

# Exogenous shock models in high dimensions

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## **Research questions**

(a) For which  $g_1, \ldots, g_d$  is the following a multivariate distribution function

$$(x_1,\ldots,x_d)\mapsto\prod_{k=1}^d g_k(x_{[k]}),$$

where  $x_{[1]} \leq x_{[2]} \leq \ldots \leq x_{[d]}$  is the ordered list of  $x_1, \ldots, x_d \in \mathbb{R}$ ?

- (b) What about stochastic representations?
- (c) Are there interesting examples / applications?





• First observation:

$$(X_1,\ldots,X_d)\sim\prod_{k=1}^d g_k(x_{[k]})$$
  $\Rightarrow$   $g_1(x)=\mathbb{P}(X_k\leq x), \quad \forall k=1,\ldots,d.$ 

- W.I.o.g. consider a copula framework, i.e.  $x_1, \ldots, x_d \in [0, 1]$  and  $g_1(x) = x$ .
- For which  $g_2, \ldots, g_d$  is the following a copula?

$$C(x_1,...,x_d) := x_{[1]} \cdot \prod_{k=2}^d g_k(x_{[k]})$$

• If C is a copula and  $g_1$  any univariate d.f. (resp.  $\bar{g}_1$  any univariate s.f.),

$$C(g_1(x_1), \dots, g_1(x_d)) := g_1(x_{[1]}) \cdot \prod_{k=2}^d g_k \circ g_1(x_{[k]}) \quad (\text{resp. } C(\bar{g}_1(x_1), \dots, \bar{g}_1(x_d)))$$

is a multivariate distribution function (resp. survival function).



## Some examples

• Independence:  $g_k \equiv id_{[0,1]}$  yields

$$C(x_1,\ldots,x_d)=\prod_{k=1}^d x_k.$$

• Exchangeable Marshall–Olkin copulas have  $g_k(x) := x^{a_{k-1}}$ , so

$$C(x_1, \dots, x_d) = \prod_{k=1}^d x_{[k]}^{a_{k-1}}$$
 for *d*-monotone  $a_0 = 1, a_1, \dots, a_{d-1} \ge 0$ .

• **[Durante et al. (2007)]** study copulas with  $g_k(x) := g(x), k = 2, ..., d$ , i.e.

$$C(x_1,\ldots,x_d) = x_{[1]} \prod_{k=2}^d g(x_{[k]}).$$



#### Set $\boldsymbol{\mathcal{D}}$ of univariate d.f.'s:

 $\mathcal{D} := \{F : [0,1] \to [0,1] : \text{ continuous, non-decreasing, strictly positive on } (0,1], F(1) = 1\}$  $= \{d.f.s \text{ of absolutely continuous r.v.s on } (0,1) \text{ with possibly an extra atom at } 0\}.$ 





#### **Bivariate case:**

• For  $g_1, g_2 : [0, 1] \to [0, 1], g_1(1) = g_2(1) = 1$ , consider

 $C(x_1, x_2) := g_1(x_{[1]}) g_2(x_{[2]}).$ 

• [Durante et al. (2008)]: C is a copula if and only if  $g_1 \equiv id_{[0,1]}$  and

(*i*)  $g_2(y) - g_2(x) \ge 0$ ,  $x, y \in [0, 1], x < y$ .

(*ii*)  $g_1(y)g_2(y) - 2g_1(x)g_2(y) + g_1(x)g_2(x) \ge 0$ ,  $x, y \in [0, 1], x < y$ .

- Analytical interpretation:
  - $\rightarrow$  (*i*)  $\Leftrightarrow$   $g_2$  is increasing.

 $\rightarrow$  (*ii*)  $\Leftrightarrow$   $g_2$  is continuous, strictly positive on (0, 1], and  $g_1/g_2$  is increasing.

• **Remark**:  $g_1/g_2 \in \mathcal{D}$ .

#### Necessary condition:

• Assume that  $(U_1, \ldots, U_d) \sim x_{[1]} \cdot \prod_{k=2}^d g_k(x_{[k]})$ . Then

$$g_j(x) = \frac{x \cdot \prod_{k=2}^j g_k(x)}{x \cdot \prod_{k=2}^{j-1} g_k(x)}$$
  
= 
$$\frac{\mathbb{P}(\max\{U_1, \dots, U_j\} \le x)}{\mathbb{P}(\max\{U_1, \dots, U_{j-1}\} \le x)} \implies x \cdot \prod_{i=2}^j g_i(x) = \mathbb{P}(\max\{U_1, \dots, U_j\} \le x).$$

- So it is necessary that  $x \cdot \prod_{i=2}^{j} g_i(x) \in \mathcal{D}$  for all j = 2, ..., d.
- But we need more.



**Theorem 1:** The following statements are equivalent:

(i)  $C(x_1, ..., x_d) = x_{[1]} \cdot \prod_{k=2}^d g_k(x_{[k]})$  is a copula.

(ii) For all  $0 < x < y \le 1$ ,  $(k, j) \in \mathbb{N}_0 \times \mathbb{N}$  with  $k + j \le d$  we have

$$\sum_{i=0}^{j} \binom{j}{i} (-1)^{i} \prod_{\ell=1}^{i} g_{\ell+k}(x) \prod_{\ell=i+1}^{j} g_{\ell+k}(y) \ge 0.$$

(iii) For  $m = 1, \ldots, d$  we have  $G_m \in \mathcal{D}$ , where

$$G_m(x) := \prod_{i=0}^{m-1} g_{d-m+1+i}^{(-1)^i \binom{m-1}{i}}(x), \quad x \in [0,1] \quad \text{(with } x = 0 \text{ as limit and } g_1(x) = x\text{)}.$$





## **Research questions**

(a) For which  $g_1, \ldots, g_d$  is the following a multivariate distribution function

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where  $x_{[1]} \leq x_{[2]} \leq \ldots \leq x_{[d]}$  is the ordered list of  $x_1, \ldots, x_d \in \mathbb{R}$ ?

#### (b) What about stochastic representations?

(c) Are there interesting examples / applications?





#### Theorem 2:

Let (i)–(iii) of Theorem 1 be valid. For each  $E \subset \{1, \ldots, d\}$  consider a r.v.  $Z_E$  s.t.

- the distribution function of  $Z_E$  equals  $G_{|E|}$ ,
- all  $Z_E$  are independent.

With an arbitrary univariate survival function  $\bar{g}_1$ , define the random vector  $(X_1, \ldots, X_d)$  by the following "exogenous shock model":

 $X_k := \min \{ \bar{g}_1^{-1}(Z_E) : k \in E \}, \quad k = 1, \dots, d.$ 

 $\Rightarrow$  The survival function of  $(X_1, \ldots, X_d)$  is given by  $C(\bar{g}_1(x_1), \ldots, \bar{g}_1(x_d))$ .

**Remark:** Conversely, any choice for the laws  $G_1, \ldots, G_d \in \mathcal{D}$  of the  $Z_E$  uniquely determines associated functions  $g_2, \ldots, g_d$ .





With  $\bar{g}_1$  a continuous survival function on  $(0, \infty)$  consider

 $X_k := \min \{ \bar{g}_1^{-1}(Z_E) : k \in E \}, \quad k = 1, \dots, d.$ 

- The  $Z_E$  are abs. continuous on (0, 1) with potential extra atom at 0.
- The  $\bar{g}_1^{-1}(Z_E)$  are abs. continuous on  $(0, \infty)$  with possible extra atom at  $\infty$ .
- $\bar{g}_1^{-1}(Z_E) = \text{arrival time point of exogenous shock killing all components in E.}$
- $X_k$  = first time point when a shock kills component k.
- **Example:**  $\bar{g}_1(x) = \exp(-x)$  and  $G_m(x) = x^{\lambda_m}$  for some  $\lambda_m > 0$   $\Rightarrow \bar{g}_1^{-1}(Z_E)$  exponential with rate  $\lambda_{|E|}$  $\Rightarrow g_k(x) = x^{a_k}$  for special sequences  $(a_2, \dots, a_d)$ .





**Schematic overview:** dimension = 2

 $g_1$   $g_2$   $\sim$   $Z_E, |E| = 1$  $g_1/g_2$   $\sim$   $Z_E, |E| = 2$ 





**Schematic overview:** dimension = 3

 $g_1$   $g_2$   $g_3$  ~  $Z_E, |E| = 1$  $g_1/g_2$   $g_2/g_3$  ~  $Z_E, |E| = 2$  $g_1g_3/g_2^2$  ~  $Z_E, |E| = 3$ 





**Schematic overview:** dimension = *d* 



Challenge: Dimensionality reduction?



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# Stochastic representations de Finetti representation





Let  $H = \{H_t\}_{t \in [t_0, t_1]}$  be an increasing additive process with  $H_{t_0} = 0$  and  $H_{t_1} = \infty$ , i.e.  $\Psi_t(x) := -\log (\mathbb{E}[\exp(-xH_t)])$  defines a family of Bernstein functions  $\{\Psi_t\}_{t \ge 0}$ .

- Draw one sample  $\mathcal{F}(\omega)$  from the random d.f.  $\mathcal{F} := \left\{1 e^{-H_t}\right\}_{t \in [t_0, t_1]}$ .
- Let  $X_1, X_2, \ldots \in (t_0, t_1)$  be i.i.d. random variables drawn from  $\mathcal{F}(\omega)$ .

 $\Rightarrow$  The univariate survival function of  $X_k$  is given by

$$\bar{g}_1(x) := \mathbb{P}(X_k > x) = \mathbb{E}\left[e^{-H_x}\right] = \exp\left(-\Psi_x(1)\right), \quad x \ge 0.$$

 $\Rightarrow$  The survival copula of  $(X_1, \ldots, X_d)$  has form *C* with  $g_2, \ldots, g_d$  given by

$$g_k(x) := \exp\left(-\Psi_{\bar{g}_1^{-1}(x)}(k) + \Psi_{\bar{g}_1^{-1}(u)}(k-1)\right), \quad k = 2, \dots, d.$$





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#### Examples Marshall–Olkin copulas

- Let  $H = \{H_t\}_{t \ge 0}$  be a Lévy subordinator. This means
  - $\rightarrow \Psi_t = t \cdot \Psi_1$  for some fixed Bernstein function  $\Psi_1$ .
  - $\rightarrow$   $g_k(x) = x^{a_k}$  for a sequence  $(1, a_2, \dots, a_d)$  being *d*-monotone.
  - $\rightarrow$  C in Th. 3 is the survival copula of an **exchangeable Marshall–Olkin law**.
- Th. 2 with  $\bar{g}_1(x) = \exp(-x)$  yields the [Marshall–Olkin (1967)] representation, in which arrival times of exogenous shocks are exponentially distributed.

#### • Proposition:

 $C(x_1, ..., x_d) = x_{[1]} \cdot \prod_{k=2}^d g_k(x_{[k]})$  is an extreme-value copula  $\Leftrightarrow$  it is the survival copula of an exchangeable Marshall-Olkin law.





## Examples Sato-frailty copulas

- Let  $H = \{H_t\}_{t \ge 0}$  be an increasing Sato process. This means
  - →  $\Psi_t(x) = \Psi_{sd}(x t^H)$  for some H > 0 and some fixed, self-decomposable Bernstein function  $\Psi_{sd} = \Psi_1$ .
  - → With  $\varphi := \exp(-\Psi_{sd})$  denoting the Laplace transform of the associated self-decomposable law on  $(0, \infty)$

$$C(x_1, \dots, x_d) = C_{\varphi}(x_1, \dots, x_d) = x_{[1]} \cdot \prod_{k=2}^d g_{k,\varphi}(x_{[k]}),$$
  
with  $g_{k,\varphi}(x) = \frac{\varphi(k \varphi^{-1}(x))}{\varphi((k-1) \varphi^{-1}(x))}.$ 

• Theorem 4: ("Kimberling-type" copula-characterization of SD laws)

 $C_{\varphi}$  is a copula for all  $d \ge 2$  $\Leftrightarrow \varphi$  is the Laplace transform of a self-decomposable law on  $(0, \infty)$ .





# Examples

Sato-frailty copula  $g_k(u) = \varphi(k \varphi^{-1}(u)) / \varphi((k-1) \varphi^{-1}(u))$ 

• Laplace exponent of Gamma-distributed r.v.:

$$\Psi_{\mathsf{sd}}(x) = \beta \, \log(1 + \frac{x}{\eta}), \quad x, \beta, \eta > 0.$$

• There is a (unique) increasing Sato process  $\{H_t\}_{t\geq 0}$  s.t.

$$\varphi(x) := \mathbb{E}[\exp(-xH_1)] = \exp(-\Psi_{\mathsf{sd}}(x)), \quad x \ge 0.$$

• The corresponding bivariate Sato-frailty copula  $C_{\varphi}$  is

$$C_{\varphi}(x_1, x_2) = \frac{x_{[1]}}{(2 - x_{[2]}^{1/\beta})^{\beta}}.$$





# Examples

Sato-frailty copula  $C_{\varphi}(x_1, x_2) = x_{[1]}/(2 - x_{[2]}^{1/\beta})^{\beta}$ 





## Examples The Dirichlet copula

For c > 0, let  $H^{(c)} = \{H_t^{(c)}\}_{t \in [0,1]}$  be specified via  $\{\Psi_t^{(c)}\}_{t \in (0,1)}$  as  $\Psi_t^{(c)}(x) := \int_0^\infty (1 - e^{-xu}) \frac{e^{u c (1-t)} - e^{-u c}}{u (1 - e^{-u})} du.$ 



- → The random d.f.  $\mathcal{F} = \mathcal{F}_c = \left\{1 e^{-H_t^{(c)}}\right\}_{t \in [0,1]}$  is a **Dirichlet process**.
- $\rightarrow$  The resulting copula in Th. 3 is called **Dirichlet copula**:

$$C(x_1,\ldots,x_d) = C_c(x_1,\ldots,x_d) = x_{[1]} \cdot \prod_{k=2} \frac{c x_{[k]} + k - 1}{c + k - 1}.$$

 $\rightarrow$  Kendall's  $\tau$ , Spearman's  $\rho_S$ , and tail-dependence are

$$\tau = \frac{2c+3}{3(c+1)^2}, \qquad \rho_S = \frac{1}{c+1}, \qquad \text{LTD}_C = \text{UTD}_C = \frac{1}{c+1}.$$

→ **Theorem 5: (Radially symmetric exogenous shock models)** The copula  $C(x_1, ..., x_d) = x_{[1]} \cdot \prod_{k=2}^d g_k(x_{[k]})$  is radially symmetric  $\Leftrightarrow C = C_c$  is a Dirichlet copula for some  $c \in [0, \infty]$ .





## Examples The Dirichlet copula



- Dirichlet process is known in non-parametric Bayesian statistics.
- It is an interesting model for a distorted random number generator.
- **[Ferguson (1973)]**: Simulation  $(X_1, \ldots, X_d)$  from Dirichlet copula  $C_c$  is easy:
  - $\rightarrow$  Simulate  $X_1 \sim \mathcal{U}(0, 1)$ .
  - $\rightarrow$  For k = 2, ..., d simulate  $X_k$  as follows:
    - (i) Simulate discrete random variable  $N \in \{1, ..., k\}$  with

$$\mathbb{P}(N=i) = \frac{1}{c+k-1}, \quad i = 1, \dots, k-1, \quad \mathbb{P}(N=k) = \frac{c}{c+k-1}.$$

(ii) If N = k simulate  $X_k \sim \mathcal{U}(0, 1)$ , else set  $X_k := X_N$ .



## References

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