



Stress tests for lapse risk: correlation and contagion among policyholders' behaviours

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A word on the lapse risk

What is the **lapse risk** ?

→ Due to death, maturity, change of premium level, **surrender**, ...

Why so much interest ?

- 1 Among the 3 major risks in life insurance ;
- 2 Understand the behaviours, and design new products ;
- 3 Predictions : segmentation and risk management (ALM).

Lapse classification in life insurance

The 2 main historical explanations for surrenders are ([Out90])

- *liquidity needs* → idiosyncratic → structural surrenders ;
- *economic distress* → environment → temporary surrenders.

Current context : never experienced such low interest rates ⇒ impact on the underwriting of new business...

Threat : massive (temporary) surrenders due to ↗ of interest rate.

- 1 Introduction to the problem
- 2 Regulation and current approaches in insurance companies
 - Structural surrenders and segments
 - Issues
 - QIS 5 and Solvency II recommendations
 - (Partial) internal model
- 3 The dynamic contagion process
- 4 Key messages, limits and on-going research

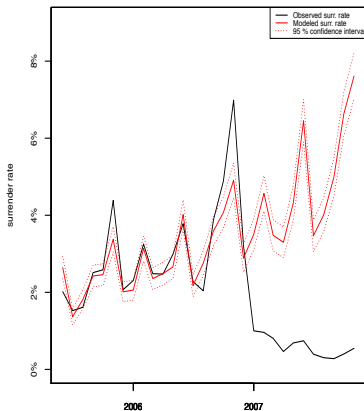
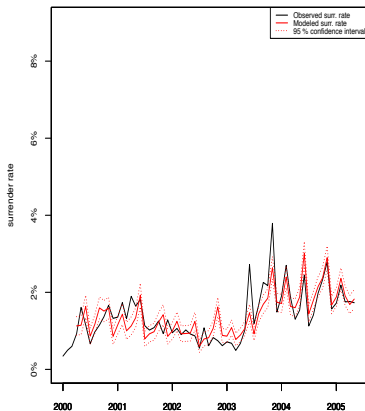
Estimation of structural lapses

Tables with profiles, e.g. yearly **structural** lapse rates (LR) for 9 segments :

	Bank	Agent	Direct
[0, 4] years	5%	8%	15%
[4, 8] years	3%	10%	15%
> 8 years	10%	19%	20%

- Segment the population with a priori discriminant risk factors.
- Empirical estimation / model-based estimates (GLM, ...).
- **Associated important assumptions...**

Model-based example : logit on spanish Endowments



Crisis **not captured** (despite financial covariates).

→ GLM / survival models **missed something....**

Regulation : QIS 5 (EIOPA)

Compute the SCR in 2 steps, and keep the max. b/w (1) and (2).

Step (1) : shocks applied to structural LR (misestimation).

$$LR_{up} = \min(100\%, 150\% \times LR) \rightarrow \text{our context !}$$

$$LR_{down} = \min(0, \max(50\% \times LR, LR - 20\%)).$$

Step (2) : mass lapse event, ~ “bank run”.

30%-loss of the sum of positive surrender strain over portfolio ;

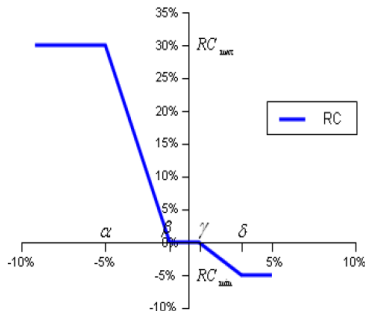
→ Empirics to calibrate mass lapse event is poor ;

→ Should be adjusted to the type of life insurance policy...

“Steps (1) AND (2) incorporate temporary lapses”

Internal model and practical approach (S-shaped)

Taux de rachats conjoncturels (RC) en fonction de l'écart entre le taux servi R et le taux concurrent TC



At the end...

$$LR_{shocked} = \min(1, \max(0, RS + RC)).$$

Still some issues to deal with

+ Pros :

- easy-to-understand, easy-to-implement,
- integrates (artificially) copycat behaviours → correlation risk ;

- Cons :

- not fully realistic ;
- this is a static model...(does not depend on time t),
- does not consider the contagion between policyholders...

⇒ We'd like to introduce a model that copes with both **correlation** and **contagion risks** to define extreme scenarios.

- 1 Introduction to the problem
- 2 Regulation and current approaches in insurance companies
- 3 The dynamic contagion process
 - An alternative to model contagion : Hawkes processes
 - Extended Hawkes processes : our context
 - Theoretical results
 - Risk management and sensitivities
- 4 Key messages, limits and on-going research

Intensity models - Hawkes process

Intensity models are often used in **mortgage prepayments** (most of them of Cox-type). **Focus here on Hawkes-type processes** [HO74].

Counting process s.t.

$$\lambda_t = \lambda_\infty + (\lambda_0 - \lambda_\infty)e^{-\beta t} + \alpha \int_0^t e^{-\beta(t-s)} dN_s,$$

- Path-dependent stochastic intensity ;
- Piecewise deterministic : “internal” source of excitation ;
- No correlation...

Dynamic contagion process for the lapse intensity

→ Extended Hawkes, [DZ11].

→ $(N_t)_{t \geq 0}$: counting process of lapses over the whole portfolio.

$$\lambda_t = \lambda_\infty + (\lambda_0 - \lambda_\infty) e^{-\beta t} + \sum_{i \geq 1} X_i e^{-\beta(t-T_i)} \mathbf{1}_{T_i \leq t} + \sum_{j \geq 1} Y_j e^{-\beta(t-\hat{T}_j)} \mathbf{1}_{\hat{T}_j \leq t} \quad (1)$$

where lapses occur at T_i , and X_i, Y_j are magnitudes of jumps.

- ① **Structural** surrender forces λ_0 , and λ_∞ (constant here),
- ② **Temporary** surrenders, with
 - endogenous shocks : **contagion**, internal ;
 - **exogenous shocks** : history of $\hat{N}_t \rightarrow$ dynamic dependence, source of **correlation** in our setting (to be defined later).

Cause of correlation : interest rate movements

→ Consider a contract with

- guaranteed return $R^g > 0$: minimum profitability,
- **credited rate** R_t^c : encompasses R^g + potential profit benefit.

At the contract inception, $R^g \simeq 0$ for most of life insurers in 2015.
Moreover, we have $R_0^c = R^g$.

→ Let $(r_t)_{t \geq 0}$ be the interest rate with GBM dynamics (μ, σ) .

Q : in critical scenarios, how the surrender decision could be affected by the level of r_t ?

Look at the following standardized spread :

$$RG_t^0 := \frac{r_t - R_0^c}{R_0^c}$$

→ Makes sense to \nearrow the propensity to lapse when $RG_t^0 \nearrow$;

→ Say that policyholders would exercise their option to surrender at time \hat{T}_1 being the first time RG_t^0 hits a constant barrier $B > 0$.

→ Assume that the company can then adjust the credited rate R_t^c depending on the interest rate level (to avoid massive lapses).

This defines the new standardized spread RG_t^1 , given by

$$RG_t^1 = \frac{r_t - R_{\hat{T}_1}^c}{R_{\hat{T}_1}^c} = \frac{r_t - r_{\hat{T}_1}}{r_{\hat{T}_1}}, \quad \hat{T}_1 \leq t < \infty.$$

Next adjustment will be operated as soon as $RG_t^1 = B$, and so on...

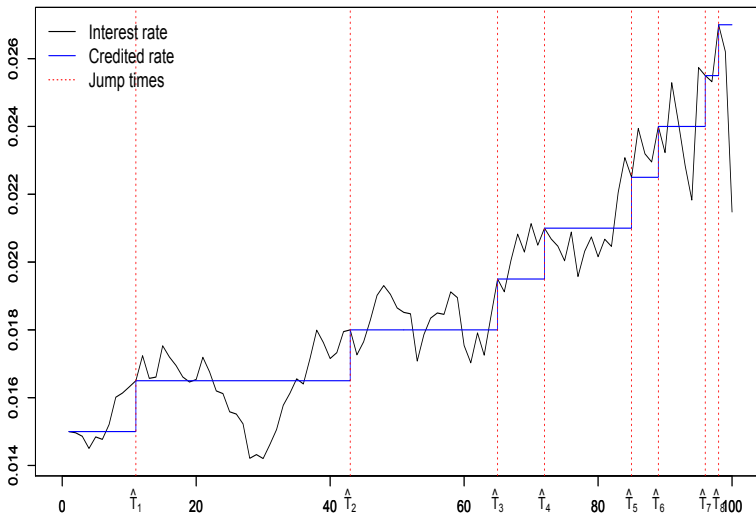
⇒ **These events thus characterize the sequence** $(\hat{T}_j)_{j=0,1,\dots}$ s.t.

$$\hat{T}_{j+1} = \inf\{t > 0, RG_t^j = B\}, \quad (2)$$

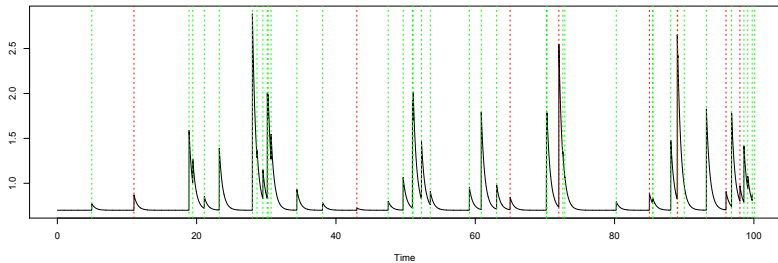
with $\hat{T}_0 = 0$ for convenience.

$(\hat{N}_t = \sum_{j \geq 1} 1_{\hat{T}_j \leq t}$: counting process associated to such events.)

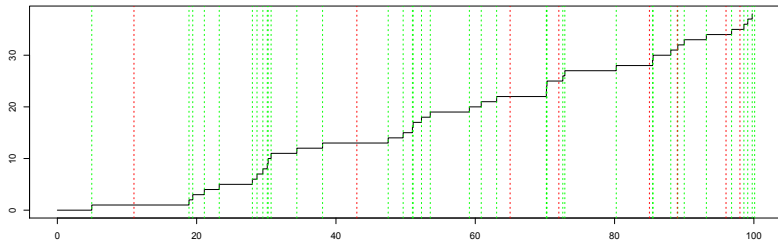
Adjustments of the credited rate



Dynamic contagion process: intensity process λ_t



Dynamic contagion process: counting process N_t



About external jumps at \hat{T}_j

→ $(r_t)_{t \geq 0}$ follows a GBM ($\log(r_t/r_0) = (\mu - \sigma^2/2)t + \sigma W_t$).

→ $(\hat{T}_j)_{j=0,1,\dots}$ are **hitting times** of the process RG_t .

⇒ The events \hat{T}_j **can also be characterized as follows**

$$\hat{T}_j = \hat{T}_{j-1} + \inf\{t \geq 0, \mu t + \sigma W_t = \log(1 + B)\}.$$

⇒ Inter-arrival times $\Delta \hat{T}_j = \hat{T}_j - \hat{T}_{j-1}$ are i.i.d., with **distribution F** ([SK91]) :

$$\Delta \hat{T}_j \sim \mathcal{IG}(\theta_1, \theta_2),$$

with $\theta_1 = 2 \log(1 + B)/(2\mu - \sigma^2)$, $\theta_2 = \log(1 + B)^2/\sigma^2$.

→ $\lambda_t, (\lambda_t, N_t), (\lambda_t, N_t, \widehat{N}_t)$ not Markovian.

→ $\hat{T}_j = \sum_{k=1}^j \Delta \hat{T}_k$.

Introduce $P(\hat{T}_j \leq t) = F^{j*}(t)$.

→ Denote by $h(t) = E[\hat{N}_t] = \sum_{j=0}^{\infty} P(\hat{N}_t \geq j)$, thus

$$h(t) = \sum_{j=0}^{\infty} F^{j*}(t). \quad (3)$$

The CDF F^{j*} is **still \mathcal{IG}** , with mean $j\theta_1$ and shape $j\theta_2$.

Moments of the lapse intensity

→ Let

$$m(t, \theta) = E[e^{\theta \lambda_t}],$$

and $m^{(n)}(t, \theta)$: n^{th} derivative of m with respect to θ .

We have

$$m^{(n)}(t, 0) = E[\lambda_t^n].$$

→ Denote respectively $\xi(t, \theta)$ and $\widehat{\xi}(t, \theta)$ the m.g.f. of

$$Z_t = \sum_{i=1}^{N_t} X_i e^{\beta T_i} \quad \text{and} \quad \widehat{Z}_t = \sum_{j=1}^{\widehat{N}_t} Y_j e^{\beta \widehat{T}_j}. \quad (4)$$

Z_t, \widehat{Z}_t are **discounted compound renewal processes** ([LGFW10]).

Similarly, $\xi^{(n)}(t, \theta)$ and $\widehat{\xi}^{(n)}(t, \theta)$ refer to the n^{th} derivative $\xi(t, \theta)$ and $\widehat{\xi}(t, \theta)$ with respect to θ .

→ λ_t can be written in the following form

$$\lambda_t = (\lambda_\infty + (\lambda_0 - \lambda_\infty)e^{-\beta t}) + e^{-\beta t}Z_t + e^{-\beta t}\widehat{Z}_t.$$

⇒ We can then derive

- 1 the m.g.f. of Z_t and \widehat{Z}_t ;
- 2 the m.g.f. of λ_t in function of those of Z_t and \widehat{Z}_t ;

⇒ At the end, we obtain a recursive formula.

(1) Moment generating functions of Z_t and \widehat{Z}_t

The m.g.f. ξ and $\widehat{\xi}$ of Z_t and \widehat{Z}_t are given by **recursive formulas** :

$$E[e^{\theta Z_t}] = \xi(t, \theta) = \dots + \int_0^t \dots \xi(t-u, \theta e^{\beta u}) m^{(1)}(u, 0) du \quad (5)$$

$$E[e^{\theta \widehat{Z}_t}] = \widehat{\xi}(t, \theta) = \dots + \int_0^t \dots \widehat{\xi}(t-u, \theta e^{\beta u}) dh(u), \quad (6)$$

- We can derive the moments of the renewal processes ;
- The first moment of the intensity λ_t is key (self-excited) ;
- Recall that $h(t) = E[\widehat{N}_t] = \sum_{j=0}^{\infty} F^{j*}(t)$.

(2) Moment generating function of λ_t

Proposition. For $n > 1$, the n^{th} derivative of the surrender intensity m.g.f. is given **recursively** :

$$\begin{aligned} m^{(n)}(t, \theta) &= K(t, \lambda_0, \lambda_\infty) m^{(n-1)}(t, \theta) \\ &\quad + \sum_{i=0}^{n-1} G(i, n) \left(l_i(t, \theta) + \widehat{l}_i(t, \theta) \right) m^{(i)}(t, \theta), \end{aligned} \tag{7}$$

with l_k and J_k for $\{k = 1, 2, \dots\}$ given by

$$\begin{aligned} l_k(t, \theta) &= l_{k-1}^{(1)}(t, \theta) + H(l_{k-1}(t, \theta)) \xi^{(1)}(t, \theta e^{-\beta t}), \\ \widehat{l}_k(t, \theta) &= \widehat{l}_{k-1}^{(1)}(t, \theta) + H'(\widehat{l}_{k-1}(t, \theta)) \widehat{\xi}^{(1)}(t, \theta e^{-\beta t}). \end{aligned}$$

Application : expected intensity process

→ The expectation $E[\lambda_t]$ is given by

$$m^{(1)}(t, 0) = \left(\lambda_0 - \frac{\beta \lambda_\infty}{\beta - 1/\gamma} \right) e^{-(\beta - \frac{1}{\gamma})t} + \frac{\beta \lambda_\infty}{\beta - 1/\gamma} + \frac{1}{\delta} \int_0^t e^{-(\beta - \frac{1}{\gamma})(t-s)} h'(s) ds.$$

Remark : $m^{(1)}(t, 0)$ comprises an infinite series associated with the external jumps component ($h'(s) = \sum_{j=0}^{\infty} f^{j*}(s)$).

Trick to get closed-form expressions : convolution of exponential and inverse gaussian r.v.

$$\rightarrow E[N_t] = E \left[\int_0^t \lambda_s ds \right] = \int_0^t m^{(1)}(s, 0) ds.$$

Limiting behaviour of the lapse intensity

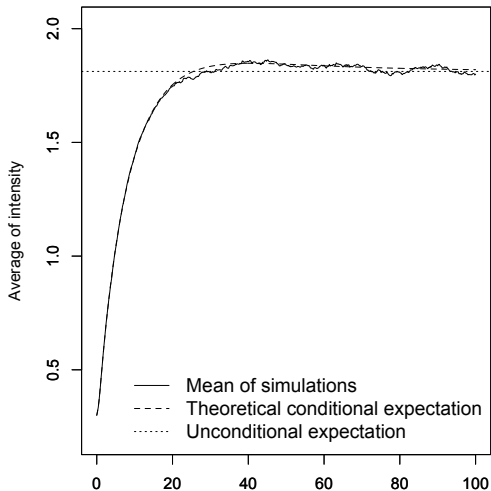
We can also compute the limit of this expectation :

$$\lim_{t \rightarrow \infty} E[\lambda_t] = \frac{\beta \lambda_\infty}{\beta - 1/\gamma} + \frac{1}{\delta \theta_1 (\beta - 1/\gamma)}. \quad (8)$$

→ The limiting behavior of the lapse intensity first moment strongly depends on the limit of the last term in the previous result.

→ Serfozo [2009] for such results.

Mean intensity process (# simu : 20 000)



Application to risk management - calibration

Some parameters can be calibrated from practitioners' knowledge :

- λ_0 (initial force of lapse) is a **constant**.
⇒ Exponential underlying lifetime distribution before lapse.
- λ_∞ can be fixed by the risk managers as their **goal**...
⇒ When the time horizon is given, this can be easily fixed.
- B depicts the sensitivity of PH to opportunities (experts).

Some parameters (e.g. GBM) should be calibrated from empirical data / history whenever possible.

Others relate to the management : β (ability to reassure the PH), γ, δ tie in with the mean size (SI) of lapsed contracts...

Stress tests : comparison with SII and S-shaped

- **Within the Solvency II framework** : run-off, 1-year horizon.
- **With regard to financial context** : focus on the upper-shock.
- **Risk measures** under consideration : VaR and $TVaR$.

	Solvency II Standard formula		S-shaped curve (ONC)		Hawkes counting process			Dynamic contagion process			
Parameters	Risk level	Shocks	Risk level	Shocks	$E[N_t]$	VaR_α	$TVaR_\alpha$	$E[N_t]$	VaR_α	$TVaR_\alpha$	
B	10%							455	1028	1142	
	30%	75	112	75	375	291	776	837	312	818	930
	50%								293	778	886
δ	0.1								2461	4286	4559
	0.5	75	112	75	375	291	776	837	702	1460	1594
	1.5								455	1028	1142

TABLE: Impact of contagion and correlation on $VaR_\alpha(N_t)$, $TVaR_\alpha(N_t)$ at level $\alpha = 99.5\%$, in a 1-year time horizon ($t = 250$).

Conclusion on stress tests

- The shock in SII looks neither consistent nor realistic.
- Stress tests in most of companies seem to be **underestimated**.
- OK for extreme scenarios (reserving), not so realistic in classical regime (pricing).
- PH' sensitivity to IR movements is obviously not linear...
- External component has a limited impact, provided that mean size of the external jumps is low ⇒ portfolio composition is crucial !

Key messages

Integrate only main risk factors + **correlation** + **contagion**.

Perspectives :

- ① Calibration on a real-life portfolio ;
- ② Use a martingale approach to retrieve the whole distribution of lapses N_t ,
- ③ Extend this approach with an adapted interest rate model.

References



Angelos Dassios and Hongbiao Zhao, *A dynamic contagion process*, Advances in Applied Probability **43** (2011), no. 3, 814–846.



Alan G Hawkes and David Oakes, *A cluster process representation of a self-exciting process*, Journal of Applied Probability (1974), 493–503.



G. Lévêillé, J. Garrido, and Y. Fang Wang, *Moment generating functions of compound renewal sums with discounted claims*, Scandinavian Actuarial Journal **2010** (2010), no. 3, 165–184.



Jean François Outreville, *Whole-life insurance lapse rates and the emergency fund hypothesis*, Insurance : Mathematics and Economics **9** (1990), 249–255.



SE Shreve and I Karatzas, *Brownian motion and stochastic calculus*, Newyork Berlin. Heidelberg. London Paris Tokyo (1991).