may give a guess whereabouts it’s probability is, and, by the usual methods computing the magnitudes of the areas there mentioned, see the chance that the guess is right. And that the same rule is the proper one to be used in the case of an event concerning the probability of which we absolutely know nothing antecedently to any trials made concerning it, seems to appear from the following consideration; viz. that concerning such an event I have no reason to think that, in a certain number of trials, it should rather happen any one possible number of times than another.”

**Price, on Bayes’ essay:** “... the rule must be to suppose the chance the same that it should lie between any two equidifferent degrees”

**Criticism:** not invariant under transformation, eg if  $p$  is uniform in a binomial setup,  $p^2$  is not uniform.

# Plausible “objective” priors

Berger (1985); Ghosh (2011)

- |                       |  |                |
|-----------------------|--|----------------|
| 1. Bayes-Laplace      | $\theta \sim \text{Uniform}$             | Beta(1,1)      |
| 2. Jeffreys/reference | $\theta \propto \sqrt{\det I(\theta)}$   | Beta(1/2, 1/2) |
| 3. Zellner            | $\theta \propto \theta^0 (1 - \theta)^0$ |                |

# Jeffreys' Prior

Example 1:  $y|\mu \sim N(\mu, \sigma^2)$ ,  $\sigma^2$  known

*What is the Jeffreys prior for  $\mu$ ?*

$$p(y|\mu) = \sqrt{2\pi\sigma^2} \exp(-(x - \mu)^2) / 2\sigma^2)$$

$$\begin{aligned} p(\mu) \propto \sqrt{I(\mu)} &= \sqrt{E \left[ \left( \frac{d}{d\mu} \log f(y|\mu) \right)^2 \right]} \\ &= \sqrt{E \left[ \left( \frac{y-\mu}{\sigma} \right)^2 \right]} = \sqrt{\int f(y|\mu) \left( \frac{y-\mu}{\sigma^2} \right)^2 dy} = 1 \end{aligned}$$

Improper, invariant

# Jeffreys' Prior

Example 2:

$$y|\sigma \sim N(\mu, \sigma^2), \mu \text{ known}$$

Jeffreys' prior:

$$p(\sigma) \propto 1/\sigma$$

Improper, invariant

Example 3: Your turn!

$$y|\theta \sim \text{Bernoulli}(\theta), y \in (0,1), 0 \leq \theta \leq 1$$

Jeffreys' prior?

# Jeffreys' Prior: Binomial

$$y|\theta \sim \text{Bernoulli}(\theta) \propto \theta^y (1 - \theta)^{1-y}$$

$$\begin{aligned} p(\theta) &\propto \sqrt{I(\theta)} = \sqrt{E \left[ \frac{y}{\theta} - \frac{(1-y)}{(1-\theta)} \right]^2} \\ &= \sqrt{\theta \left[ \frac{1}{\theta} - \frac{0}{(1-\theta)} \right]^2 + (1-\theta) \left[ \frac{0}{\theta} - \frac{1}{(1-\theta)} \right]^2} \\ &= \frac{1}{\sqrt{[\theta(1-\theta)]}} \end{aligned}$$

arcsine distribution, Beta(1/2, 1/2)

# Other priors

- Weakly Informative

$$\mu \sim N(0, \text{large variance})$$

- Informative (subjective)

$$\mu \sim N(M, V)$$

# Normal model, unknown mean unknown variance

Conjugate:

$$\sigma^{-2} \sim \text{Gamma}(\alpha, \beta)$$

Weakly/strongly informative:

$$\sigma \sim \text{Uniform}(a, b)$$

# Normal linear regression

$$y = (y_1, \dots, y_n)^T ; X = n \times k \text{ design matrix}$$

$$y_i = x_i^T \beta + \varepsilon_i ; \varepsilon_i \sim N(0, \sigma^2)$$

or equivalently

$$y_i | \beta, \sigma^2 \sim N(\mu_i, \sigma^2)$$

$$\mu_i = x_i^T \beta$$

# Conjugate prior

$$y_i | \beta, \sigma^2 \sim N(\mu_i, \sigma^2)$$

$$\mu_i = x_i^T \beta$$

Conjugate priors:

$$p(\beta, \sigma^2) = p(\beta | \sigma^2) p(\sigma^2)$$

$$p(\beta | \sigma^2) \sim N(\beta_0, \sigma^2 \Lambda_0^{-1})$$

$$p(\sigma^2) \sim IG(a_0 = \frac{\nu_0}{2}, b_0 = \frac{\nu_0 s_0^2}{2})$$

# Posterior under conjugate prior

$$p(\beta, \sigma^2 | y, X) = p(y | \beta, \sigma^2) p(\beta | \sigma^2) p(\sigma^2)$$

$$p(\beta | \sigma^2, y, X) \sim N(\mu_1, \sigma^2 \Lambda_1^{-1})$$

$$\mu_1 = (X^T X + \Lambda_0)^{-1} (\Lambda_0 \mu_0 + X^T y)$$

$$\Lambda_1 = (X^T X + \Lambda_0)$$

$$p(\sigma^2 | y, X) \sim IG(a_1, b_1)$$

$$a_1 = a_0 + n/2$$

$$b_1 = b_0 + 0.5(y^T y - \mu_1^T \Lambda_0 \mu_1 + \mu_0^T \Lambda_0 \mu_0)$$

# Zellner's g-prior

- Objective prior for  $\beta$ : MVN with covariance matrix proportional to the inverse Fisher information matrix for  $\beta$ .
- The scalar constant  $g$  controls the weight assigned to the prior.

Prior:

$$\beta | \varphi \sim MVN(\beta_0, g\varphi^{-1}(X^T X)^{-1})$$

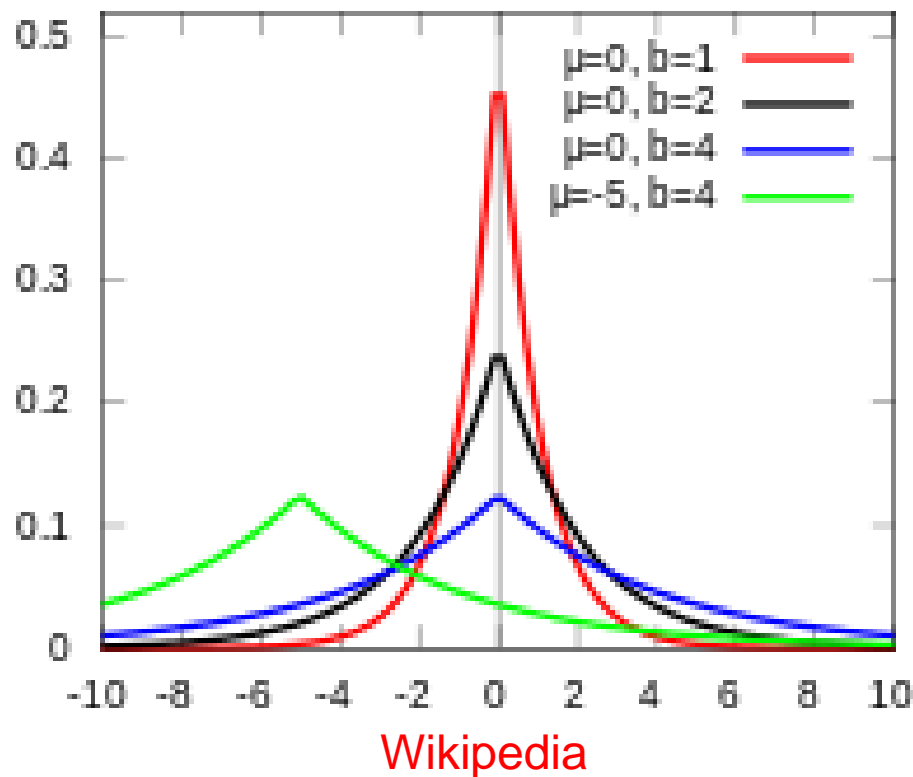
Posterior:

$$\beta | \varphi, x, y \sim MVN(q\hat{\beta} + (1 - q)\beta_0, g/\varphi(X^T X)^{-1})$$

$$q = g/(g + 1); \quad \hat{\beta} = (X^T X)^{-1}X^T y$$

# High dimensional regression

- Spike and slab priors
- Lasso regression (Laplace priors)
- Elastic net (combine lasso and ridge regression)



# Model Comparison

- Bayes factors, posterior odds

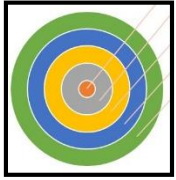
$$\frac{p(M_2|y)}{p(M_1|y)} \propto \frac{p(y|M_2) p(M_2)}{p(y|M_1) p(M_1)}$$

Bayes Factor  
(marginal likelihood ratio,  
integrated over parameters)

- BIC, DIC

$$\text{BIC} = \log P(y|\theta^*, M) - p/2 \log n$$

- Posterior predictive fit
- Reversible jump MCMC, Birth and death MCMC
- Model averaging

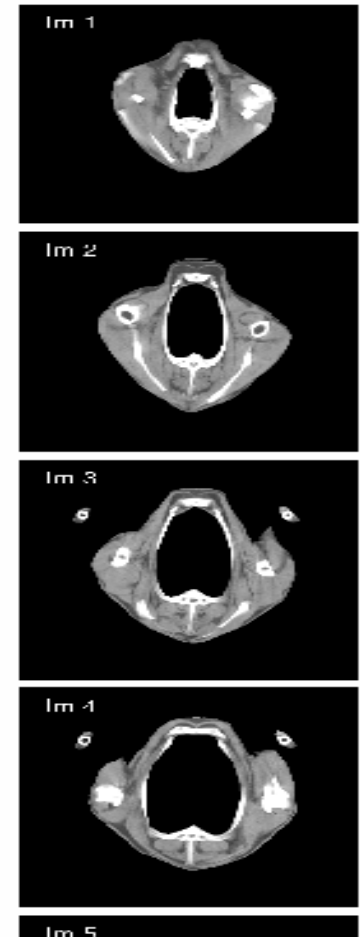


# Priors in practice: CAT-scanning sheep

with Alston, Robert,  
Besag et al.



$$p(y|\theta) = \sum_{k=1}^K w_k p(y|\theta_k)$$



# Bayesian Mixture Models

Robert

$$\text{Likelihood: } p(y|\mu, \sigma^2) = \sum_{k=1}^K w_k N(\mu_k, \sigma_k^2)$$

Priors:  $\mu_k \sim \text{Normal}$ ;  $w_k \sim \text{Dirichlet}$

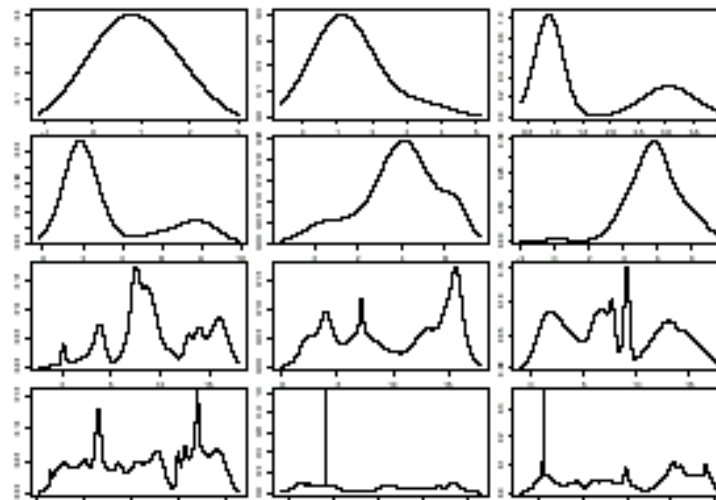
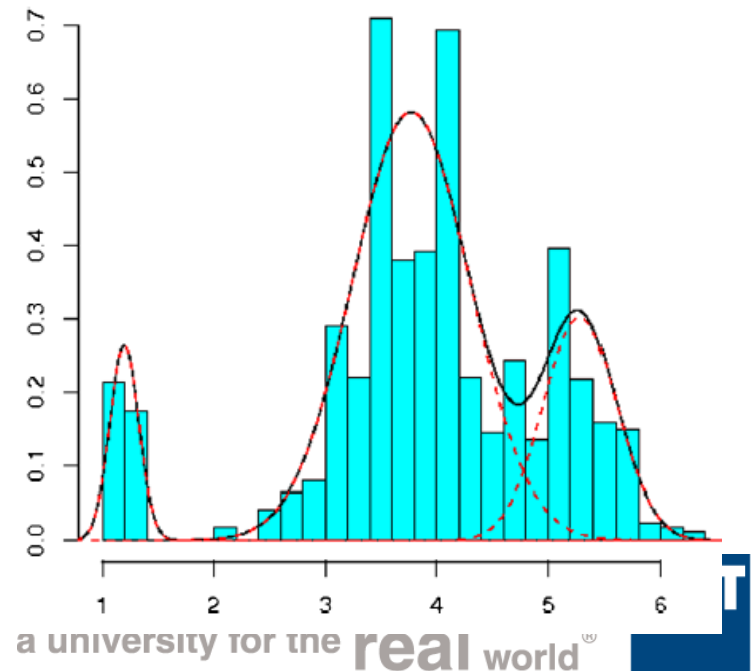
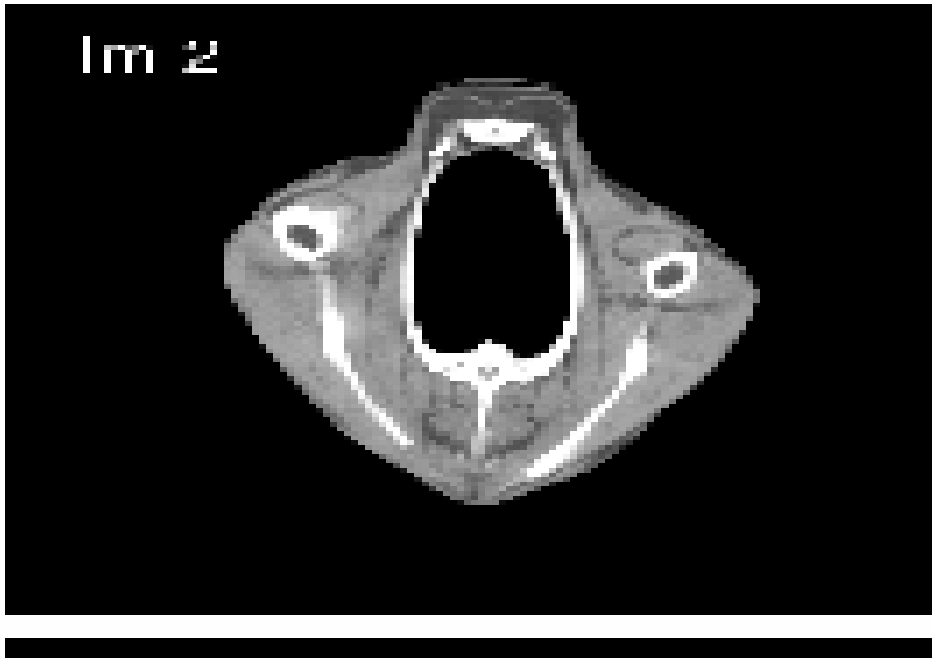


FIGURE 1. Some normal mixture densities for  $K = 2$  (first row),  $K = 5$  (second row),  $K = 25$  (third row) and  $K = 50$  (last row).

# Bayesian Mixture Models

Latent variable approach: ‘break down’ the likelihood

$$y_i | T_i = T \sim N(y | \mu_{T_i}, \sigma_{T_i}^2)$$



# Bayesian Spatial Models

Conditional autoregressive prior:

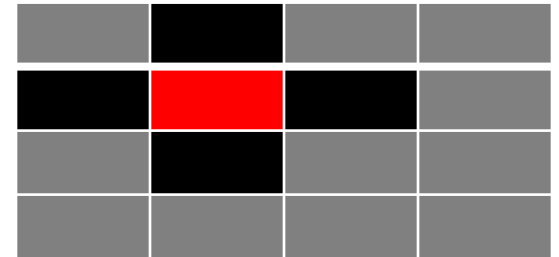
$$y_i \sim \text{Poisson}(\mu_i)$$

$$\mu_i = E_i \theta_i$$

$$\theta_i = \exp(X\beta + v_i + \varepsilon_i)$$

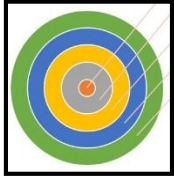
$$\varepsilon_i \sim N(0, \sigma_\varepsilon^2)$$

$$v_i | v_j \sim N(\mu_i, \sigma_i^2)$$



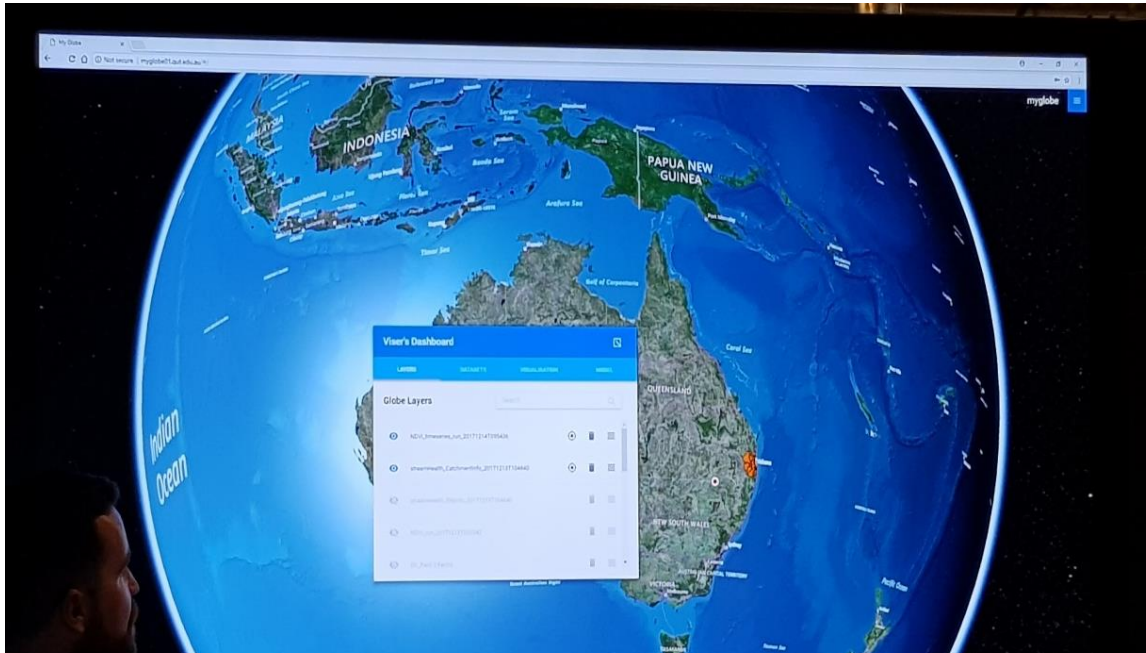
$\mu_i$  = weighted sum of neighbouring  $n$ 's / no. neighbours

$\sigma_i^2 = \sigma^2$  / no. neighbours



# From sheep to cancer

with Baade, Duncan  
Cramb et al.



<https://www.youtube.com/watch?v=sleXZ7EiSj8>

<https://atlas.cancer.org.au/>

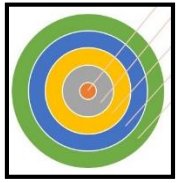


Does “place” impact on cancer?

What is the effect of agricultural/environmental exposures?

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# From cancer to ICU: predicting death after heart surgery

with Petra  
Graham

- Demographic (age, gender, hospital length of stay, etc)
- Comorbidity (other diseases)
- Physiological (heart rate, blood pressure, body temperature, etc)
- Biochemical (white cell count, potassium, sodium, etc)
- Treatment (use of mechanical ventilation, etc)
- Outcome (dead/alive after 30 days)



# Adaptive logistic regression model

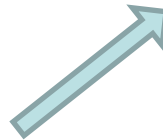
$$y_i \in (0,1); \quad i = 1, \dots, n$$

$$y_i \sim \text{Bernoulli}(p_i)$$

$$\text{logit}(p_i) = \log(p_i/(1 - p_i)) = X\beta + e_i$$

$$e_i \sim N(0, \sigma_e^2)$$

$$\beta \sim N(\beta_0, \sigma_0^2)$$



U.S.A. ICU risk score

# “Doing” Bayesian Analysis: Markov chain Monte Carlo

- “Decompose” joint posterior distribution into a sequence of conditional distributions – these are often much simpler (eg, simple univariate normals, etc)
- Simulate from each conditional distribution in turn. We use a simulation method that resembles a Markov chain (so that the new simulated value relies only on the previous value), giving a set of simulated values
$$\theta^{(1)}, \theta^{(2)}, \dots, \theta^{(i)}, \dots$$
which converges to the required conditional,  
The resulting simulations will come from the required joint distribution.
- We can use Markov chain theory to make statements about behaviour and convergence of the chain

# Markov chain Monte Carlo

- “Decompose” joint posterior distribution into a sequence of conditional distributions – these are often much simpler (eg, simple univariate normals, etc)
- Simulate from each conditional distribution in turn. We use a simulation method that resembles a Markov chain (so that the new simulated value relies only on the previous value), giving a set of simulated values
$$\theta^{(1)}, \theta^{(2)}, \dots, \theta^{(i)}, \dots$$
which converges to the required conditional,  
Under mild assumptions, the resulting simulations will come from the required joint distribution.
- Use Markov chain theory to make statements about behaviour and convergence of the chain

# MCMC Algorithms

- **Gibbs sampling**: sample from the conditionals themselves
- **Metropolis-Hastings**: sample from an “easy” distribution and accept those values that conform to the conditional distribution
- Lots of variations: reversible jump, slice sampling, particle filters, perfect sampling, adaptive rejection sampling, etc
- Need to ensure conditions, eg *detailed balance*, *reversibility*

# Gibbs sampling

Suppose you have a joint posterior  $p(\theta_1, \theta_2 | y, \dots)$

0. Choose starting values  $\theta_1^{(0)}, \theta_2^{(0)}$

1. At  $i$ th iteration  $(i)$

Sample  $\theta_1^{(i)}$  from  $p(\theta_1 \mid \theta_2, y, \dots)$

Sample  $\theta_2^{(i)}$  from  $p(\theta_2 \mid \theta_1, y, \dots)$

2. Repeat step 1 many times

3. Make inferences based on simulated values

# Example of Gibbs Sampler

Consider a single observation  $(y_1, y_2)$  from a bivariate normal population with unknown mean  $(\theta_1, \theta_2)$  and known variance-covariance matrix

$$\begin{pmatrix} \sigma_1^2 & \rho^2 \\ \rho^2 & \sigma_2^2 \end{pmatrix}$$

Let  $\sigma_1^2 = \sigma_2^2 = 1$

With a uniform prior on  $\theta_1, \theta_2$ , the posterior distribution is

$$\begin{pmatrix} \theta_1 \\ \theta_2 \end{pmatrix} \mid y \sim N \left( \begin{pmatrix} y_1 \\ y_2 \end{pmatrix}, \begin{pmatrix} \sigma_1^2 & \rho^2 \\ \rho^2 & \sigma_2^2 \end{pmatrix} \right)$$

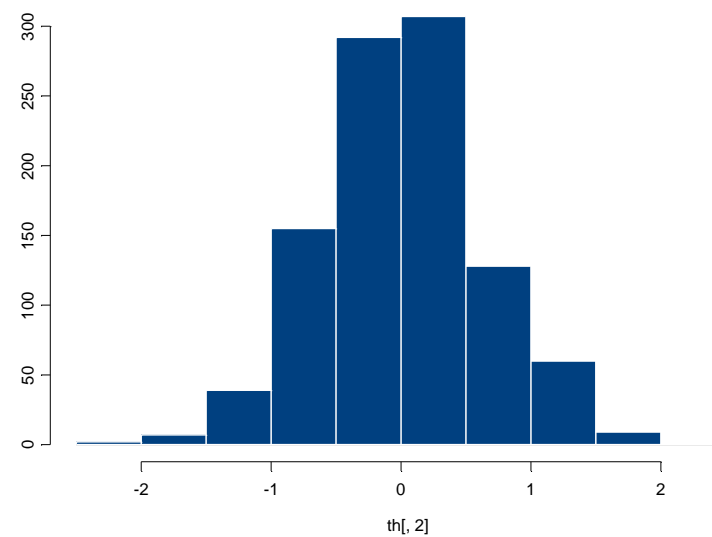
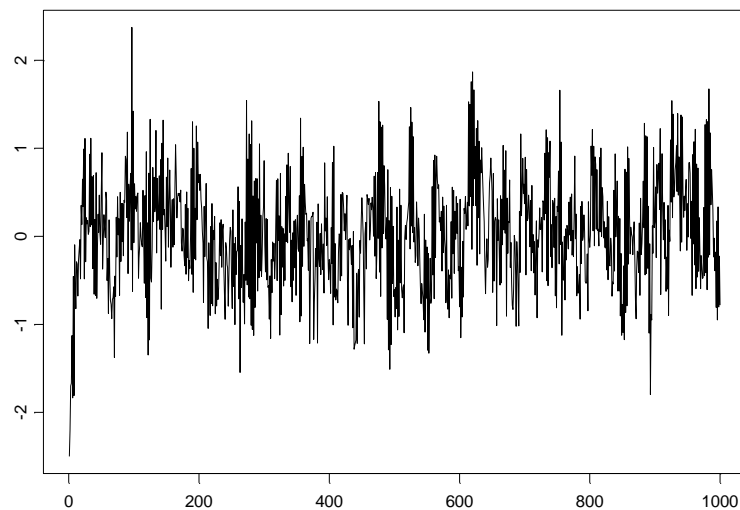
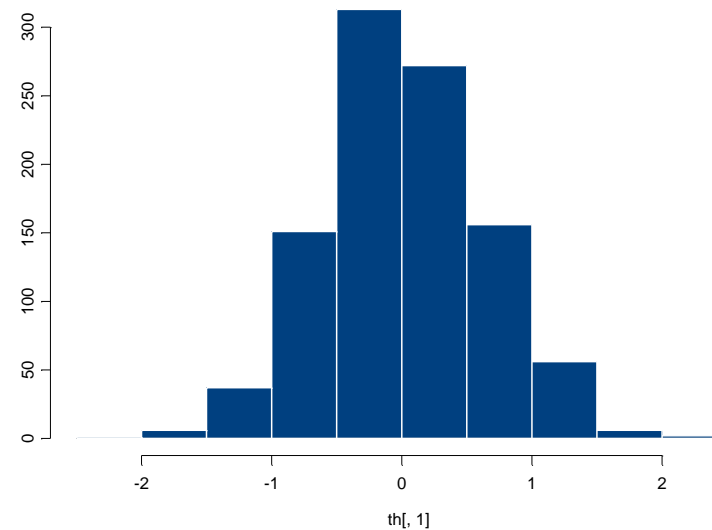
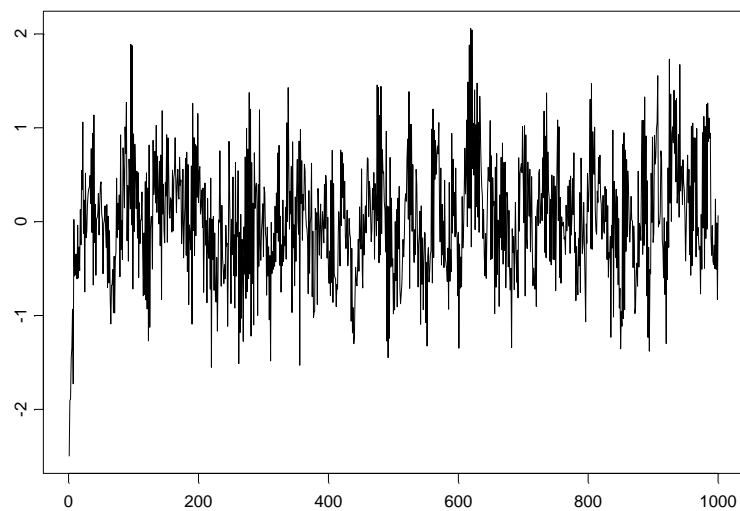
# Gibbs sampler

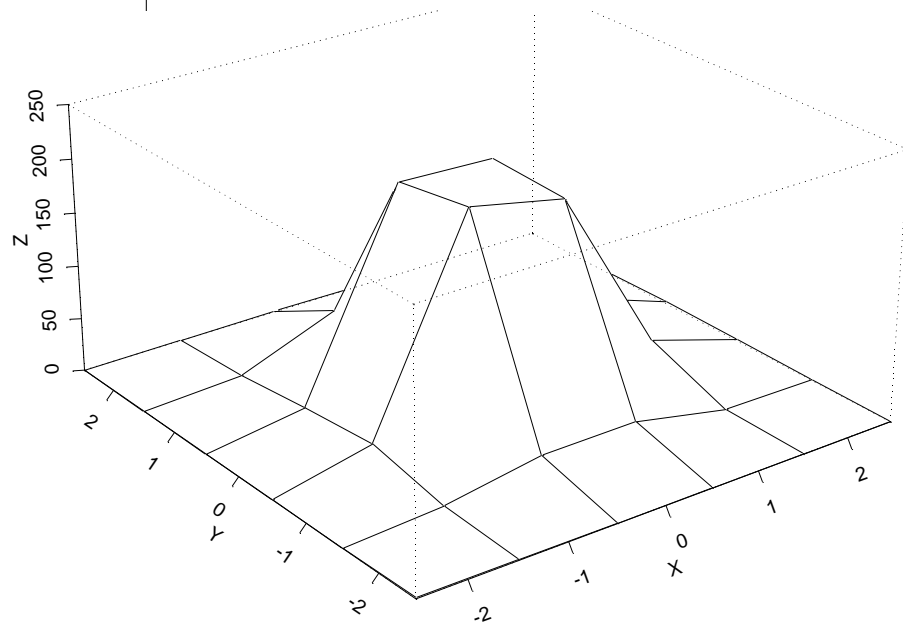
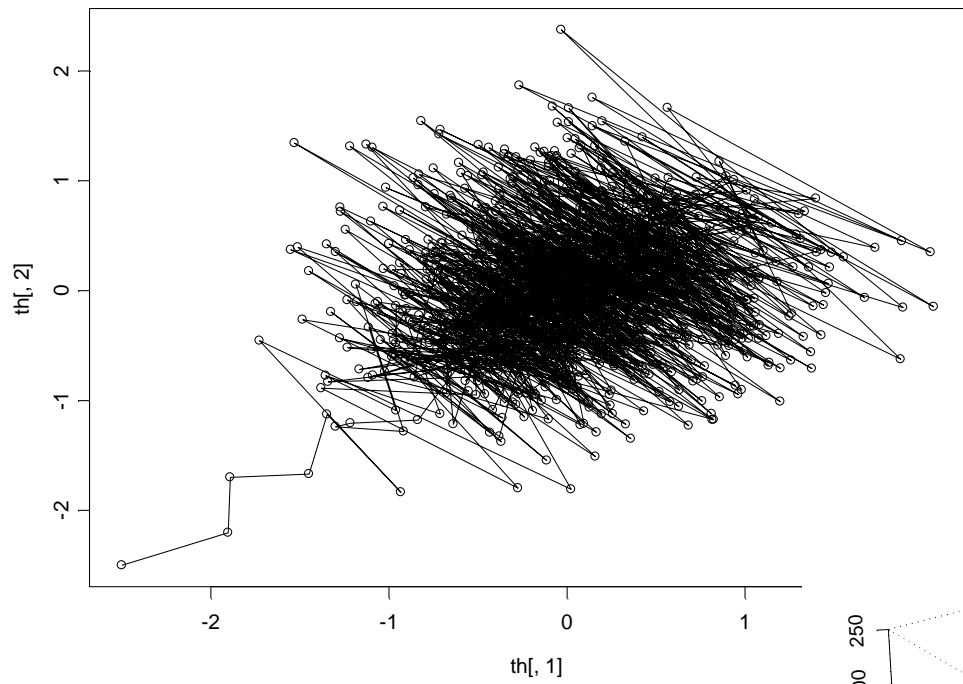
Sample from

$$\theta_1 \mid \theta_2, y \sim N(y_1 + \rho(\theta_2 - y_2), 1 - \rho^2)$$

Then sample from

$$\theta_2 \mid \theta_1, y \sim N(y_2 + \rho(\theta_1 - y_1), 1 - \rho^2)$$





# Gibbs sampling for mixtures

0. *Initialisation*: Choose  $\underline{p}^{(0)}$  and  $\underline{\theta}^{(0)}$  arbitrarily

*For*  $t=1, \dots$

1.1 Allocate observations to components:

Generate  $T^{(t)}$  for each observation

1.2 Generate new weights for the components:

Generate  $\underline{p}^{(t)}$

1.3 Generate new parameters for each component:

Generate  $\underline{\theta}^{(t)}$

# Other computational methods

- Other MCMC algorithms
  - Metropolis Hastings
  - Slice Sampling
- Approximations
  - Variational Bayes
  - INLA
  - ABC
  - SMC
  - Gaussian Processes

# Bayesian Software

**BUGS: Bayesian analysis using Gibbs sampling (OpenBUGS)**

<http://www.mrc-bsu.cam.ac.uk/software/bugs/>

<http://www.openbugs.net>

**JAGS: Just Another Gibbs Sampler (RJAGS)**

<http://mcmc-jags.sourceforge.net/>

<https://www.r-bloggers.com/getting-started-with-jags-rjags-and-bayesian-modelling/>

**STAN (interface with R, Matlab, etc)**

<http://mc-stan.org/>

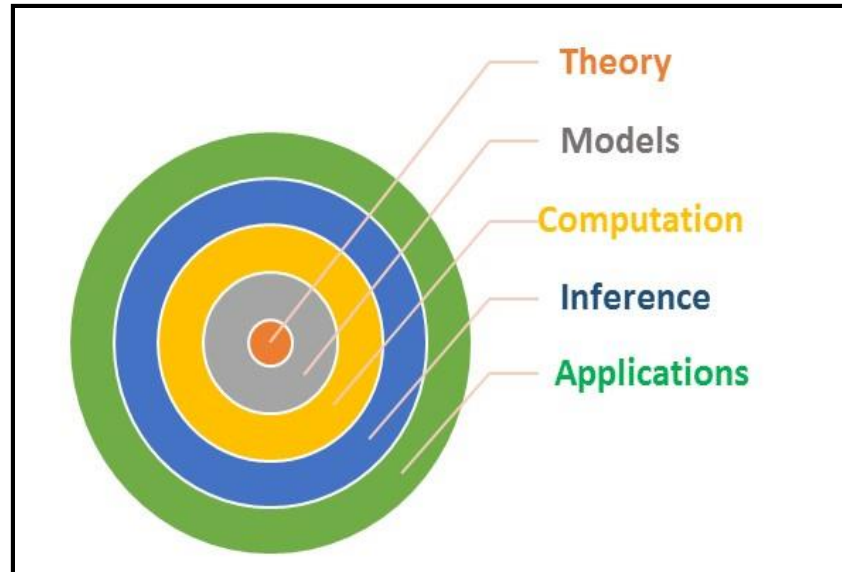
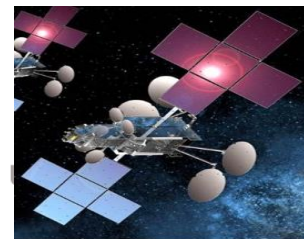
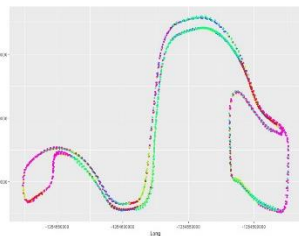
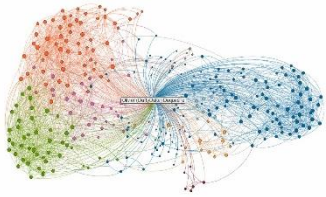
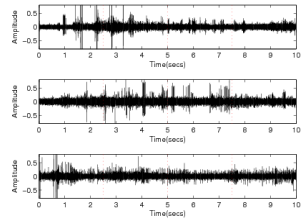
**INLA: integrated nested Laplace approximation**

[www.r-inla.org](http://www.r-inla.org)

**R packages**

<http://cran.r-project.org/web/views/Bayesian.html>

# Bayes + Big Data



age	sex
21.79603	female
21.71116	female
21.27584	male
27.37577	female
19.66598	male
24.37509	female

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# Meeting the challenge

## Models:

- Probabilistic
- Regularised
- Flexible
- Robust
- Transferable
- Adaptive

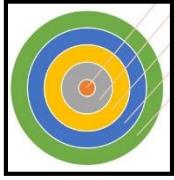
## Computation:

- Scalable  
(parallelisable)
- Subsampling
- Pre-computable
- Approximations  
(eg. ABC, SMC,  
VI, GP)

## Inference:

- Estimation
- Optimisation
- Uncertainty  
quantification
- Testing
- Model averaging

*"In the past ten years, it's hard to find anything that doesn't advocate a Bayesian approach." -Nate Silver*



# Monitoring the Great Barrier Reef

with Peterson et al.

**Crown of thorns starfish pest 'the  
greatest impact on Great Barrier  
Reef'**  
By Kirsty Nancarrow  
Posted Fri 14 Jun 2013

Jan 29 2014  
**Concerned scientists demand  
rejection of permit for dumping  
dredge spoil in Great Barrier  
Reef's waters**

**Cyclone damage devastating  
to the Great Barrier Reef**

Sep 2 2013  
**Thousands gather for Rally for the Reef**

The world's largest living  
structure  
Almost 3000 reefs  
Over 2,600 km



# Traditional data source: surveys



# MAP OF THE GREAT BARRIER REEF

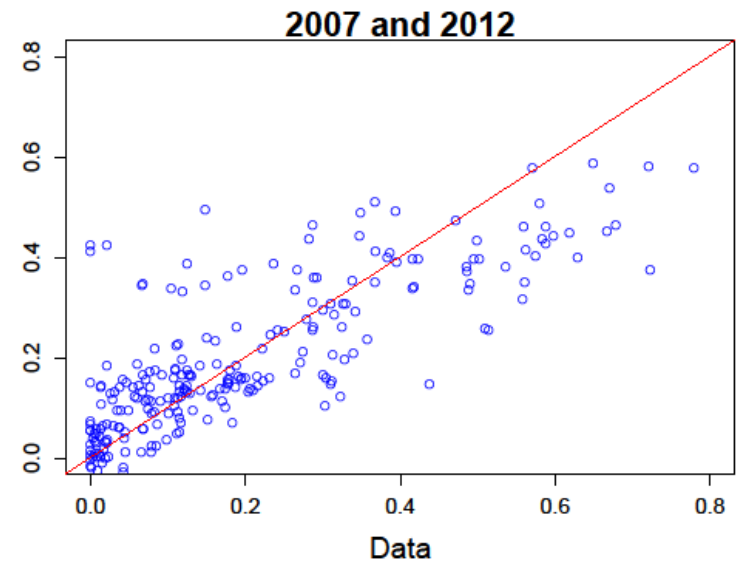
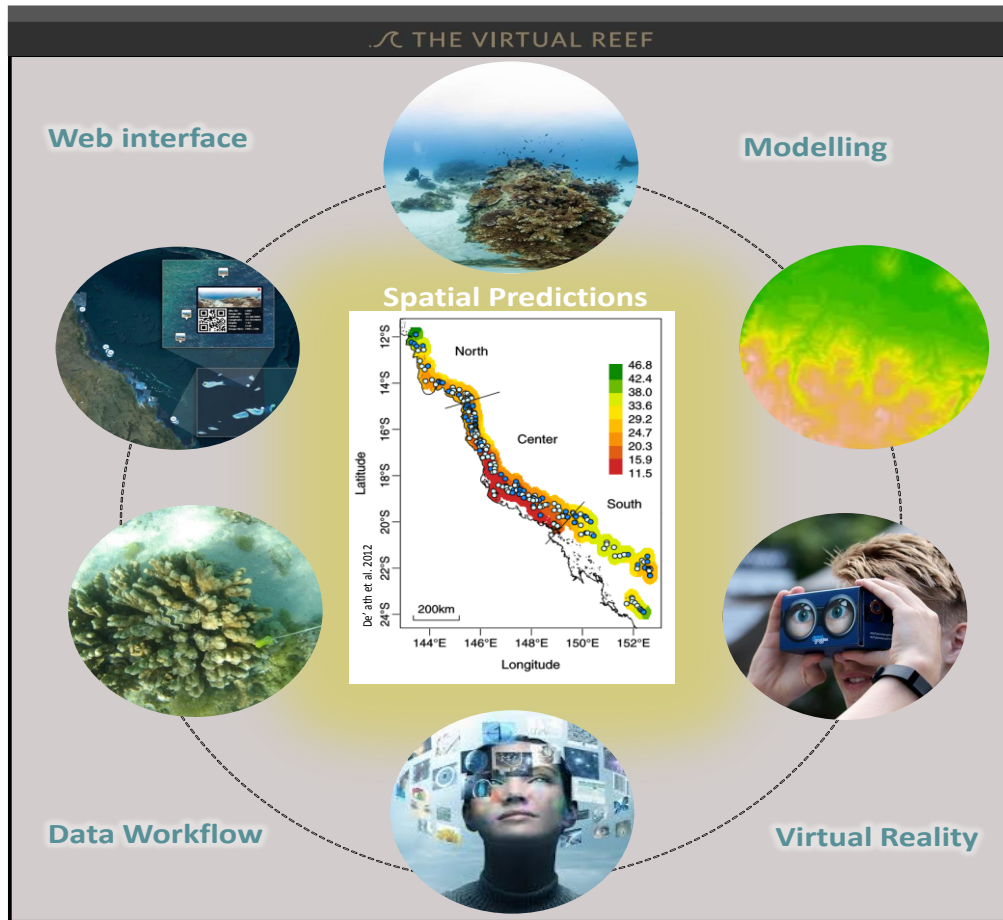


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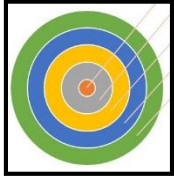
# Virtual Reef Diver



<https://www.virtualreef.org.au/>

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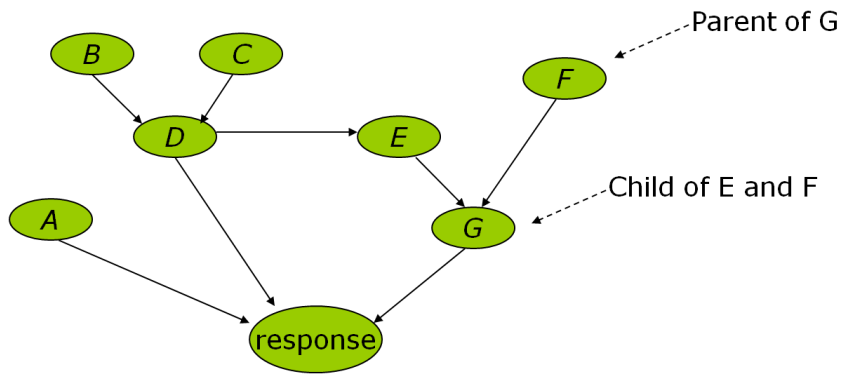


# Modelling complex systems

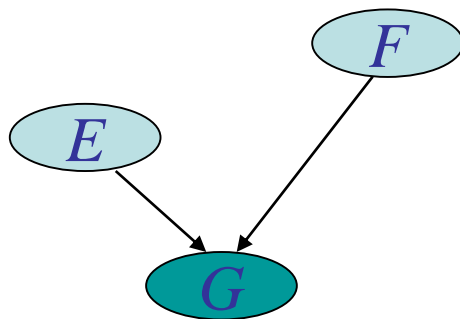
Environment  
Biology  
Social  
Govt  
Economic  
Human impact  
Unknowns  
External factors



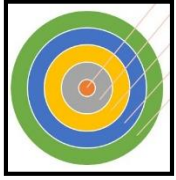
# Bayesian Networks



<i>F</i>	
<i>low</i>	0.7
<i>medium</i>	0.2
<i>high</i>	0.1



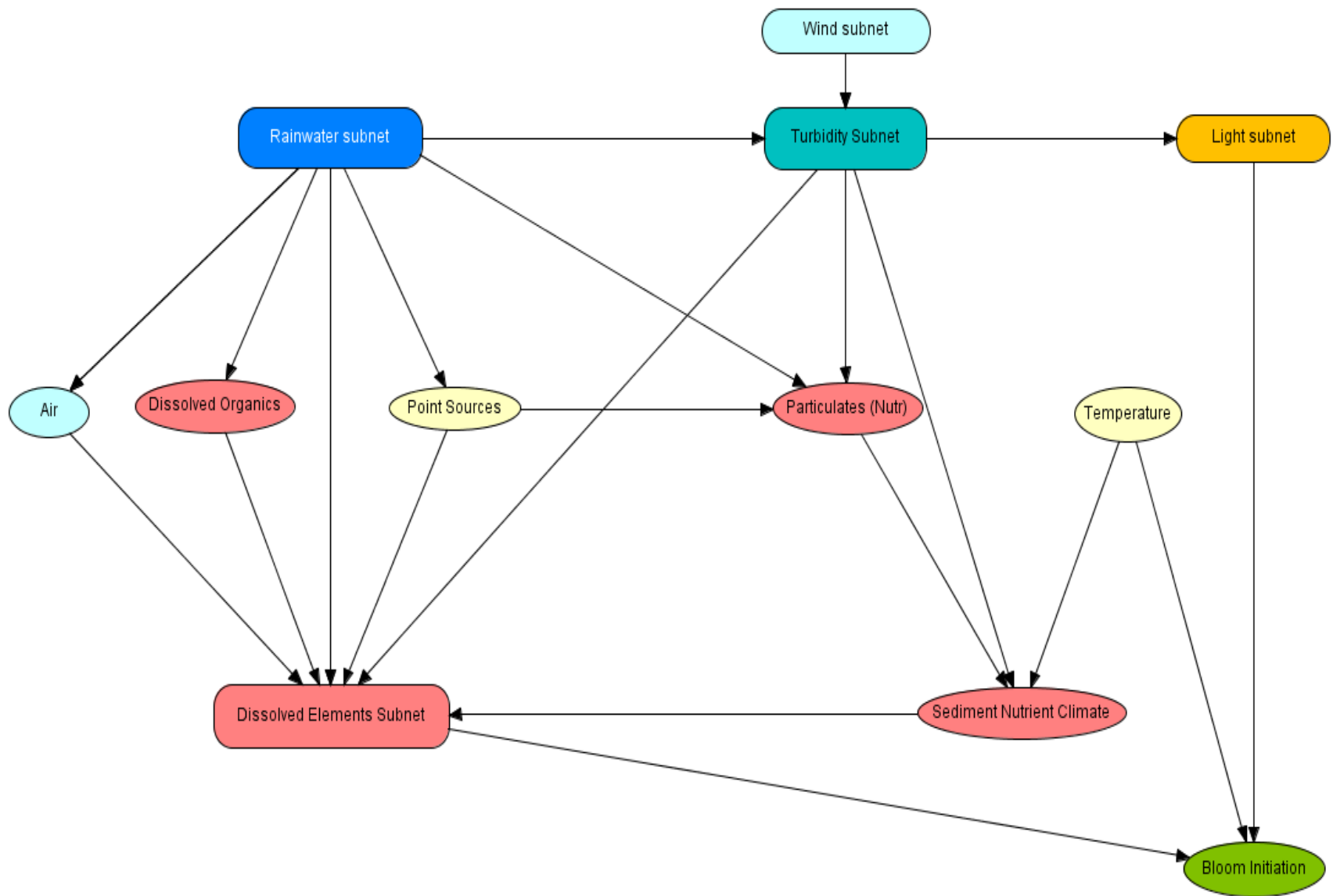
		<i>G</i>	
<i>E</i>	<i>F</i>	<i>normal</i>	<i>high</i>
<i>yes</i>	<i>low</i>	0.4	0.6
	<i>medium</i>	0.2	0.8
	<i>high</i>	0.1	0.9
<i>no</i>	<i>low</i>	0.5	0.5
	<i>medium</i>	0.6	0.4
	<i>high</i>	0.4	0.6



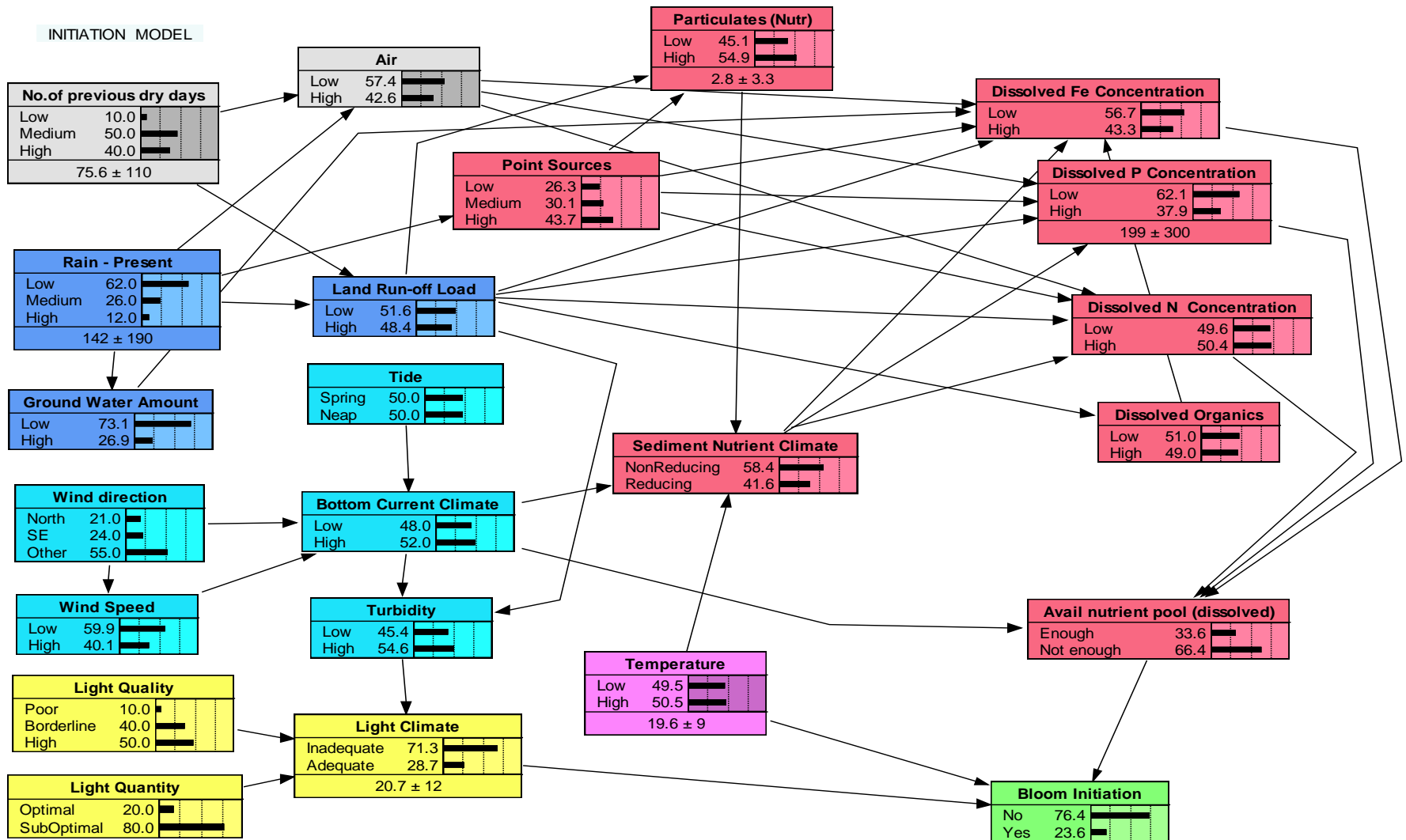
# Managing lyngbya in Moreton Bay, Australia

- *What is the overall scientific consensus about the drivers of lyngbya?*
- *What management actions should be taken to reduce lyngbya in the Bay?*





# INITIATION MODEL



# Most influential factors

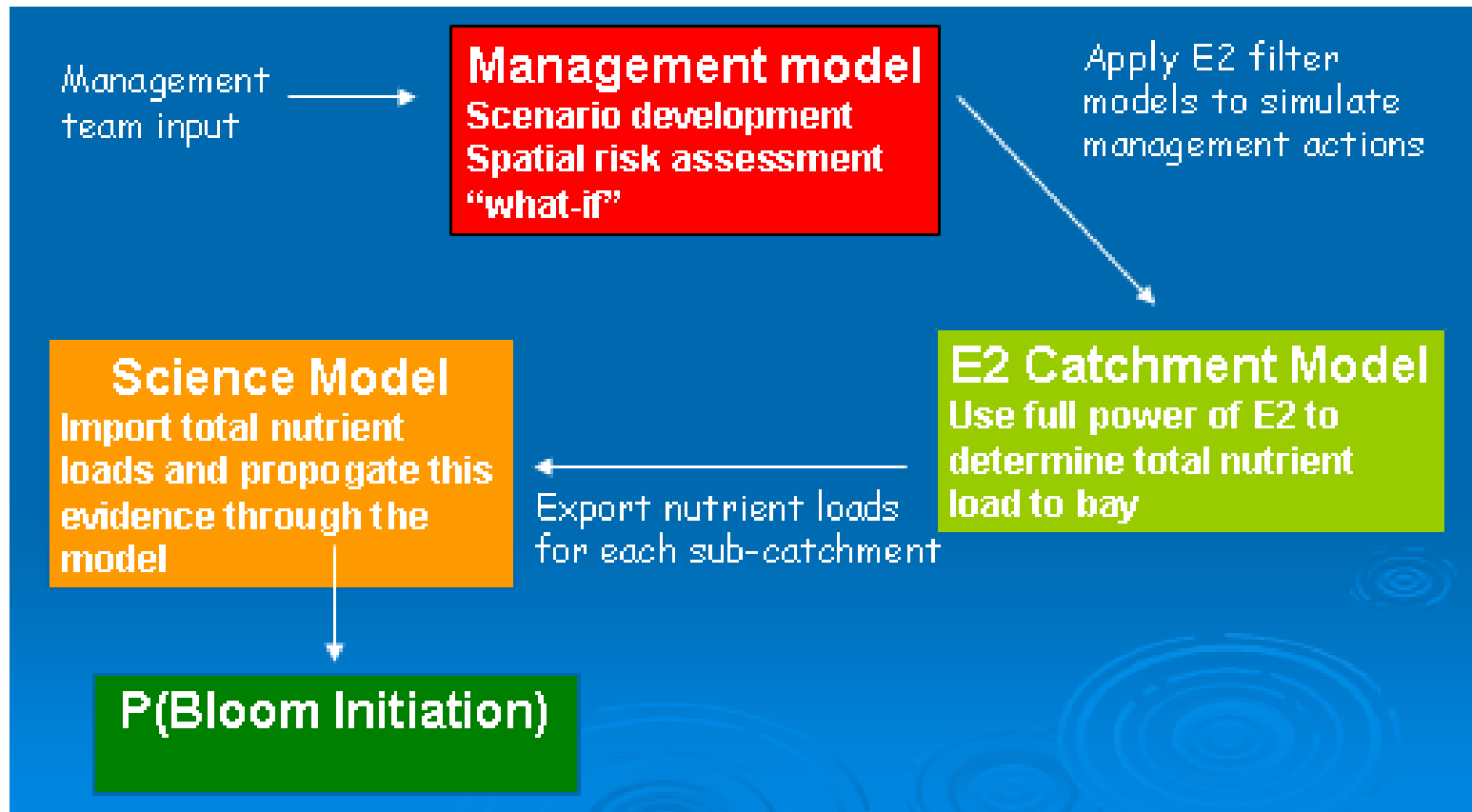
1. Available Nutrient Pool
2. Bottom Current Climate
3. Sediment Nutrients
4. Dissolved Iron
5. Dissolved Phosphorous
6. Light
7. Temperature

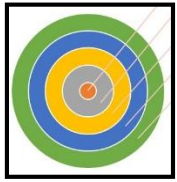


# “What-if” scenarios

Factor	Change in P(Bloom) (%)
Available Nutrient Pool	77 (3% - 80%)
Bottom Current Climate	28 (15% - 43%)
Sediment Nutrient Climate	17 (21% - 38%)
Dissolved Fe	16 (21% - 37%)
Dissolved P	15 (23% - 38%)
Light Climate	14 (18% - 32% )
Temperature	14 (21% - 35%)
Dissolved N	13 (22% - 35%)
Rain – present	10 (25% - 35%)
Light Quantity	9 (21% - 30%)

# From Science to Management





# Biosecurity: “Beyond Compliance”

An integrated approach to pest risk management

STDF – WTO funded project

5 SEA partners  
+ OC: + QUT



Mumford *et al.*

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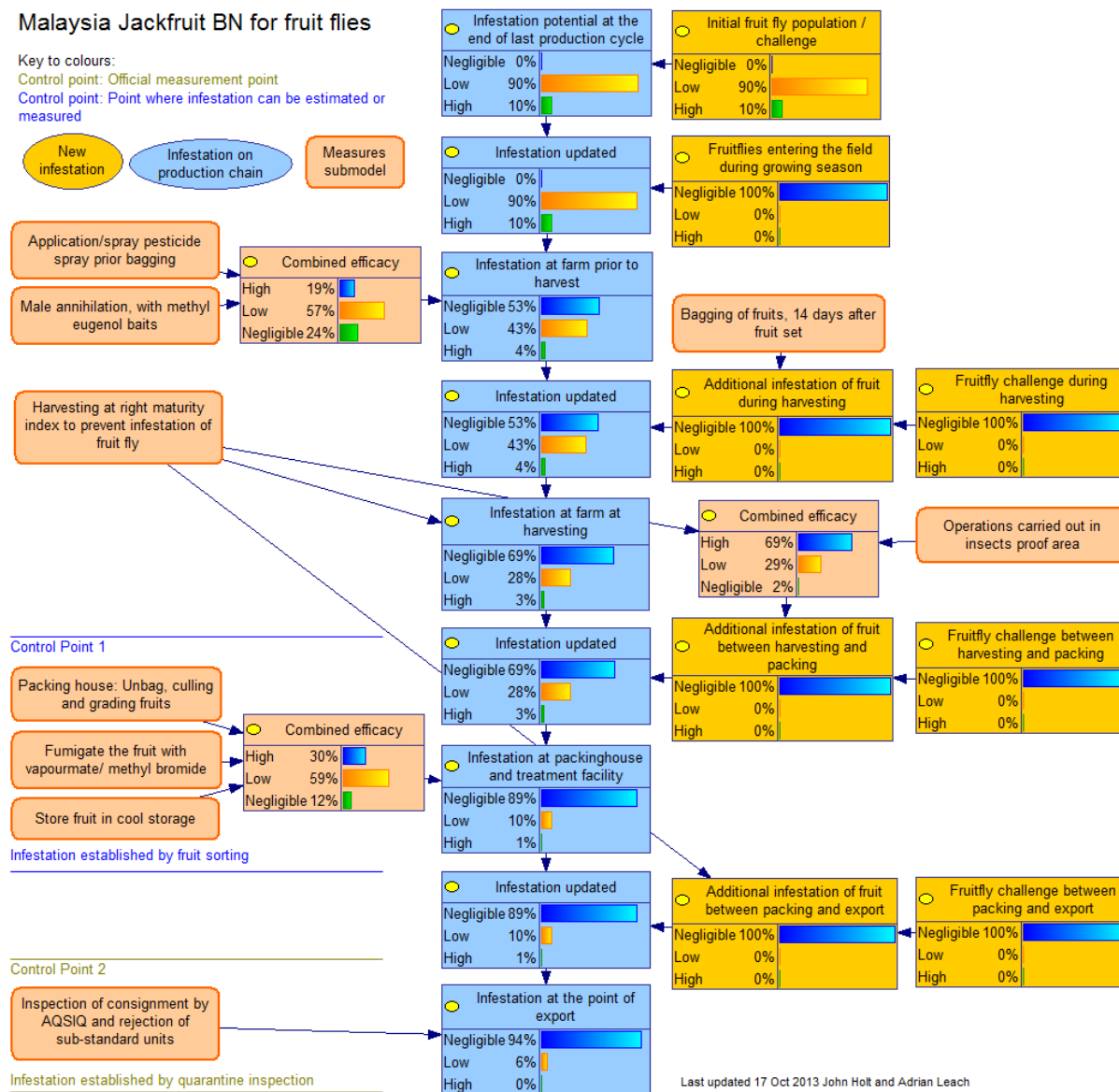
# CP-BN

## Malaysia Jackfruit BN for fruit flies

Key to colours:

Control point: Official measurement point

Control point: Point where infestation can be estimated or measured



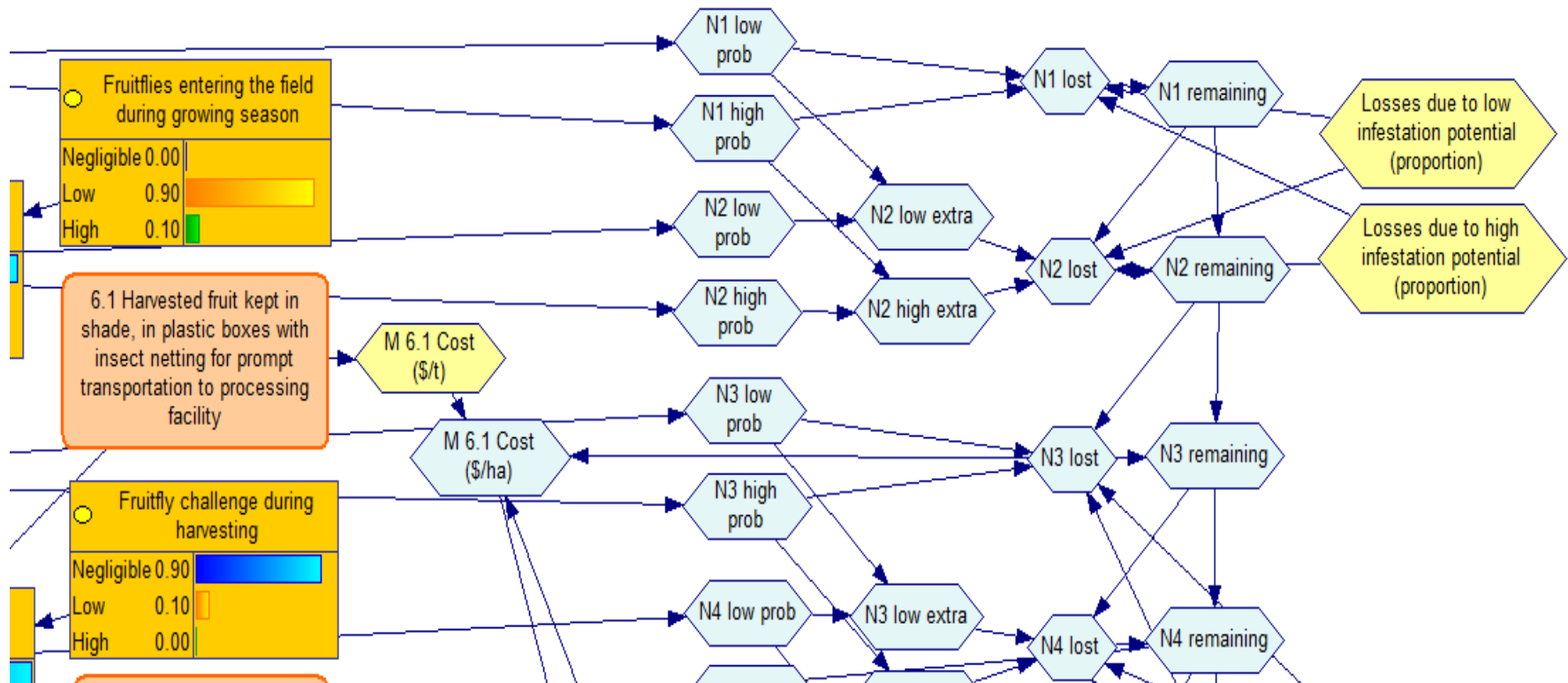
Last updated 17 Oct 2013 John Holt and Adrian Leach

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# Economics

## adding costs and losses utility nodes



J. Holt, A. W. Leach, S. Johnson, D. M. Tu, D. T. Nhu, N. T. Anh, L. N. Quang, M. M. Quinlan, P. J. L. Whittle, K. Mengersen and J. D. Mumford (in prep.)  
Bayesian networks to compare pest control interventions on commodities along agricultural production chains.

# Ongoing Questions

1. How to elicit information from experts?
2. How to combine information from multiple experts?
3. How to incorporate uncertainty into BNs?
4. How to combine BNs?

# ***Reflection and Final Comments***