

A two-locus model with interacting mutation rates and fluctuating selection

Peter Pfaffelhuber

University of Freiburg
Abteilung für Mathematische Stochastik
joint work with Franz Baumdicker, Elisabeth Huss

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- An individual is **first-order-fit** if it produces on average more offspring than others.
- An individual is **second-order-fit** if its offspring produce on average more offspring than others.
- Example (from Mao et al (1997)):
In *E. coli*, the fraction of cells lacking a certain DNA repair mechanism, is $\sim 10^{-5}$.
→ they carry a mutator allele
Treat a population with antibiotics. After four rounds of treatments, only mutators have survived.
- Conclusion: **Mutators can be second-order fit.**

Mini-review

Second-order selection in bacterial evolution: selection acting on mutation and recombination rates in the course of adaptation

Olivier Tenaillon*, François Taddei, Miroslav Radman, Ivan Matic

Inserm E9916, Faculté de médecine Necker Enfants Malades, Université Paris V, 156, rue de Vaugirard, 75730 Paris cedex 15, France

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Abstract – The increase in genetic variability of a population can be selected during adaptation, as demonstrated by the selection of mutator alleles. The dynamics of this phenomenon, named second-order selection, can result in an improved adaptability of bacteria through regulation of all facets of mutation and recombination processes. © 2001 Éditions scientifiques et médicales Elsevier SAS

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Adaptation through genetic time travel? Fluctuating selection can drive the evolution of bacterial transformation

Jan Engelstädter¹ and Danesh Moradigaravand²

¹School of Biological Sciences, The University of Queensland, Brisbane, Queensland 4072, Australia

²Institute of Biogeochemistry and Pollutant Dynamics, ETH Zurich, Zurich 8092, Switzerland

The model

There is an A -locus (types: ℓ, h) and a B -locus (types: 0,1)

The A -allele determines the mutation rate at the B -locus

		B -locus	
		0	1
A -locus	ℓ	$X_{\ell 0} \longleftrightarrow X_{\ell 1}$	
	h	$X_{h 0} \longleftrightarrow X_{h 1}$	

$$dX_{\ell 0} = \theta_{\ell}(X_{\ell 1} - X_{\ell 0})dt + \sqrt{X_{\ell 0}(1 - X_{\ell 0})}dW_{\ell 0},$$

$$dX_{\ell 1} = \theta_{\ell}(X_{\ell 0} - X_{\ell 1})dt + \sqrt{X_{\ell 1}(1 - X_{\ell 1})}dW_{\ell 1},$$

$$dX_{h 0} = \theta_h(X_{h 1} - X_{h 0})dt + \sqrt{X_{h 0}(1 - X_{h 0})}dW_{h 0},$$

$$dX_{h 1} = \theta_h(X_{h 1} - X_{h 0})dt + \sqrt{X_{h 1}(1 - X_{h 1})}dW_{h 1}$$

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Environment favors 1 over 0
with $\sigma n Z$ for $Z \in \{-1, 1\}$

$Z \rightarrow -Z$ at rate $\gamma n^2/2$

$$dX_{\ell 0} = -\sigma n Z X_{\ell 0} X_1 dt + \theta_{\ell} (X_{\ell 1} - X_{\ell 0}) dt + \sqrt{X_{\ell 0} (1 - X_{\ell 0})} dW_{\ell 0},$$

$$dX_{\ell 1} = \sigma n Z X_{\ell 1} X_0 dt + \theta_{\ell} (X_{\ell 0} - X_{\ell 1}) dt + \sqrt{X_{\ell 1} (1 - X_{\ell 1})} dW_{\ell 1},$$

$$dX_{h 0} = -\sigma n Z X_{h 0} X_1 dt + \theta_h (X_{h 1} - X_{h 0}) dt + \sqrt{X_{h 0} (1 - X_{h 0})} dW_{h 0},$$

$$dX_{h 1} = \sigma n Z X_{h 1} X_0 dt + \theta_h (X_{h 1} - X_{h 0}) dt + \sqrt{X_{h 1} (1 - X_{h 1})} dW_{h 1}$$

- The generator of (X, Z) is

$$Gf(x, z) = \underbrace{G_0 f(x, z)}_{\substack{\text{mutation,} \\ \text{resampling}}} + \underbrace{nG_1 f(x, z)}_{\text{selection}} + \underbrace{n^2 G_2 f(x, z)}_{\text{environment}}$$

- Any Markov process can be characterized via a martingale problem. Here, this means that

$$\begin{aligned} f(X_t, Z_t) - \int_0^t Gf(X_s, Z_s) ds \\ = f(X_t, Z_t) - \int_0^t (G_0 f + nG_1 f + n^2 G_2 f)(X_s, Z_s) ds \end{aligned}$$

is a martingale for all smooth, bounded f .

- Is there a limit of X as $n \rightarrow \infty$?

- Goal: Averaging out a fast variable (environment).
- Dates back at least to Khashminskii (1966)

A limit theorem for the solutions of differential equations with random right-hand sides. *Theor. Probability Appl.*, 11(11):390–406, 1966.

- General reference is Kurtz (1992)

Averaging for martingale problems and stochastic approximation. In *Applied stochastic analysis* (New Brunswick, NJ, 1991), volume 177 of *Lecture Notes in Control and Inform. Sci.*, 186–209. Springer, Berlin, 1992.

- For processes on three time-scales:

- Semigroup approach: See Theorem 1.7.6 of Ethier, Kurtz (1986)

Markov processes: Characterization and convergence. Wiley Series in Probability and Mathematical Statistics: Probability and Mathematical Statistics. John Wiley & Sons Inc., New York, 1986.

- Special case of diffusion operators E. Pardoux and A. Yu. Veretennikov. On Poisson equation and diffusion approximation 1 and 2. *Ann. Probab.*, 29:1061–1085, 2001 and 31:1166–1192, 2003.

- Martingale-problem approach Hutzentaler, Pfaffelhuber, Printz (2018) Stochastic averaging for multiscale Markov processes with an application to a Wright-Fisher model with fluctuating selection. *Submitted*.

- The generator is

$$G = G_0 + nG_1 + n^2 G_2$$

with $G_2 f = 0$ if f only depends on x .

- For $n \rightarrow \infty$,

$$\begin{aligned} f(X_t) - \int_0^t G f(X_s, Z_s) ds \\ \approx f(X_t) - \int_0^t G_0 f(X_s) + nG_1 f(X_s, Z_s) ds \end{aligned}$$

is a martingale.

- For $n \rightarrow \infty$, Z is a fast process with equilibrium π_x , so

$$\begin{aligned} & (f + \tfrac{1}{n}h)(X_t, Z_t) - \int_0^t G(f + \tfrac{1}{n}h)(X_s, Z_s) ds \\ & \approx f(X_t) - \int_0^t G_0 f(X_s) + nG_1 f(X_s, Z_s) \\ & \quad + G_1 h(X_s, Z_s) + nG_2 h(X_s, Z_s) ds \\ & \approx f(X_t) - \int_0^t \underbrace{G_0 f(X_s) + \overbrace{\mathbb{E}_{\pi_{X_s}}[G_1 h(X_s, Z)]}_{=: \tilde{G}_1 h(X_s)}}_{\text{potential generator of limit process}} ds \end{aligned}$$

is a martingale.

- The generator is

$$Gf(x, z) = \underbrace{G_0 f(x, z)}_{\text{mutation, resampling}} + \underbrace{n G_1 f(x, z)}_{\text{selection}} + \underbrace{n^2 G_2 f(x, z)}_{\text{environment}}$$

- Let f only depend on x and find h with $G_2 h = -G_1 f$
- For the limit process X ,

$$f(X_t) - \int_0^t G_0 f(X_s) + \bar{G}_1 h(X_s) ds$$

is a martingale.

- The generator is

$$Gf(x, z) = \underbrace{G_0 f(x, z)}_{\text{mutation, resampling}} + \underbrace{n G_1 f(x, z)}_{\text{selection}} + \underbrace{n^2 G_2 f(x, z)}_{\text{environment}}$$

$$G_2 f(x, z) = \mathbb{E}_\pi[f(x, Z)] - f(x, z)$$

- Let f only depend on x and find h with $G_2 h = -G_1 f$
- For the limit process X ,

$$f(X_t) - \int_0^t G_0 f(X_s) + \bar{G}_1 h(X_s) ds$$

is a martingale.

- For $h = G_1 f$,

$$G_2 h(x, z) = \mathbb{E}_\pi[Z]g(x) - G_1 f(x, z) = -G_1 f(x, z)$$

- Limit has generator, for smooth f ,

$$Gf = G_0 f + \mathbb{E}_\pi[G_1 G_1 f(x, Z)]$$

Theorem (Baumdicker, Huss, P, 2018)

As $n \rightarrow \infty$, $X = X^n$ converges weakly to the unique solution of

$$\begin{aligned}
 dX_{\ell 0} &= \frac{\sigma^2}{\gamma} X_{\ell 0} X_1 (X_1 - X_0) dt + \theta_\ell (X_{\ell 1} - X_{\ell 0}) dt \\
 &\quad + \sqrt{X_{\ell 0} (1 - X_{\ell 0})} dW_{\ell 0} + \sqrt{\frac{2\sigma^2}{\gamma}} X_{\ell 0} X_1 dW, \\
 dX_{\ell 1} &= \frac{\sigma^2}{\gamma} X_{\ell 1} X_0 (X_0 - X_1) dt + \theta_\ell (X_{\ell 0} - X_{\ell 1}) dt \\
 &\quad + \sqrt{X_{\ell 1} (1 - X_{\ell 1})} dW_{\ell 1} - \sqrt{\frac{2\sigma^2}{\gamma}} X_{\ell 1} X_0 dW, \\
 dX_{h0} &= \frac{\sigma^2}{\gamma} X_{h0} X_1 (X_1 - X_0) dt + \theta_h (X_{h1} - X_{h0}) dt \\
 &\quad + \sqrt{X_{h0} (1 - X_{h0})} dW_{h0} + \sqrt{\frac{2\sigma^2}{\gamma}} X_{h0} X_1 dW, \\
 dX_{h1} &= \frac{\sigma^2}{\gamma} X_{h1} X_0 (X_0 - X_1) dt + \theta_h (X_{h0} - X_{h1}) dt \\
 &\quad + \sqrt{X_{h1} (1 - X_{h1})} dW_{h1} - \sqrt{\frac{2\sigma^2}{\gamma}} X_{h1} X_0 dW,
 \end{aligned}$$

where W is an independent Brownian motion.

Corollary 1 (Karlin-Levikson model)

If $\theta = \theta_h = \theta_\ell$, $X_0 = X$ satisfies

$$dX = \frac{2\sigma^2}{\gamma} X(1-X)\left(\frac{1}{2} - X\right)dt + \theta(1-2X)dt \\ + \sqrt{X(1-X)}dW + \sqrt{\frac{2\sigma^2}{\gamma}} X(1-X)dW'$$

with independent Brownian motions W, W' .

Corollary 2 (High versus low mutators)

$$dX_h = \frac{\sigma^2}{\gamma} (X_{h0}X_{\ell1} - X_{h1}X_{\ell0})(X_1 - X_0)dt \\ + \sqrt{X_h X_\ell}dW + \sqrt{\frac{2\sigma^2}{\gamma}} (X_{h0}X_{\ell1} - X_{h1}X_{\ell0})dW'$$

with independent Brownian motions W, W' .

When is a high mutation rate beneficial?

Theorem (Baumdicker, Huss, P; 2018)

If $X_h(0) = x$, $X_{h0}(0) = px$, $X_{\ell 0} = q(1 - x)$,

$$\mathbf{P}_x(X_h(\infty) = 1) = x + \frac{\sigma^2}{8\gamma} x(1 - x)f + o(\sigma^2/\gamma)$$

with

$$\begin{aligned} f = & (1 - 2x) \left(\frac{1}{3} - \frac{4}{(3 + \theta_h + \theta_\ell)} (1 - 2p)(1 - 2q) \right) \\ & + \frac{3}{(3 + 2\theta_\ell)} \left((1 - x)(1 - 2q)^2 - \frac{1}{1 + \theta_\ell} \right) \\ & - \frac{3}{(3 + 2\theta_h)} \left(x(1 - 2p)^2 - \frac{1}{1 + \theta_h} \right) \\ & + 3 \frac{\theta_h - \theta_\ell}{(1 + \theta_\ell)(1 + \theta_h)}. \end{aligned}$$

When is a high mutation rate beneficial?

- Assume $p = q = \frac{1}{2}$, i.e.

$$\mathbf{P}_x(X_h(\infty) = 1) \approx x + \frac{\sigma^2}{8\gamma} x(1-x).$$

$$\left(\frac{1}{3}(1-2x) - \frac{3}{(3+2\theta_\ell)} \frac{1}{1+\theta_\ell} + \frac{3}{(3+2\theta_h)} \frac{1}{1+\theta_h} + 3 \frac{\theta_h - \theta_\ell}{(1+\theta_\ell)(1+\theta_h)} \right)$$

- For a maximum, solve

$$-4(\theta_\ell - 1)\theta_h^2 + 8(2 - \theta_\ell)\theta_h + 2(7 - 2\theta_\ell) = 0,$$

$$\theta_h = \frac{4 - 2\theta_\ell + \sqrt{2(\theta_\ell + 1)}}{2(\theta_\ell - 1)}.$$

\Rightarrow there is a fixed point at $\theta \approx 1.78$.

- For small σ^2/γ ,

$$\begin{aligned}\mathbf{P}_x^{\sigma^2/\gamma}[X_h(\infty) = 1] &= \mathbf{E}_x^{\sigma^2/\gamma}[X_h(\infty)] = x + \int_0^\infty \mathbf{E}_x^{\sigma^2/\gamma}[\bar{G}X_h(s)]ds \\ &= x + \frac{\sigma^2}{\gamma} \int_0^\infty \mathbf{E}_x^{\sigma^2/\gamma}[(X_{h0}(t)X_{\ell 1}(t) - X_{h1}(t)X_{\ell 0}(t))(X_1(t) - X_0(t))]dt\end{aligned}$$

When is a high mutation rate beneficial?

- For small σ^2/γ ,

$$\begin{aligned}
 \mathbf{P}_x^{\sigma^2/\gamma}[X_h(\infty) = 1] &= \mathbf{E}_x^{\sigma^2/\gamma}[X_h(\infty)] = x + \int_0^\infty \mathbf{E}_x^{\sigma^2/\gamma}[\bar{G}X_h(s)]ds \\
 &= x + \frac{\sigma^2}{\gamma} \int_0^\infty \mathbf{E}_x^{\sigma^2/\gamma}[(X_{h0}(t)X_{\ell1}(t) - X_{h1}(t)X_{\ell0}(t))(X_1(t) - X_0(t))]dt \\
 &= x + \frac{\sigma^2}{\gamma} \int_0^\infty \underbrace{\mathbf{E}_x^0[(X_{h0}(t)X_{\ell1}(t) - X_{h1}(t)X_{\ell0}(t))(X_1(t) - X_0(t))]}_{\text{can be computed using Kingman's coalescent}} dt \\
 &\quad + o(\sigma^2/\gamma)
 \end{aligned}$$

- Second-order selection favors types which have fit offspring (rather than being fit themselves).
- Fluctuating selection can be treated with stochastic averaging.
- For small selection strength or fast fluctuating environments, we computed the optimal mutation rate.