Functional data reduction issue for robust inversion

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### The framework : automotive pollution control



where  $D \subset \mathbb{R}^{p}$  is a compact and  $\mathcal{F}$  a functional space.

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## The framework : automotive pollution control



where  $D \subset \mathbb{R}^p$  is a compact and  $\mathcal{F}$  a functional space.

Probabilistic description of uncertainty : V is a random variable valued in  $\mathcal{F}$ .

For a fixed  $s \in \mathbb{R}$ , define

$$\begin{array}{lll} \Gamma^*_{a.s.} &:= & \{\mathbf{x} \in D \ s.t. \ g(\mathbf{x}, \mathbf{V}) \leq s & almost \ surely\}, \\ \Gamma^*_{\alpha} &:= & \{\mathbf{x} \in D \ s.t. \ \mathbb{P}(g(\mathbf{x}, \mathbf{V}) \leq s) \geq 1 - \alpha\}, \\ \Gamma^* &:= & \{\mathbf{x} \in D \ s.t. \ f(x) = \mathbb{E}[g(\mathbf{x}, \mathbf{V})] \leq s\} &:= f^{-1}(T) \ , \text{where} \ T = (-\infty, s]. \end{array}$$

Aim : estimate  $\Gamma^*$ , the excursion set of f below t.

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#### The framework

Model :  $g : D \times \mathcal{F} \to \mathbb{R}$  with  $D \subset \mathbb{R}^p$  and  $\mathcal{F}$  a functional space.

Objective : estimate  $\Gamma^* \subset \mathbb{R}^p$  from *n* evaluations of function  $f(\cdot) = \mathbb{E}[g(\cdot, \mathbf{V})], \mathbf{f}_n = (f(\mathbf{x}^1), f(\mathbf{x}^2), ..., f(\mathbf{x}^n)).$ 

#### Issues :

- each evaluation  $f(\mathbf{x})$  requires the estimation of  $\mathbb{E}[g(\mathbf{x}, \mathbf{V})]$ ,
- V is known through  $\kappa$  realizations  $\mathbf{v}^1, \ldots \mathbf{v}^{\kappa}$ ,
- each evaluation  $g(\mathbf{x}, \mathbf{v})$  is costful.

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## An analytical example (1/5)

Let  $g: D imes \mathcal{F} 
ightarrow \mathbb{R}$  defined as

$$g(\mathbf{x}, \mathbf{V}) = |0.1 \cos(x_1 \max_t V_t) \sin(x_2) \cdot (x_1 + x_2 \min_t V_t)^2| \cdot \int (30 + V_t)^{\frac{x_1 \cdot x_2}{20}} dt \cdot \max_t V_t$$

with  $\mathbf{x} = (x_1, x_2) \in D \subset \mathbb{R}^2$  and  $\mathbf{V} = (V_t, t \ge 0)$  a BM with constant drift equal to 2.



$$\begin{split} \Gamma^* &= \{ \mathbf{x} \in [1.5,5] \times [3.5,5] \ : \ f(\mathbf{x}) \le 1.2 \} \\ &= \{ \mathbf{x} \in [1.5,5] \times [3.5,5] \ : \ \mathbb{E} \big[ g(\mathbf{x},\mathbf{V}) \big] \le 1.2 \} \end{split}$$



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# Bayesian Approach / Kriging<sup>4</sup>

*f* is seen as a realization of a Gaussian Process  $(Z_x)_{x \in D}$  with prior mean *m* and covariance kernel *k*.

$$\Gamma^*$$
 is a realization of  $\Gamma := \{\mathbf{x} \in D : Z(\mathbf{x}) \leq s\}.$ 

Let  $\mathcal{X}_n = (\mathbf{x}^1, \dots, \mathbf{x}^n)$ . Given the evaluations  $\mathbf{f}_n = f(\mathcal{X}_n)$ , the posterior field  $Z | Z(\mathcal{X}_n) = \mathbf{f}_n$  follows a Gaussian distribution with mean and covariance kernel :

$$m_n(\mathbf{x}) = m(\mathbf{x}) + k(\mathbf{x}, \mathcal{X}_n)k(\mathcal{X}_n, \mathcal{X}_n)^{-1}(\mathbf{f}_n - m(\mathcal{X}_n))$$
  

$$k_n(\mathbf{x}, \mathbf{y}) = k(\mathbf{x}, \mathbf{y}) - k(\mathbf{x}, \mathcal{X}_n)k(\mathcal{X}_n, \mathcal{X}_n)^{-1}k(\mathcal{X}_n, \mathbf{y})$$

What about the probability distribution of  $\Gamma | Z(X_n) = \mathbf{f}_n$ ?

4. Clément CHEVALIER (2013). « Fast uncertainty reduction strategies relying on Gaussian process models ». Thèse de doct. Citeseer  $\Box \rightarrow \langle \Box \rangle = \langle \Box \rangle$ 

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## An analytical example (2/5)

What about the probability distribution of  $\Gamma | Z(X_n) = \mathbf{f}_n$ ?

Let us consider :

- priors on the Gaussian process : constant mean function, 5/2 Matern covariance kernel;
- initial DoE  $\mathcal{X}_n = (\mathbf{x}^1, \dots, \mathbf{x}^n)$ , n = 9 (black triangles).

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How to summarize the probability distribution of  $\Gamma | Z(X_n) = f_n$ ?

What are, for random sets, the notions of

- expectation,
- deviation.

**Remark :** let  $\mu$  be a Borel measure on D. Then,

$$\mathbb{E}\left(\mu(\Gamma)|Z(\mathcal{X}_n)=\mathbf{f}_n\right)=\int\mathbb{P}_n(\mathbf{x}\in\Gamma)\mu(d\mathbf{x})$$

with

$$\mathbb{P}_n(\mathbf{x} \in \Gamma) = \mathbb{P}(Z(\mathbf{x}) \leq \mathbf{s} | Z(\mathcal{X}_n) = \mathbf{f}_n) \cdot$$

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Theory of random sets, Vorob'ev quantiles <sup>5 6</sup>

The conditional coverage function of  $\pmb{\Gamma}$  is defined as

 $p_n: \mathbf{x} \in D \longrightarrow \mathbb{P}_n(\mathbf{x} \in \Gamma) = \mathbb{P}(Z(\mathbf{x}) \leq s \mid Z(\mathcal{X}_n) = \mathbf{f}_n) \in [0, 1].$ 

Under our Gaussian framework :

 $p_n(\mathbf{x}) = \phi\Big(\frac{m_n(\mathbf{x}) - t}{\sqrt{k_n(\mathbf{x}, \mathbf{x})}}\Big),$ 

with  $\phi$  the Gaussian cumulative distribution function.

For any  $\alpha \in [0,1]$ , let us define

 $Q_{\alpha} = \{\mathbf{x} \in D : p_n(\mathbf{x}) \geq \alpha\}.$ 



5. Ilya MOLCHANOV (2006). Theory of random sets.

6. O Yu Vorob'ev (1984). Srednemernoje modelirovanie (mean-measure modelling). < 🖻 🛌 🔗

Stepwise Uncertainty Reduction Modeling Functional Uncertainties Theory of Random Sets

#### Theory of random sets, Vorob'ev Expectation

The Vorob'ev Expectation of  $\Gamma$  is the set  $Q_{\alpha_n^*}$  such that

 $\mu(Q_{\alpha_n^*}) = \mathbb{E}_n(\mu(\Gamma)) = \mathbb{E}(\mu(\Gamma)|Z(\mathcal{X}_n) = \mathbf{f}_n).$ 



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## Vorob'ev deviation

$$\mathbb{E}_{n}[\mu(Q_{\alpha_{n}^{*}}\Delta\Gamma)] = \mathbb{E}_{n}\Big[\mu(\Gamma \cap Q_{\alpha_{n}^{*}}^{c}) + \mu(Q_{\alpha_{n}^{*}}\cap\Gamma^{c})\Big]$$
$$= \int_{Q_{\alpha_{n}^{*}}^{c}} p_{n}(\mathbf{x})\mu(\mathrm{d}\mathbf{x}) + \int_{Q_{\alpha_{n}^{*}}} (1 - p_{n}(\mathbf{x}))\mu(\mathrm{d}\mathbf{x})$$

with  $\mathbb{E}_n = \mathbb{E}[. \mid Z(\mathcal{X}_n) = \mathbf{f}_n].$ 

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#### Vorob'ev deviation

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with  $\mathbb{E}_n = \mathbb{E}[. \mid Z(\mathcal{X}_n) = \mathbf{f}_n].$ 

The Vorob'ev expectation  $Q_{\alpha_n^*}$  is the minimizer, among all closed subsets Q satisfying  $\mu(Q) = \mathbb{E}_n[\mu(\Gamma)]$ , of  $\mathbb{E}_n[\mu(Q\Delta\Gamma)]$ .

Infill strategy : sampling strategy to choose new simulation points of f in order to reduce the Vorob'ev deviation.

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Introduction Theory of Random Sets **Stepwise Uncertainty Reduction** Application

Stepwise Uncertainty Reduction (SUR strategy)

## Uncertainty function : Vorob'ev deviation

$$\begin{aligned} \mathcal{H}_{n}^{\text{uncert}} &= \mathbb{E}_{n}[\mu(Q_{\alpha_{n}^{*}}\Delta\Gamma)] \\ &= \int_{D} \left( p_{n}(\mathbf{x})\mathbb{1}_{\{p_{n}(\mathbf{x}) < \alpha_{n}^{*}\}} + (1 - p_{n}(\mathbf{x}))\mathbb{1}_{\{p_{n}(\mathbf{x}) \geq \alpha_{n}^{*}\}} \right) \mu(\mathrm{d}\mathbf{x}) \end{aligned}$$

# Stepwise Uncertainty Reduction<sup>7</sup>

find the best next evaluation point  $\mathbf{x}^{n+1}$  that optimally reduces the expected uncertainty  $\mathcal{H}_{n+1}$  on the future estimate, i.e.,

$$\mathbf{x}^{n+1} = \operatorname{argmin}_{\mathbf{x} \in D} \mathbb{E}_{n,\mathbf{x}}[\mathcal{H}_{n+1}^{\mathsf{uncert}}(\mathbf{x})].$$

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# An analytical example (3/5)

(left) Vorob'ev deviation

(right) boundary for the true  $\Gamma^*$  (red line) , estimation of  $\Gamma^*$  (green set)



On that application, the evaluations of f are computed analytically.

Our second objective is to include in the study the estimation of  $f(\mathbf{x})$  from a low number of evaluations  $g(\mathbf{x}, \mathbf{v}^{j}), j = 1, ..., N$ .

Crude scenario approach Modeling based on Functional Principal Components Analysis A sequential hybrid methodology Algorithm Application

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# An analytical example (4/5)

$$\Gamma^* = \{x \in D \;,\; f(x) \leq t\} \;,\; ext{where}\; f(x) = \mathbb{E}[g(x, \mathbf{V})]$$



FIGURE – A finite sample of  $\kappa$  realizations of V



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#### Crude scenario approach

The uncertainty of the functional variable V is represented by a subset of M curves  $\mathbf{v}^1, \ldots, \mathbf{v}^M$  (randomly chosen among the initial sample of curves). Then we draw randomly with replacement N curves  $\mathbf{w}^1, \ldots, \mathbf{w}^N$  in  $\{\mathbf{v}^1, \ldots, \mathbf{v}^M\}$ .

$$egin{aligned} &\Gamma^* = \{\mathbf{x} \in D \;,\; f(\mathbf{x}) = \mathbb{E}[g(\mathbf{x},\mathbf{V})] \leq s\} \ &\simeq \{\mathbf{x} \in D \;,\; f(\mathbf{x}) \simeq rac{1}{N} \sum_{i=1}^N g(\mathbf{x},\mathbf{w}^i) \leq s\}. \end{aligned}$$

# **Functional Principal Components**

Functional PCA is a statistical method for investigating the dominant modes of variation of functional data.

Let  $\mathcal{T}$  be a finite and closed interval of  $\mathbb{R}$ . Assume  $\mathcal{F} \subset \mathbb{L}^2(\mathcal{T})$ .

Then,  $\mathbf{v}_1, \ldots, \mathbf{v}_\kappa$  can be viewed as independent realizations of a stochastic process  $\mathbf{V}$  with unknown mean function  $\mathbb{E}[V(t)] = \mu_{\mathbf{V}}(t)$  and covariance function  $\operatorname{Cov}(V(s), V(t)) = G_{\mathbf{V}}(s, t)$ .

Then, if  $G_V$  is continuous on  $\mathcal{T} \times \mathcal{T}$ , it is well known that there exists an orthogonal expansion of  $G_V$  (in the  $\mathbb{L}^2$  sense) in terms of eigenfunctions  $\phi_k$  with associated eigenvalues  $\lambda_k$  (arranged in non increasing order), that is,

$$G_{\mathsf{V}}(s,t) = \sum_{k\geq 1} \lambda_k \phi_k(s) \phi_k(t).$$

The random function V(t),  $t \in T$  can be decomposed into an orthogonal expansion

$$V(t) = \mu_V(t) + \sum_{k\geq 1} \sqrt{\lambda_k} \eta_k \phi_k(t)$$

with  $\eta_k$  uncorrelated standardized random variables.

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A truncation argument 
$$V(t) = \mu_V(t) + \sum_{k=1}^m \sqrt{\lambda_k} \eta_k \phi_k(t).$$

**Probability density estimation for the vector**  $\eta = (\eta_1, \dots, \eta_m)^{\mathsf{T}}$  (non) parametric estimation.

Then, by sampling the vector  $\eta$ , we get :

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$$egin{aligned} & \mathbf{x} \in D \;,\; f(\mathbf{x}) = \mathbb{E}[g(\mathbf{x},\mathbf{V})] \leq s \} \ & \simeq \{\mathbf{x} \in D \;,\; f(\mathbf{x}) \simeq rac{1}{N} \sum_{i=1}^N g(\mathbf{x}, oldsymbol{\eta}^i) \leq s \}. \end{aligned}$$



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# A sequential hybrid methodology<sup>8</sup>

First Step : the adaptive selection of representative curves  $\mathbf{v}^1,\ldots,\mathbf{v}^M$  among the initial sample of curves

- Let  $\mathbf{v}^1, \ldots, \mathbf{v}^{\kappa}$  the initial sample of realizations of V.
- We project the initial sample on the first *m* components in the functional PCA :  $(\phi_k)_{k=1,...,m} \to \eta^1, \ldots, \eta^{\kappa} \kappa$  points in  $\mathbb{R}^m$ .



8. Mohamed Reda EL AMRI et al. (2018). ≪ Data-driven stochastic inversion under functional uncertainties ≫. URL: https://hal.inria.fr/hal-01704189



- We sequentially construct a space filling design in  $\mathbb{R}^m$  by picking points in  $\{\eta^1, \ldots, \eta^\kappa\}$ .
- At each selected point  $\eta$  (on the right), we select the corresponding curve (on the left).



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• The points  $\eta$  are chosen sequentially, with a min dist criterion in  $\mathbb{R}^m$ .



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We stop with N points η<sup>1</sup>,..., η<sup>N</sup> in ℝ<sup>m</sup> and the N corresponding curves v<sup>1</sup>,..., v<sup>N</sup> by considering a stopping criterion based on the the stability in the estimation of ℝ[g(x, V)].



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**Second Step :** the estimation of  $\mathbb{E}[g(\mathbf{x}, \mathbf{V})]$  with a weighted sum

• The computation of the weights is based on the Voronoi diagram.

$$\begin{aligned} & \mathsf{T}^* = \{ x \in D \ , \ f(x) = \mathbb{E}[g(\mathbf{x}, \mathbf{V})] \leq s \} \\ & \simeq \{ x \in D \ , \ f(\mathbf{x}) \simeq \frac{\sum_{i=1}^N w_i \ g(\mathbf{x}, \mathbf{v}^i)}{\sum_{i=1}^N w_i} \leq s \}. \end{aligned}$$



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### **Global algorithm**

#### Algorithm 1: Create an initial design of experiments (Doe) at n points in the control space X2: $l \leftarrow 2$ 3: while Stopping criterion not met (SUR) do $x_{n+1} \leftarrow \text{Using sampling criterion } \mathcal{J}_n$ 4: while Stopping criterion not met (Expectation Estimation) do 5: 6. Create design $D_l$ and weight calculus $w_l$ using (1) $test \leftarrow stopping criterion$ 7. $l \leftarrow l + 1$ 8. end while 9: Update Doe 10: 11: $n \leftarrow n + 1$ 12: end while 13: end

Stopping criterion are based on the stability of the estimation of  $\mathbb{E}[g(\mathbf{x}, \mathbf{V})]$  (hybrid weighted scenario approach) and on the stability of the Vorob'ev deviation (SUR strategy).

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# Analytical example (5/5)



FIGURE – For Fpca and W.scenario, the truncation argument m = 7 is chosen to explain 97% of the variance of **V**.

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Test Case IFPEN : Selective Catalytic Reduction



• Thresholds :

 $\begin{array}{l} \bullet \quad \mathrm{NO}_{\mathrm{x}}^{\mathrm{out}} & \leq & \mathsf{80} \ mg.km^{-1} \\ \bullet \quad \mathrm{NH}_{3}^{\mathrm{max}} & \leq & \mathsf{30} \ ppm \end{array}$ 

• Objects of interest :

- { $(x_1, x_2, ...)$  :  $\mathbb{E}_{(B,V)}[NO_x^{out}] \le 80 \text{ mg.km}^{-1}$ } ● { $(x_1, x_2, ...)$  :  $\mathbb{E}_{(B,V)}[NH_3^{max}] \le 30 \text{ ppm}$ }
- Sensitivity Analysis : the driving cycle V is the most influential variable

IFPEN test case : control strategy for an automotive  $NO_x$  depollution system

•  $\mathbf{v}^1, \ldots, \mathbf{v}^{100}$  random driving cycles.



•  $\text{NH}_3^{\text{max}}$  function of  $x_1$  and  $x_2$ ,  $D = [0, 0.6]^2$ 

$$\Gamma^* = \{\mathbf{x} = (\mathbf{x}_1, \mathbf{x}_2) \in D : \mathbb{E}[\mathrm{NH}_3^{\max}(\mathbf{x}, \mathbf{V})] \leq s\}.$$

The model is expensive to evaluate (chemical kinetic model).



IFPEN test case, iteration 1, initial DoE : n = 9 black triangles  $\langle \Box \rangle \land \langle \Box \rangle \land \langle \Xi \rangle$ 



Nbre of curves for expectation estimation

Stopping Criterion : L0 = 4, epsilon = 5e-3



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## Conclusion

- Robust Inversion Problem :
  - Robustness measure : expectation ;
  - Sequential strategy SUR.
- With functional Uncertainty :
  - Hybrid sequential strategy based on functional PCA, a space filling strategy and a weighted scenario approach.

## Perpectives

• Robustness measure : high probability, with some potentially small risk  $\alpha$ 

$$\Gamma_{\alpha} := \{ \mathbf{x} \in D : \mathbb{P} (g(\mathbf{x}, \mathbf{V}) \leq s) \geq 1 - \alpha \};$$

• SUR strategy to choose both the next **x** and the next **v**.

### Thanks for your attention !