Optimal designs for inhibition models (with Holger Dette, Katrin Kettelhake, Tilman Möller)

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The Michaelis-Menten model is given by

$$Y_i = \eta(S_i, \tilde{\theta}) + \varepsilon_i = rac{VS_i}{S_i + K_m} + \varepsilon_i, \quad i = 1, \dots, n$$





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- the dose levels $S_i \in \mathcal{S} = [S_{\min}, S_{\max}]$, $0 \leq S_{\min}$,
- $\tilde{\theta} = (V, K_m)^T \in \mathbb{R}^2$ is the unknown parameter, • $\varepsilon_i \stackrel{i.i.d}{\sim} \mathcal{N}(0, \sigma^2), \ \sigma^2 > 0.$



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BUT:

- Many diseases require co-administration of several drugs.
- New drugs are often also screened for their inhibitory potential.
- Adequate modeling has to reflect this fact.
 - We extend the model by including the effect of inhibitor concentration.



Instead of the Michaelis-Menten model

$$Y_i = \eta(S_i, \tilde{\theta}) + \varepsilon_i = \frac{VS_i}{S_i + K_m} + \varepsilon_i, \quad i = 1, \dots, n$$

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The non-competitive inhibition model

We consider the non-competitive inhibition model

$$\begin{aligned} Y_i &= \eta(S_i, I_i, \theta) + \varepsilon_i \\ &= \frac{V \cdot S_i}{(K_m + S_i)(1 + \frac{I_i}{K_{ic}})} + \varepsilon_i , \ i = 1, \dots, n, \end{aligned}$$

where

- $(S_i, I_i), \in S = [S_{\min}, S_{\max}] \times [I_{\min}, I_{\max}],$ $0 \le S_{\min} < S_{\max}$ and $0 \le I_{\min} < I_{\max},$
- $heta = (V, \mathcal{K}_m, \mathcal{K}_{ic})^{\mathcal{T}} \in \mathbb{R}^3$ is the unknown parameter,

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The structure of the information matrix

Let ξ be an approximate design with finite support in S, that is,

$$\xi = \begin{pmatrix} (S_1, I_1) & \cdots & (S_k, I_k) \\ \xi_1 & \cdots & \xi_k \end{pmatrix}.$$



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$$\xi = \left(\begin{array}{ccc} (S_1, I_1) & \cdots & (S_k, I_k) \\ \xi_1 & \cdots & \xi_k \end{array}\right).$$

The information matrix of the design ξ is then given by

$$M(\xi,\theta) = \int_{\mathcal{S}} \frac{\partial \eta(\mathcal{S}, I, \theta)}{\partial \theta} \Big(\frac{\partial \eta(\mathcal{S}, I, \theta)}{\partial \theta} \Big)^{\mathsf{T}} d\xi(\mathcal{S}, I),$$

where

$$\frac{\partial \eta(S,I,\theta)}{\partial \theta} = \frac{S}{(K_m+S)} \frac{1}{(1+I/K_{ic})} \left(1, -\frac{V}{K_m+S}, \frac{V \cdot I/K_{ic}^2}{1+I/K_{ic}}\right)^T.$$

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Considered optimality criteria

We are now interested in determining

• D-optimal designs, that is designs such that

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• D-optimal designs, that is designs such that

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• e_j -optimal designs, j = 1, 2, 3, that is designs that that

$$\Phi_{e_j}\{M(\xi,\theta)\} = (e_j M^-(\xi,\theta)e_j)^{-1}, \ j = 1,2,3.$$

is maximised with respect to ξ under the condition that $e_j \in \text{Range}(M(\xi, \theta))$.

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The non-competitive inhibition model and optimal design

Main problem:

• Model is highly non-linear and therefore the criteria are difficult to analyse mathematically.

That is the reason why:

- Not much literature on optimal designs for this type of models. (see Youdim et al. (2010); Bogacka et al. (2011); Atkinson and Bogacka (2013); Chen et al. (2017))
- Most results exist for D-optimal and D_s -optimal designs.
- In most cases optimal designs have to be found numerically.

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- In most cases optimal designs have to be found numerically.

Our idea:

- Use a non-linear transformation of the variables (*S*, *I*) to achieve multivariate polynomial regression model.
 - Then the analysis of the model becomes much easier.

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We define a one-to-one transformation of the variable (S, I) by

$$\begin{pmatrix} x \\ y \end{pmatrix} = \psi(S, I) = \begin{pmatrix} \frac{S}{K_{m}+S} \\ \frac{1}{1+I/K_{ic}} \end{pmatrix}.$$

where $(x, y) \in \mathcal{X} = [x_{\min}, x_{\max}] \times [y_{\min}, y_{\max}].$

The boundary points of the two intervals are defined by

$$x_{\min} = \frac{S_{\min}}{K_m + S_{\min}}; x_{\max} = \frac{S_{\max}}{K_m + S_{\max}}; y_{\min} = \frac{1}{1 + l_{\max}/K_{ic}}; y_{\max} = \frac{1}{1 + l_{\min}/K_{ic}}.$$

Note that $x_{\min}, x_{\max}, y_{\min}, y_{\max} \in [0, 1]$.

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Transformation of the gradient

Using the transformation the gradient

$$\frac{\partial \eta(S,I,\theta)}{\partial \theta} = \frac{S}{(K_m+S)} \frac{1}{(1+I/K_{ic})} \left(1, -\frac{V}{K_m+S}, \frac{V \cdot I/K_{ic}^2}{1+I/K_{ic}}\right)^T$$

can be represented by

$$\frac{\partial \eta(S,I,\theta)}{\partial \theta} = A(\theta)f(x,y),$$

where the non-singular matrix $A(\theta)$ and the vector f(x, y) are given by

$$A(\theta) = \begin{pmatrix} 1 & 0 & 0 \\ -\frac{V}{K_m} & \frac{V}{K_m} & 0 \\ \frac{V}{K_{ic}} & 0 & -\frac{V}{K_{ic}} \end{pmatrix}, \quad f(x,y) = xy \begin{pmatrix} 1 \\ x \\ y \end{pmatrix}.$$

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The vector f(x, y) corresponds to the regression function of a multivariate polynomial regression model.

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Transformation of the design and the information matrix

We transform the design ξ , that is,

$$\xi$$
 design on $\mathcal{S} \underset{\psi^{-1}(x,y)}{\overset{\psi(S,I)}{\rightleftharpoons}} \tilde{\xi}$ induced design on \mathcal{X} .



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The information matrix can then be represented by

$$M(\xi,\theta) = A(\theta)\tilde{M}(\tilde{\xi})A^{\mathsf{T}}(\theta),$$

where the matrix $\tilde{M}(\tilde{\xi})$ is defined by

$$\tilde{M}(\tilde{\xi}) = \int_{\mathcal{X}} f(x, y) f^{\mathsf{T}}(x, y) d\tilde{\xi}(x, y).$$



maximising $\Phi(M(\xi,\theta)) \Leftrightarrow$ maximising $\Phi(A(\theta)\tilde{M}(\tilde{\xi})A^{T}(\theta))$.



maximising $\Phi(M(\xi, \theta)) \Leftrightarrow$ maximising $\Phi(A(\theta)\tilde{M}(\tilde{\xi})A^{T}(\theta))$. In the case of *D*-optimality, we get:

 $\Phi_D(M(\xi,\theta)) = \Phi_D(A(\theta)\tilde{M}(\tilde{\xi})A^T(\theta)) = (\det A(\theta))^2 (\det \tilde{M}(\tilde{\xi})).$



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In the case of e_j -optimality, we get for j = 1, 2, 3:

$$\begin{split} \Phi_{e_j}(M(\xi,\theta)) &= (e_j^T(A(\theta)\tilde{M}(\tilde{\xi})A^T(\theta))^- e_j)^{-1} \\ &= ((A^{-1}(\theta)e_j)^T\tilde{M}^-(\tilde{\xi})(A^{-1}(\theta))^{-1}e_j))^{-1} \\ &= (\tilde{e}_j^T\tilde{M}^-(\tilde{\xi})\tilde{e}_j)^{-1}. \end{split}$$

UNIVERSITÄT BOCHUM maximising $\Phi(M(\xi, \theta)) \Leftrightarrow$ maximising $\Phi(A(\theta)\tilde{M}(\tilde{\xi})A^{T}(\theta))$. In the case of *D*-optimality, we get:

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e_j-optimal designs



Structure of D-optimal designs I

Theorem

Let $\mathcal{X} = [x_{\min}, x_{\max}] \times [y_{\min}, y_{\max}]$ and $x_{\min}^0 = \frac{x_{\min}}{x_{\max}}$, $y_{\min}^0 = \frac{y_{\min}}{y_{\max}}$. The D-optimal design is supported at three points if and only if (x_{\min}^0, y_{\min}^0) is within the set $\mathcal{D} \subset \mathbb{R}^2$.





Structure of D-optimal designs II

Theorem

Let
$$\mathcal{X} = [x_{\min}, x_{\max}] \times [y_{\min}, y_{\max}]$$
 and $x_{\min}^0 = \frac{x_{\min}}{x_{\max}}$, $y_{\min}^0 = \frac{y_{\min}}{y_{\max}}$.
If $(x_{\min}^0, y_{\min}^0) \in \mathcal{D} \subset \mathbb{R}^2$, the D-optimal design is given by

$$\tilde{\xi}^* = \begin{pmatrix} \left(\max\{x_{\min}, \frac{x_{\max}}{2}\}, y_{\max}\right) & \left(x_{\max}, \max\{\frac{y_{\max}}{2}, y_{\min}\}\right) & \left(x_{\max}, y_{\max}\right) \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \end{pmatrix}.$$





Remarks to D-optimal designs

• Bogacka et al. (2011) stated that the *D*-optimal design is always supported by three points.



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Remarks to D-optimal designs

- Bogacka et al. (2011) stated that the *D*-optimal design is always supported by three points.
- Chen et al. (2017) found a parameter combination where the corresponding design with three support points fails to be *D*-optimal.
- Because of the transformation we were able to derive explicit conditions under which the saturated design is *D*-optimal.





Example

Let $\theta = (V, K_m, K_{ic})^T = (1, 4, 2)^T$ and $S = [7, 30] \times [30, 60]$. Then the regression function is given by $\eta(S, I, \theta) = \frac{S}{(4+S)(1+\frac{I}{2})}$.





Let $\theta = (V, K_m, K_{ic})^T = (1, 4, 2)^T$ and $S = [7, 30] \times [30, 60]$.

The transformed design space is then given by: $\mathcal{X} = [\frac{7}{11}, \frac{15}{17}] \times [\frac{1}{31}, \frac{1}{16}].$



Let $\theta = (V, K_m, K_{ic})^T = (1, 4, 2)^T$ and $S = [7, 30] \times [30, 60]$. The transformed design space is then given by: $\mathcal{X} = [\frac{7}{11}, \frac{15}{17}] \times [\frac{1}{31}, \frac{1}{16}]$. Therefore: $x_{\min}^0 = \frac{119}{165}$ and $y_{\min}^0 = \frac{16}{31}$ and $(x_{\min}^0, y_{\min}^0) \in \mathcal{D}$.



Let $\theta = (V, K_m, K_{ic})^T = (1, 4, 2)^T$ and $S = [7, 30] \times [30, 60]$. The transformed design space is then given by: $\mathcal{X} = [\frac{7}{11}, \frac{15}{17}] \times [\frac{1}{31}, \frac{1}{16}]$. Therefore: $x_{\min}^0 = \frac{119}{165}$ and $y_{\min}^0 = \frac{16}{31}$ and $(x_{\min}^0, y_{\min}^0) \in \mathcal{D}$.

The *D*-optimal design is (using the transformation to the original space):

$$\xi^* = \begin{pmatrix} (7,30) & (30,30) & (30,60) \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \end{pmatrix}.$$



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Let $\theta = (V, K_m, K_{ic})^T = (1, 4, 2)^T$ and $S = [15, 30] \times [30, 60]$. The transformed design space is then given by: $\mathcal{X} = [\frac{15}{19}, \frac{15}{17}] \times [\frac{1}{31}, \frac{1}{16}]$.

Therefore:
$$x_{\min}^0 = \frac{17}{19}$$
 and $y_{\min}^0 = \frac{16}{31}$ and $(x_{\min}^0, y_{\min}^0) \notin \mathcal{D}$.



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In the case of e_j -optimality, we get for j = 1, 2, 3:

$$\Phi_{e_j}(M(\xi,\theta)) = ((A^{-1}(\theta)e_j)^T \tilde{M}^-(\tilde{\xi})(A^{-1}(\theta))^{-1}e_j))^{-1} = (\tilde{e}_j^T \tilde{M}^-(\tilde{\xi})\tilde{e}_j)^{-1}.$$

What does \tilde{e}_j look like for the different cases?



e_i -optimality in the transformed model

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What does \tilde{e}_j look like for the different cases?

$$\tilde{e}_{j} = \begin{cases} (1, 1, 1)^{T} , & j = 1 \\ \frac{K_{m}}{V} e_{2} , & j = 2 . \\ \frac{K_{ic}}{V} e_{3} , & j = 3 \end{cases}$$



e_i -optimality in the transformed model

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Theorem

The optimal design maximising $(\tilde{e}_2^T \tilde{M}^- (\tilde{\xi}) \tilde{e}_2)^{-1}$ is of the form

$$\tilde{\xi} = \begin{pmatrix} (x_{\max}, y_{\max}) & (\overline{x}, y_{\max}) \\ \frac{\overline{x}}{1 + \overline{x}} & \frac{1}{1 + \overline{x}} \end{pmatrix},$$

where $\overline{x} = \max \{x_{\min}, (\sqrt{2}-1)x_{\max}\}.$



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where $\overline{x} = \max \{x_{\min}, (\sqrt{2}-1)x_{\max}\}.$

We have to transform the design $\tilde{\xi}$ to the original space S to get the optimal design for estimating the Michaelis-Menten constant K_m .

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Optimal design for estimating the parameter K_m

Corollary

The optimal design for estimating the Michaelis-Menten constant K_m is given by

$$\xi = \begin{pmatrix} \left(\textbf{S}_{\mathsf{max}}, \textbf{I}_{\mathsf{min}} \right) & \left(\overline{\textbf{S}}, \textbf{I}_{\mathsf{min}} \right) \\ 1 - \omega & \omega \end{pmatrix},$$

where

$$\overline{S} = \max \left\{ S_{\min} , \frac{K_m S_{\max}(\sqrt{2} - 1)}{K_m + (2 - \sqrt{2})S_{\max}} \right\},$$

$$\omega = \left(1 + \max \left\{ \frac{S_{\min}}{K_m + S_{\min}} , \frac{(\sqrt{2} - 1)S_{\max}}{K_m + S_{\max}} \right\} \right)^{-1}.$$



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Let $\theta = (V, K_m, K_{ic})^T = (1, 4, 2)^T$ and $S = [7, 30] \times [30, 60]$.

Then the design for estimating the constant K_m is given by

$$\xi = \begin{pmatrix} (7,30) & (30,30) \\ 0.6\overline{1} & 1 - 0.6\overline{1} \end{pmatrix}$$





Conclusion and further comments

- The non-linear transformation is useful to make the optimisation problem more tractable.
- Similar transformations are possible for other inhibition models and can simplify the (numerical) optimisation problem.
- For instance, we can use the algorithm presented by Fabrice Gamboa today.



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- The non-linear transformation is useful to make the optimisation problem more tractable.
- Similar transformations are possible for other inhibition models and can simplify the (numerical) optimisation problem.
- For instance, we can use the algorithm presented by Fabrice Gamboa today.

Thank you very much for your attention!



- Atkinson, A. C. and Bogacka, B. (2013). Robust experimental design for choosing between models of enzyme inhibition. In D. Ucinski, A. C. Atkinson, M. P., editor, mODa 10 - Advances in Model-Oriented Design and Analysis. Contributions to Statistics. Springer.
- Bogacka, B., Patan, M., Johnson, P. J., Youdim, K., and Atkinson, A. C. (2011). Optimum design of experiments for enzyme inhibition kinetic models. *Journal of Biopharmaceutical Statistics*, 21(3):555–572.
- Chen, P.-Y., Chen, R.-B., Tung, H.-C., and Wong, W. K. (2017). Standardized maximin *D* -optimal designs for enzyme kinetic inhibition models. *Chemometrics and Intelligent Laboratory Systems*, 169:79 – 86.
- Youdim, K. A., Atkinson, A. C., Patan, M., Bogacka, B., and Johnson, P. J. (2010). Potential application of *D*-optimal designs in the efficient investigation of cytochrome "P450" inhibition kinetic models. *Drug Metabolism and Disposition*, 38(7):1019–1023.

