

# Bayesian estimation of the intensity function for Hawkes processes

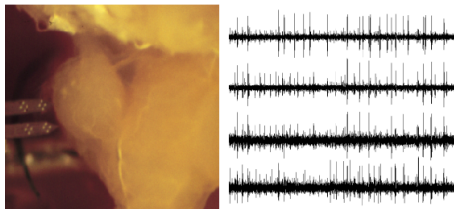
Luminy

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# Spikes and point processes

- ▶ neuronal activity : electric impulses = spikes



From C Pouzat, CNRS UMR 8118

- ▶ The interesting information from these spikes is related to the time of their appearance and their length (rather than their shape or intensity)
- ⇒ Point processes On  $\mathbb{R}^+$ .

# Poisson processes and Hawkes processes

On  $(\Omega, \mathcal{F}, \mathbb{P})$ ,  $\mathbf{N}$  is a point process

- ▶ Its distribution is characterized by its conditional intensity process

$$\begin{aligned}\lambda(t) &= \lambda(t|\mathcal{F}_t) = \lim_{h \rightarrow 0} \frac{1}{h} \mathbb{P}(\text{an event takes place in } [t, t+h] | \mathcal{F}_t) \\ &= \text{"}\mathbb{E}[dN_t | \mathcal{F}_t]\text{"} \quad (\mathcal{F}_t) = \text{adapted filtration}\end{aligned}$$

- ▶ **Poisson Process**

- ▶  $\lambda(t)$  deterministic
- ▶ intervals between events are independent

- ▶ **Hawkes Process**

- ▶  $\lambda(t) = \left\{ \nu + \sum_{T_i < t} h(t - T_i) \right\}_+$
- ▶ When  $h$  is positive : self-excitation, when  $h$  is negative inhibition

# Multidimensional Hawkes processes

- ▶  $M$  neurones interacting : self or inter excitation or inhibition. We observe  $M$  point processes non independent.
- ▶ Conditional intensity of the process  $m$  :

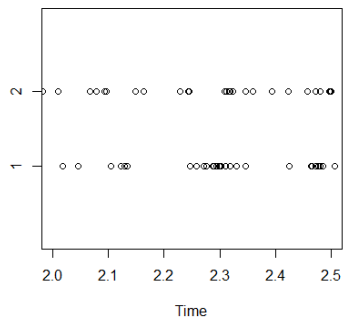
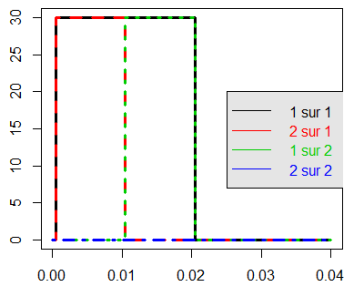
$$\lambda^{(m)}(t) = \left\{ \nu^{(m)} + \sum_{\ell=1}^M \sum_{t_i^{(\ell)} < t} h_{\ell}^{(m)}(t - t_i^{(\ell)}) \right\}_+$$

## ▶ Remarks

- ▶  $h_{\ell}^{(m)}$  is the excitation function of  $m$  by  $\ell$ .
- ▶  $h_{\ell}^{(m)}$  can be negative.
- ▶ Stationary process if the spectral radius of matrix  $I = \left( \int_0^{\infty} |h_{\ell}^{(m)}(u)| du \right)_{1 \leq \ell, m \leq M}$  is  $< 1$ .
- ▶ The functions  $h_{\ell}^{(m)}$  have support :  $[0, s_{max}]$

[Reynaud-Bouret et al., 2013]

# Example of bi-dimensionnal Hawkes processes



# Objectives

- ▶ Bayesian estimation  $\theta = \left\{ \left\{ \nu^{(m)} \right\}_{m=1 \dots M}, \left\{ h_{\ell}^{(m)} \right\}_{(\ell, m) \in \{1, \dots, M\}^2} \right\}$ .
- ▶ exists : Lasso estimates [Hansen et al., 2014] .
- ▶ Hardly anything in the Bayesian framework [Rasmussen, 2013], K. Heller

# Bayesian inference

- ▶  $\theta = \left\{ \left\{ \nu^{(m)} \right\}_{m=1 \dots M}, \left\{ h_{\ell}^{(m)} \right\}_{(\ell, m) \in \{1, \dots, M\}^2} \right\}$
- ▶ Observations  $\mathbf{N}$  on  $M$  neurones during  $[0, T]$
- ▶ **Likelihood** :  $\mathbf{n}_m$  = number of jumps for neuron  $m$ , and  $T_1^{(m)}, \dots, T_{\mathbf{n}_m}^{(m)}$  jump times for neurone  $m$ ,  $m \leq M$

$$L(\mathbf{N}; \theta) = \exp \left[ \sum_{m=1}^M \sum_{j=1}^{\mathbf{n}_m} \log \lambda^{(m)}(T_j^{(m)}) - \int_0^T \lambda^{(m)}(u) du \right]$$

- ▶ Posterior (pseudo) :

$$\pi(d\theta | \mathbf{N}) = \frac{\pi(d\theta) L(\mathbf{N}; \theta)}{p(\mathbf{N})}$$

- ▶ Depends on  $\mathbf{N}_{[-s_{\max}: T]}$

## Posterior concentration on $\theta : \mathcal{T} \rightarrow +\infty$

- **Posterior concentration** If  $\theta^*$  = true parameter,  $\epsilon_{\mathcal{T}} = o(1)$

$$\mathbb{E}_{\theta^*} [\mathbb{P}(d(\theta, \theta^*) > \epsilon_{\mathcal{T}} | \mathbf{N})] = o(1)$$



# General Theorem on posterior concentration rates G& VdV (07)

## ► Kullback- Leibler Condition

$$\mathbb{P}(\{K(p_{\theta^*, T}, p_{\theta, T}) \leq T\epsilon_T^2; V(p_{\theta^*, T}, p_{\theta, T}) \lesssim T\epsilon_T^2\}) > e^{-cT\epsilon_T^2}$$

## ► Tests $\exists \Theta_T \pi(\Theta_T^c) \leq e^{-(c+2)T\epsilon_T^2}$ s.t.

$$\exists \phi_T; \mathbb{E}_{\theta^*}(\phi_T) = o(1); \quad \sup_{\theta \in \Theta_T^c; d(\theta, \theta^*) > B\epsilon_T} \mathbb{E}_{\theta}(1 - \phi_T) \leq e^{-(c+2)T\epsilon_T^2}$$

Then

$$\mathbb{P}(d(\theta, \theta^*) > M\epsilon_T | \mathbf{N})$$

# Application to Hawkes processes

- ▶ Parameter space :

$$\Theta = \{(\nu_\ell, h_{k,\ell}, k, \ell \leq M); h_{k,\ell} \geq 0, \nu_\ell > 0, \underbrace{\|\rho\| < 1}_{\text{For stationarity}} \}$$

$$\rho = (\rho_{k,\ell})_{k,\ell \leq M}, \quad \rho_{k,\ell} = \int h_{k,\ell}(x) dx$$

- ▶ Metrics :  $L_1$  and  $L_1$  stochastic

$$d_{1,T}(\theta, \theta^*) = \int_0^T \sum_\ell |\lambda_{\ell,\theta} - \lambda_{\ell,\theta^*}|(t) dt,$$

$$\|\theta - \theta^*\|_1 = \sum_\ell |\nu_\ell - \nu_\ell^*| + \sum_{\ell,k} \|h_{k,\ell} - h_{k,\ell}^*\|_1$$

$$\lambda_{\ell,\theta}(t) = \nu_\ell + \sum_{k=1}^M \int_{t-s_{\max}}^{t^-} h_{k,\ell}(t-u) dN_u^k$$

- ▶ Observations  $\mathbf{N}_{[-s_{\max}, T]} = (N_t^\ell, t \leq T, \ell \leq M)$ .

## Concentration in $d_{1,T}$

- True  $\theta^* = (\nu^*, h^*)$  with  $\nu^* > 0$  and  $\|\rho^*\| < 1$ . If

- ▶ **KL** :  $\exists c > 0$  such that

$$\Pi(|\nu - \nu^*| < \epsilon_T; \max_{\ell,k} \|h_{\ell,k} - h_{\ell,k}^*\|_2 \leq \epsilon_T (\log \log T)^{-1/2}) > e^{-cT\epsilon_T^2}$$

- ▶ **Sieve** :  $\exists \Theta_T \subset \Theta$ ,  $\Pi(\Theta_T^c) = o(e^{-CT\epsilon_T^2})$

- ▶ **Entropy** :

$$\log N(\epsilon_T, \Theta_T, \|\cdot\|_1) \lesssim T\epsilon_T^2$$

Then

$$\mathbb{E}^* [\pi(d_{1,T}(\theta^*, \theta) > M\epsilon_T | \mathbf{N})] \lesssim \frac{1}{T\epsilon_T^2}$$

## How does it work : (1) concentration in $d_{1,T}$

- Prove that

$$\forall \theta \in \{|\nu - \nu^*| < \epsilon_T; \max_{\ell,k} \|h_{\ell,k} - h_{\ell,k}^*\|_2 \leq \epsilon_T / \log \log T\}$$

$$KL_T(\theta^*, \theta) \leq \kappa T \epsilon_T^2, \quad \mathbb{P}^* (\ell_T(\theta) - \ell_T(\theta^*) < -(\kappa + 1) T \epsilon_T^2) \leq \frac{\log T}{T \epsilon_T^2}$$

- tests :  $\theta_1 \in \Theta_T$ ,  $S_j = \{\theta, d_{1,T}(\theta^*, \theta) \in (j\epsilon_T, (j+1)\epsilon_T)\}$ .  
 $A_1 = \{t; \lambda_{\theta_1} \geq \lambda_{\theta^*}(t)\}$ ,

$$\phi_{\theta_1} = \max_{\ell} (\mathbb{I}\{N^{\ell}(A_1) - \Lambda^{\ell}(A_1; \theta^*) \geq jT\epsilon_T/8\} \vee \mathbb{I}\{N^{\ell}(A_1^c) - \Lambda^{\ell}(A_1^c; \theta^*) \geq jT\epsilon_T/8\})$$

then

$$\mathbb{E}^* (\mathbb{I}_{\Omega_T} \phi_{\theta_1}) + \sup_{\|\theta - \theta_1\|_1 \leq j\epsilon_T} \mathbb{E}^* [\mathbb{E}_{\theta} (\mathbb{I}_{\Omega_T} \mathbb{I}_{\theta \in S_j} (1 - \phi_{\theta_1}) | \mathcal{G}_0)] \leq 2e^{-x_0 T j \epsilon_T [1 \wedge j \epsilon_T]}$$

## weak conditions - examples of priors

$$\rho \sim \pi_\rho; \quad \bar{h}_{k,\ell} = h_{k,\ell}/\rho_{k,\ell} \stackrel{iid}{\sim} \pi_h, \quad \nu_\ell \stackrel{iid}{\sim} \pi_\nu$$

### ► Histograms

$$\bar{h}_{k,\ell} = \delta_{k,\ell} g_{k,\ell}, \quad g_{k,\ell} \stackrel{iid, \mathcal{D}}{=} g, \quad \delta_{k,\ell} \stackrel{iid}{\sim} Be(p)$$

and  $g$  has distribution given by

$$g = \sum_{j=1}^J \frac{z_j w_j}{|I_j|}, \quad I_j = (s_j, s_{j+1}), \quad z_j \stackrel{iid}{\sim} Be(q), \quad \sum_{j:z_j=1} w_j = 1$$

$$J-1 \sim \mathcal{P}(a), \quad (w_j, j \in \{z_j = 1\}) \sim \mathcal{D}(\alpha_J, \dots, \alpha_J), \\ (s_1, \dots, s_J) \sim \Pi_s$$

If  $\bar{h}^*$  are Hölder  $\beta$  then

$$\epsilon_T \lesssim (T/\sqrt{\log \log T \log T})^{-\beta/(2\beta+1)}, \quad \beta \leq 1$$

# Mixtures of Betas : Adaptive estimation over Hölder classes

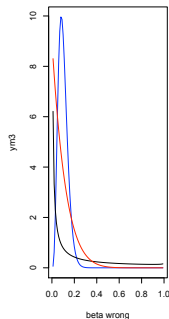
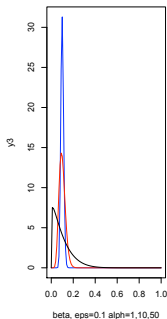
## ► Prior model

$$h_{k,\ell} = \left( \int_0^1 g_{\alpha,\epsilon} dM(\epsilon) \right)_+, \quad g_{\alpha,\epsilon}(x) = \text{Beta} \left( \frac{\alpha}{1-\epsilon}, \frac{\alpha}{\epsilon} \right)$$

$M$  : signed bounded measure on  $[0, 1]$ .

## ► Why?

$$E \left[ \text{Beta} \left( \frac{\alpha}{1-\epsilon}, \frac{\alpha}{\epsilon} \right) \right] = \epsilon, \quad \text{var} \left[ \text{Beta} \left( \frac{\alpha}{1-\epsilon}, \frac{\alpha}{\epsilon} \right) \right] = \frac{\epsilon(1-\epsilon)}{\alpha}$$



# Why positive part of mixtures of Betas ?

- ▶ (R. 08) :  $\forall f \in \mathcal{H}(\beta, L)$ , there exists  $f_1$  such that

$$\|g_{\alpha, f_1} - f\|_{\infty} = O(\alpha^{-\beta/2}), \quad g_{\alpha, f_1}(x) = \int_0^1 g_{\alpha, \epsilon}(x) f_1(\epsilon) d\epsilon$$

and  $\exists P = \sum_{j=1}^J w_j \delta_{(\epsilon_j)}$ ,  $J \lesssim \sqrt{\alpha}(\log \alpha)^{3/2}$

$$\|g_{\alpha, f_1} - g_{\alpha, P}\|_2 = O(\alpha^{-\beta/2})$$

- ▶ Prior

$$J \sim \mathcal{P}(a), \quad (|w_1|, \dots, |w_J|) \sim \mathcal{D}(\alpha/J, \dots, \alpha/J), \quad \sqrt{\alpha} \sim \Gamma(a_0, b_0)$$

with

$$w_j = \zeta_j |w_j|, \quad \zeta_j \in \{-1, 1\} \quad iid, \quad \epsilon_j \stackrel{iid}{\sim} \text{Beta}(a_1, b_1)$$

- ▶ Posterior concentration rate : **adative**

$$\epsilon_T \lesssim T^{-\beta/(2\beta+1)} (\log T)^q$$

Concentration in  $L_1$  :

$$\|\theta - \theta_0\|_1 = \sum_{\ell} |\nu_{\ell} - \nu_{\ell}^*| + \sum_{\ell,k} \|h_{\ell,k} - h_{\ell,k}^*\|_1$$

$$d_{1,T}(\theta, \theta^*) = \int_0^T |\lambda_{\theta}(t) - \lambda_{\theta^*}(t)| dt : \text{non explicit}$$

### Theorem

*Under the same assumptions and if*

$$\mathbb{P}(\|\rho\| > 1 - u_0 \epsilon_T) \leq e^{-cT\epsilon_T^2}$$

*then  $\exists M_0 > 0$*

$$\mathbb{E}^* [\mathbb{P}(\|\theta - \theta^*\|_1 > M_0 \epsilon_T | \mathbf{N})] = o(1)$$



From  $d_{1,T}$  to  $L_1$

$$\begin{aligned} & \mathbb{E}^* [\mathbb{P}(\{d_{1,T}(\theta, \theta^*) \leq \epsilon_T\} \cap \{\|\theta - \theta^*\|_1 > M_{0\epsilon_T} | \mathbf{N}\})] \\ & \leq \mathbb{P}^* (D_T < e^{-cT\epsilon_T^2}) + e^{cT\epsilon_T^2} \int_{\|\theta - \theta^*\|_1 > M_{0\epsilon_T}} \mathbb{E}^* [\mathbb{P}_\theta [d_{1,T}(\theta^*, \theta) \leq \epsilon_T | \mathcal{G}_0]] d\Pi(\theta) \end{aligned}$$

## Some consequences

- Point estimator  $\hat{f} = E[f|\mathbf{N}]$  satisfies

$$\mathbb{P}_{\theta^*} \left( \|f^* - \hat{f}\|_1 \lesssim \epsilon_T \right) \rightarrow 1$$

- *Easy conditions* Compared to Hansen et al. – who had a stronger condition on the type dictionary

# Simulation study

- **Prior** : random histogram

$$h_{\ell,k} = \underbrace{\delta_{\ell,k}}_{\text{Be}(1/2)} \sum_{j=1}^{M^{\ell,k}} \mathbb{I}_{I_j^{\ell,k}} \frac{w_j^{\ell,k}}{|I_j^{\ell,k}|}, \quad M^{\ell,k} \stackrel{iid}{\sim} 1 + \mathcal{P}(a)$$

$$\log \nu_{\ell} \sim \mathcal{N}(3, 1), \quad w_j^{\ell,k} \sim \text{Be}(1/2) \times \ln \mathcal{N}(\mu_{\alpha}, s_{\alpha}^2)$$

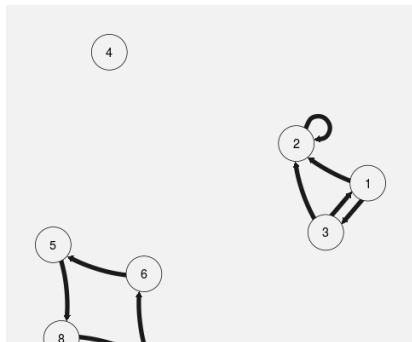
## Scenarios

- ▶  $\nu_\ell = 20$
- ▶ *Scenario 1* : We first consider  $K = 2$  neurons and piecewise constant interactions :

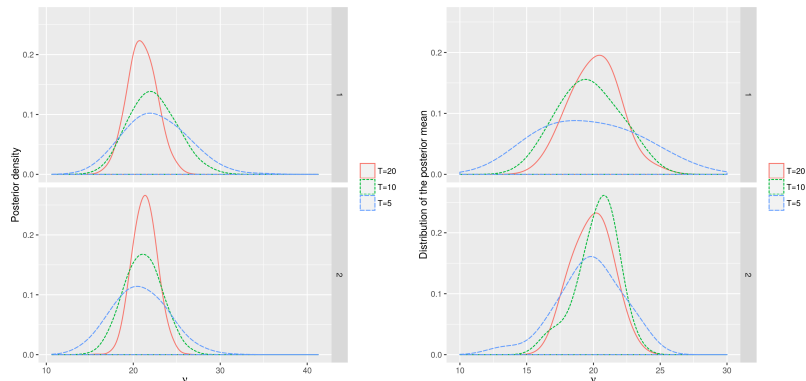
$$h_{1,1} = 30 \cdot \mathbb{I}_{(0,0.02]}, \quad h_{2,1} = 30 \cdot \mathbb{I}_{(0,0.01]}, \quad h_{1,2} = 30 \cdot \mathbb{I}_{(0.01,0.02]}, \quad h_{2,2} = 0.$$

- ▶ *Scenario 2* : In this scenario, we mimic  $K = 8$  neurons belonging to three independent groups. The non-null interactions are the piecewise constant functions defined as :

$$h_{2,1} = h_{3,1} = h_{2,2} = h_{1,3} = h_{2,3} = h_{8,5} = h_{5,6} = h_{6,7} = h_{7,8} = 30 \cdot \mathbb{I}_{(0,0.02]}.$$

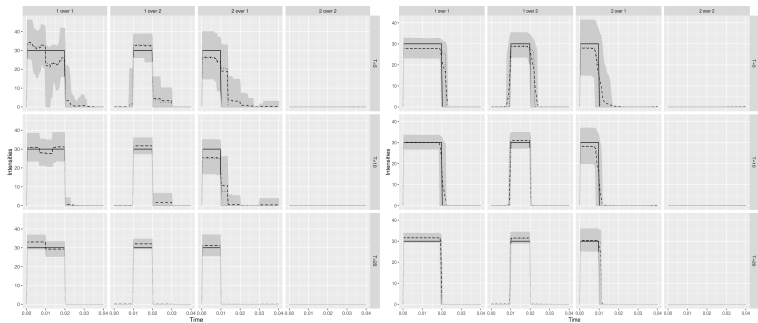


## Results Scenario 1 : $K=2 - \nu$



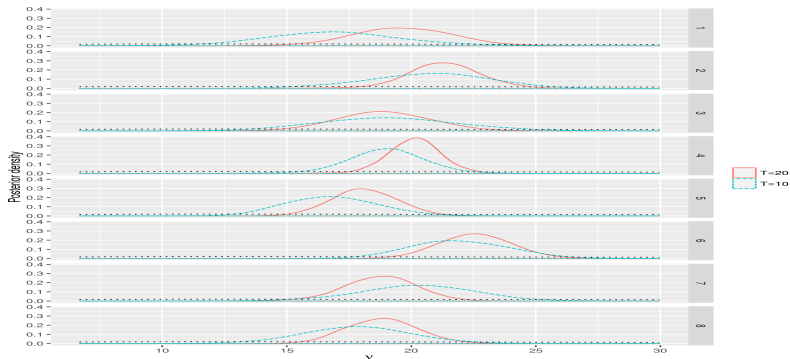
**Figure: Results for scenario 1.** *On the left*, posterior distribution of  $(\nu_1, \nu_2)$  with  $T = 5$ ,  $T = 10$  and  $T = 20$  for one dataset. *On the right*, distribution of the posterior mean of  $(\nu_1, \nu_2)$   $\left( \widehat{\mathbb{E}} \left[ \nu_k | (N_t^{sim})_{t \in [0, T]} \right] \right)_{sim=1 \dots 25}$  over the 25 simulated datasets.

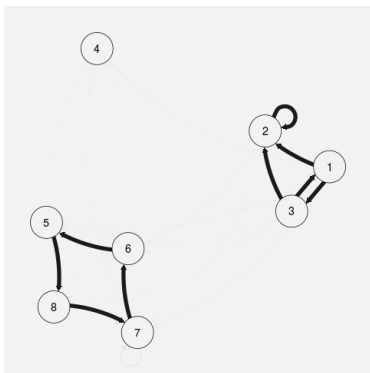
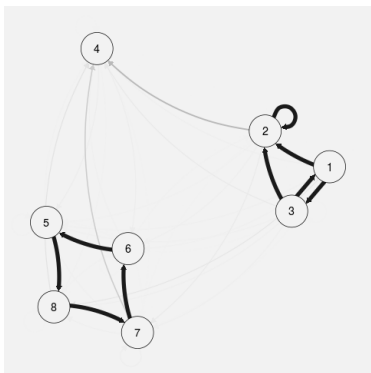
# Results Scenario 1 : $K=2 - h_{k,\ell}$



**Figure: Scenario 1,  $K=2$ .** Estimation of the  $(h_{\ell,k})_{\ell,k=1,2}$  using the regular prior (left) continuous prior (right). The gray region indicates the credible region for  $h_{\ell,k}(t)$  (delimited by the 5% and 95% percentiles of the posterior distribution). The true  $h_{\ell,k}$  is in plain line, the posterior expectation and posterior median for  $h_{\ell,k}(t)$  are in dotted and dashed lines respectively.

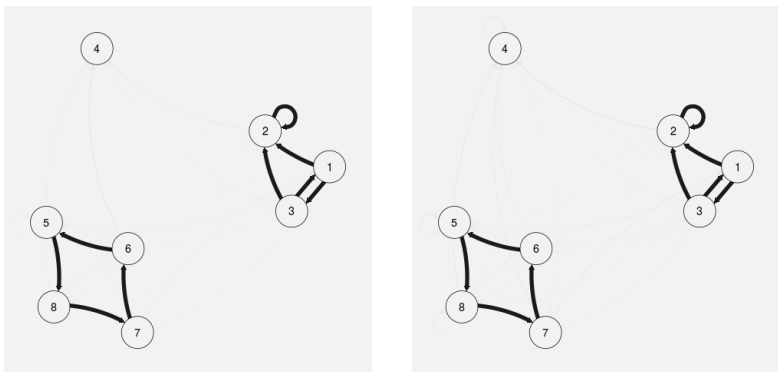
# Results for Scenario 2 : $K=8-\nu_\ell$





**Figure:** Results for scenario 2 for one given dataset. Posterior estimation of the interaction graph for  $T = 10$  on the left and  $T = 20$  on the right, for one randomly chosen dataset. Level of grey and width of the edges proportional to the posterior estimated probability of  $\hat{\mathbb{P}}(\delta_{\ell,k} = 1 | (N_t^{sim})_{t \in [0, T]})$ .





**Figure:** Results for scenario 2 over the 25 simulated datasets. Posterior estimation of the interaction graph for  $T = 10$  on the left and  $T = 20$  on the right. Level of grey and width of the edges are proportional to the posterior estimated probability of  $\frac{1}{25} \sum_{sim=1}^{25} \hat{\mathbb{P}}(\delta_{\ell,k} = 1 | (N_t^{sim})_{t \in [0, T]})$ .

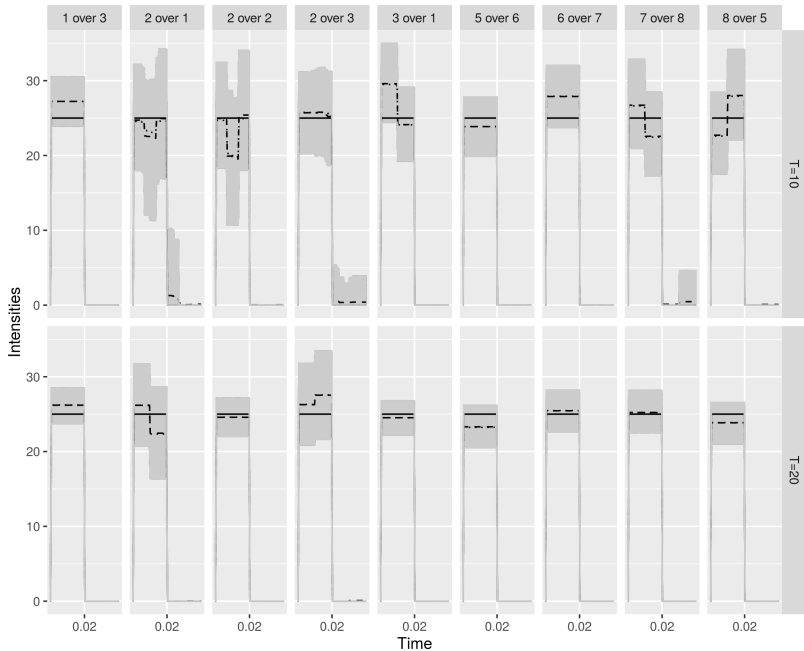


Figure: Results for scenario 2 for one given dataset. Estimation of the non null



# Conclusion

- ▶ Theory for  $L_1$  posterior concentration rates under quite weak assumptions
- ▶ Only for non negative  $h_{k,\ell}$
- ▶ Simulations : too slow for the moment to treat many neurons. Ok for  $\leq 10$  (just) – How about the mixture of Beta or other priors?
- ▶ Understanding credible regions

## Thank You



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