Télécom ParisTech CNRS

Universality of the limiting spectral distribution for matrices with correlated entries

$Marwa \; BANNA \\ \texttt{joint work with F. Merlevêde and M. Peligrad}$

Journées du Gdr Analyse Fonctionnelle, Harmonique et Probabilités

3 December 2015



Let X₁,..., X_ρ ∈ ℝ^N be i.i.d centered random vectors with covariance matrix Σ_N = ℝ(X₁X₁^T) = ... = ℝ(X_ρX_ρ^T).



- ► Let $\mathbf{X}_1, \ldots, \mathbf{X}_p \in \mathbb{R}^N$ be i.i.d centered random vectors with covariance matrix $\Sigma_N = \mathbb{E}(\mathbf{X}_1\mathbf{X}_1^T) = \ldots = \mathbb{E}(\mathbf{X}_p\mathbf{X}_p^T)$.
- ► The sample covariance matrix **B**_{N,p} is defined by

$$B_{N,p} = \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{T}$$

- Let X₁,..., X_ρ ∈ ℝ^N be i.i.d centered random vectors with covariance matrix Σ_N = ℝ(X₁X₁^T) = ... = ℝ(X_ρX_ρ^T).
- ► The sample covariance matrix **B**_{N,p} is defined by

$$B_{N,p} = \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{T}$$

- $\mathbb{E}(\mathbf{B}_{N,p}) = \Sigma_N$
- ► For fixed *N*, the strong law of large numbers implies

$$\lim_{p \to +\infty} \mathbf{B}_{N,p} = \lim_{p \to +\infty} \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{T} = \Sigma_{N} \quad \text{a.s.}$$

- Let X₁,..., X_ρ ∈ ℝ^N be i.i.d centered random vectors with covariance matrix Σ_N = ℝ(X₁X₁^T) = ... = ℝ(X_ρX_ρ^T).
- ► The sample covariance matrix **B**_{N,p} is defined by

$$B_{N,p} = \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{T}$$

- ► $\mathbb{E}(\mathbf{B}_{N,p}) = \Sigma_N$
- For fixed N, the strong law of large numbers implies

$$\lim_{p \to +\infty} \mathbf{B}_{N,p} = \lim_{p \to +\infty} \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{T} = \Sigma_{N} \quad \text{a.s.}$$

▶ What happens once $N := N_p, p \to \infty$ s.t. $p \to \infty$?

Introduction Sample covariance matrices

► Let
$$\mathbf{B}_{N} := \mathbf{B}_{N,p} = \frac{1}{p} \mathcal{X}_{N,p} \mathcal{X}_{N,p}^{T} = \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{T}$$

$$\mathcal{X}_{N,p} = \begin{pmatrix} X_{1,1} & X_{1,2} & \dots & X_{1,p} \\ \vdots & \vdots & & \vdots \\ X_{N,1} & X_{N,2} & \dots & X_{N,p} \end{pmatrix}$$
$$\mathbf{X}_{1} \qquad \mathbf{X}_{2} \qquad \dots \qquad \mathbf{X}_{p}$$

Introduction Sample covariance matrices

► Let
$$\mathbf{B}_{N} := \mathbf{B}_{N,p} = \frac{1}{p} \mathcal{X}_{N,p} \mathcal{X}_{N,p}^{T} = \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{T}$$

$$\mathcal{X}_{N,p} = \begin{pmatrix} X_{1,1} & X_{1,2} & \dots & X_{1,p} \\ \vdots & \vdots & & \vdots \\ X_{N,1} & X_{N,2} & \dots & X_{N,p} \end{pmatrix}$$
$$\mathbf{X}_{1} \qquad \mathbf{X}_{2} \qquad \dots \qquad \mathbf{X}_{p}$$

The empirical spectral measure of B_N is defined by

$$\mu_{\mathbf{B}_N} = \frac{1}{N} \sum_{k=1}^N \delta_{\lambda_k}$$

where $\lambda_1, \ldots, \lambda_N$ are the eigenvalues of **B**_N.

Introduction Sample covariance matrices

► Let
$$\mathbf{B}_N := \mathbf{B}_{N,p} = \frac{1}{p} \mathcal{X}_{N,p} \mathcal{X}_{N,p}^T = \frac{1}{p} \sum_{k=1}^p \mathbf{X}_k \mathbf{X}_k^T$$

$$\mathcal{X}_{N,p} = \begin{pmatrix} X_{1,1} & X_{1,2} & \dots & X_{1,p} \\ \vdots & \vdots & & \vdots \\ X_{N,1} & X_{N,2} & \dots & X_{N,p} \end{pmatrix}$$
$$\mathbf{X}_1 \qquad \mathbf{X}_2 \qquad \dots \qquad \mathbf{X}_p$$

The empirical spectral measure of B_N is defined by

$$\mu_{\mathbf{B}_N} = \frac{1}{N} \sum_{k=1}^N \delta_{\lambda_k}$$

where $\lambda_1, \ldots, \lambda_N$ are the eigenvalues of **B**_N.

▶ We shall suppose that $c_N := \frac{N}{\rho} \xrightarrow[N \to +\infty]{} c \in (0,\infty).$

Theorem (Marčenko, Pastur '67)

Let $(X_{ij})_{i,j \ge 1}$ be a family of *i.i.d.* random variables such that

 $\mathbb{E}(X_{11}) = 0$ and $Var(X_{11}) = \sigma^2$.

If $\lim_N N/p = c \in (0,\infty)$, then

 $\mu_{\mathbf{B}_{N}} \xrightarrow[N \to \infty]{w} \mu_{\mathrm{MP}}$ a.s.

Theorem (Marčenko, Pastur '67)

Let $(X_{ij})_{i,j \ge 1}$ be a family of *i.i.d.* random variables such that

$$\mathbb{E}(X_{11})=0$$
 and $\operatorname{Var}(X_{11})=\sigma^2.$

If $\lim_N N/p = c \in (0,\infty)$, then

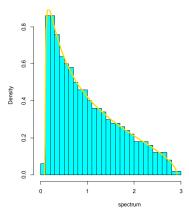
$$\mu_{\mathbf{B}_N} \xrightarrow[N \to \infty]{w} \mu_{\mathrm{MP}}$$
 a.s.

whose density is given by

$$\left(1-\frac{1}{c}\right)_+\delta_0+\frac{1}{2\pi c\sigma^2 x}\sqrt{(b-x)(x-a)}\mathbf{1}_{[a,b]}(x)dx$$

with $.+ := \max(0, .)$, $a = \sigma^2 (1 - \sqrt{c})^2$ and $b = \sigma^2 (1 + \sqrt{c})^2$.

Marčenko-Pastur Theorem



Wishart Matrix, N= 500 , n= 1000 , c= 0.5

Figure : In blue, histogram of the eigenvalues. In yellow, the density of the Marčenko-Pastur distribution

Sample covariance matrix

We consider

$$R_{N}^{1/2}\mathbf{B}_{N}R_{N}^{1/2} = \frac{1}{\rho}R_{N}^{1/2}\mathcal{X}_{N,\rho}\mathcal{X}_{N,\rho}^{T}R_{N}^{1/2}$$

where $R_N^{1/2}$ is a perturbation of the identity matrix:

 $R_N = \mathbf{I}_N + \theta \mathbf{u} \mathbf{u}^T, \quad \theta > 0$

Sample covariance matrix

We consider

$$R_{N}^{1/2}\mathbf{B}_{N}R_{N}^{1/2} = \frac{1}{\rho}R_{N}^{1/2}\mathcal{X}_{N,\rho}\mathcal{X}_{N,\rho}^{T}R_{N}^{1/2}$$

where $R_N^{1/2}$ is a perturbation of the identity matrix:

$$R_N = \mathbf{I}_N + \theta \mathbf{u} \mathbf{u}^T, \quad \theta > 0$$

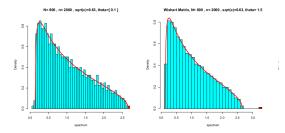


Figure : Histogram of the empirical eigenvalues and MP distribution

The Stieltjes transform $S_G : \mathbb{C}_+ \to \mathbb{C}$ of a measure ν on \mathbb{R} is defined by

$$S_{\nu}(z) := \int \frac{1}{x-z} \, d\nu(x)$$

The Stieltjes transform $S_G : \mathbb{C}_+ \to \mathbb{C}$ of a measure ν on \mathbb{R} is defined by

$$S_{\nu}(z) := \int \frac{1}{x-z} \, d\nu(x)$$

• $|S_{\nu}(z)| \leq 1/\Im \mathfrak{m}(z)$ and $\Im \mathfrak{m}(S_{\nu}(z)) \geq 0$

The Stieltjes transform $S_G : \mathbb{C}_+ \to \mathbb{C}$ of a measure ν on \mathbb{R} is defined by

$$S_{\nu}(z) := \int \frac{1}{x-z} \, d\nu(x)$$

• $|S_{\nu}(z)| \leq 1/\Im \mathfrak{m}(z)$ and $\Im \mathfrak{m}(S_{\nu}(z)) \geq 0$

• The function S_{ν} is analytic over \mathbb{C}_+ and characterizes ν

$$\nu([a,b]) = \lim_{y\downarrow 0} \frac{1}{\pi} \int_a^b \Im \mathfrak{m} S_\nu(x+iy) \, dx$$

The Stieltjes transform $S_G : \mathbb{C}_+ \to \mathbb{C}$ of a measure ν on \mathbb{R} is defined by

$$S_{\nu}(z) := \int \frac{1}{x-z} \, d\nu(x)$$

• $|S_{\nu}(z)| \leq 1/\Im \mathfrak{m}(z)$ and $\Im \mathfrak{m}(S_{\nu}(z)) \geq 0$

• The function S_{ν} is analytic over \mathbb{C}_+ and characterizes ν

$$\nu([a,b]) = \lim_{y\downarrow 0} \frac{1}{\pi} \int_a^b \Im \mathfrak{m} S_\nu(x+iy) \, dx$$

▶ For a sequence of measures $(\nu_n)_n$ on \mathbb{R} , we have

$$\left(\nu_n \xrightarrow[n \to \infty]{} \nu\right) \Leftrightarrow \left(\forall z \in \mathbb{C}_+, \ S_{\nu_n}(z) \xrightarrow[n \to \infty]{} S_{\nu}(z)\right).$$

The Stieltjes transform Characterization of the limiting distribution

To prove the M-P Theorem, it is equivalent to prove

$$orall z \in \mathbb{C}_+, \ S_{\mu_{\mathsf{B}_N}}(z) o S_{\mu_{MP}}(z)$$
 a.s.

where

$$S_{\mu_{MP}}(z)=rac{\sigma^2(1-c)-z+\sqrt{(z-b)(z-a)}}{2cz\sigma^2}$$

is the Stieltjes transform of the Marčenko-Pastur distribution

To prove the M-P Theorem, it is equivalent to prove

$$orall z \in \mathbb{C}_+, \ S_{\mu_{\mathsf{B}_N}}(z) o S_{\mu_{MP}}(z)$$
 a.s.

where

$$S_{\mu_{MP}}(z)=rac{\sigma^2(1-c)-z+\sqrt{(z-b)(z-a)}}{2cz\sigma^2}$$

is the Stieltjes transform of the Marčenko-Pastur distribution satisfying $\forall z \in \mathbb{C}_+$ the equation:

$$S_{\mu_{MP}}(z) = rac{1}{-z+1-c-czS_{\mu_{MP}}(z)}$$

having independent columns

Theorem (Bai-Zhou '08, Yao '12)

 $\mathbf{B}_{N} = \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{T} \text{ where the } \mathbf{X}_{k} \text{ 's are independent copies of } (X_{1}, \dots, X_{N})^{T} \text{ where } X_{i} = \sum \mathbf{a}_{j} \varepsilon_{i-j} .$

i≥0

1. the
$$\varepsilon_i$$
's are iid, centered and in L^4 .

2. $\sum_{j\geq 1} |a_j| < \infty$

having independent columns

Theorem (Bai-Zhou '08, Yao '12)

$$\begin{split} \mathbf{B}_{N} &= \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{T} \text{ where the } \mathbf{X}_{k} \text{ 's are independent copies of } \\ (X_{1}, \ldots, X_{N})^{T} \text{ where} \\ X_{i} &= \sum_{j \geq 0} \mathbf{a}_{j} \varepsilon_{i-j} \,. \end{split}$$
1. the ε_{i} 's are iid, centered and in \mathbf{L}^{4} .
2. $\sum_{i \geq 1} |\mathbf{a}_{i}| < \infty$

then, a.s. $\mu_{\mathbf{B}_{N}} \xrightarrow[N \to \infty]{w} \mu$ such that $S := S_{\mu}(z)$ verifies the equation

$$Z=-rac{1}{\underline{S}}+rac{c}{2\pi}\int_{0}^{2\pi}rac{1}{\underline{S}+\left(2\pi f(\lambda)
ight)^{-1}}d\lambda\,,$$

where $f(\cdot)$ is the spectral density of $(X_i)_{i \in \mathbb{Z}}$

having independent columns

Theorem (Bai-Zhou '08, Yao '12)

$$\begin{split} \mathbf{B}_{N} &= \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_{k} \mathbf{X}_{k}^{\mathsf{T}} \text{ where the } \mathbf{X}_{k} \text{ 's are independent copies of } \\ (X_{1}, \dots, X_{N})^{\mathsf{T}} \text{ where} \\ \mathbf{X}_{i} &= \sum \mathbf{a}_{j} \varepsilon_{i-j} \,. \end{split}$$

i≥0

- 1. the ε_i 's are iid, centered and in L^4 .
- 2. $\sum_{j\geq 1} |a_j| < \infty$ then, a.s. $\mu_{\mathbf{B}_N} \xrightarrow{\mathbf{w}} \mu$ such that $S := S_{\mu}(z)$ verifies the equation $z = -\frac{1}{\underline{S}} + \frac{c}{2\pi} \int_0^{2\pi} \frac{1}{S + (2\pi f(\lambda))^{-1}} d\lambda$,

where $f(\cdot)$ is the spectral density of $(X_i)_{i \in \mathbb{Z}}$ $f(x) = \frac{1}{2\pi} \sum_k Cov(X_0, X_k) e^{ixk}, x \in \mathbb{R}$

along both columns and rows

Let
$$\mathbf{B}_{N} = \frac{1}{p} \mathcal{X}_{N,p} \mathcal{X}_{N,p}^{T}$$

$$\mathcal{X}_{N,p} = \begin{pmatrix} X_{1,1} & X_{1,2} & \dots & X_{1,p} \\ \vdots & \vdots & & \vdots \\ X_{N,1} & X_{N,2} & \dots & X_{N,p} \end{pmatrix}$$

 Boutet de Monvel, Khorunzhy and Vasilchuk '96 Correlated Gaussian entries

along both columns and rows

Let
$$\mathbf{B}_{N} = \frac{1}{p} \mathcal{X}_{N,p} \mathcal{X}_{N,p}^{T}$$

$$\mathcal{X}_{N,p} = \begin{pmatrix} X_{1,1} & X_{1,2} & \dots & X_{1,p} \\ \vdots & \vdots & & \vdots \\ X_{N,1} & X_{N,2} & \dots & X_{N,p} \end{pmatrix}$$

- Boutet de Monvel, Khorunzhy and Vasilchuk '96 Correlated Gaussian entries
- Hachem, Loubaton and Najim '05

$$X_{i,j} = \sum_{(k,\ell)\in\mathbb{Z}^2} a_{k,\ell} G_{i-k,j-\ell}$$

Matrices whose entries are functions of iid random variables The Model

- Let $(\xi_{i,j})_{i,j\in\mathbb{Z}}$ be an array of iid random variables
- ▶ For all $(k, \ell) \in \mathbb{Z}^2$,

$$X_{k,\ell} := g(\xi_{k-i,\ell-j}; (i,j) \in \mathbb{Z}^2),$$

where $g: \mathbb{R}^{\mathbb{Z}^2} \to \mathbb{R}$ is a measurable function.

Matrices whose entries are functions of iid random variables The Model

- Let $(\xi_{i,j})_{i,j\in\mathbb{Z}}$ be an array of iid random variables
- ▶ For all $(k, \ell) \in \mathbb{Z}^2$,

$$X_{k,\ell} := g(\xi_{k-i,\ell-j}; (i,j) \in \mathbb{Z}^2),$$

where $g : \mathbb{R}^{\mathbb{Z}^2} \to \mathbb{R}$ is a measurable function. • $\mathbb{E}(X_{0,0}) = 0$ et $\mathbb{E}(X_{0,0}^2) < \infty$ Matrices whose entries are functions of iid random variables The Model

- Let $(\xi_{i,j})_{i,j\in\mathbb{Z}}$ be an array of iid random variables
- ▶ For all $(k, \ell) \in \mathbb{Z}^2$,

$$X_{k,\ell} := g(\xi_{k-i,\ell-j}; (i,j) \in \mathbb{Z}^2),$$

where $g: \mathbb{R}^{\mathbb{Z}^2}
ightarrow \mathbb{R}$ is a measurable function.

•
$$\mathbb{E}(X_{0,0}) = 0$$
 et $\mathbb{E}(X_{0,0}^2) < \infty$

• Let $\mathbf{B}_N = \frac{1}{p} \mathcal{X}_{N,p} \mathcal{X}_{N,p}^T = \frac{1}{p} \sum_{k=1}^{p} \mathbf{X}_k \mathbf{X}_k^T$

$$\mathcal{X}_{N,p} = \begin{pmatrix} X_{1,1} & X_{1,2} & \dots & X_{1,p} \\ \vdots & \vdots & & \vdots \\ X_{N,1} & X_{N,2} & \dots & X_{N,p} \end{pmatrix} \\ \mathbf{X}_{1} & \mathbf{X}_{2} & \dots & \mathbf{X}_{p} \end{pmatrix}$$

Let (G_{i,j})_{i,j∈ℤ} be an array of centered Gaussian random variables s.t. ∀ (i,j), (k, ℓ) ∈ ℤ²,

 $\mathbb{E}(G_{k,\ell}G_{i,j}) = \mathbb{E}(X_{k,\ell}X_{i,j})$

• Let $\mathbf{H}_N = \frac{1}{p} \mathcal{G}_{N,p} \mathcal{G}_{N,p}^T$

$$\mathcal{G}_{N,p} = \begin{pmatrix} G_{1,1} & G_{1,2} & \dots & G_{1,p} \\ \vdots & \vdots & & \vdots \\ G_{N,1} & G_{N,2} & \dots & G_{N,p} \end{pmatrix} \\ \mathbf{Z}_{1} & \mathbf{Z}_{2} & \dots & \mathbf{Z}_{p} \end{pmatrix}$$

Theorem (B., Merlevède and Peligrad '14)

Provided that $N, p \to \infty$ s.t. $N/p \to c \in (o, \infty)$, we have $\forall z \in \mathbb{C}_+$,

$$\lim_{N\to\infty} \left| S_{\mathbf{B}_N}(z) - \mathbb{E}(S_{\mathbf{H}_N}(z)) \right| = 0 \text{ a.s.}$$

Theorem (B., Merlevède and Peligrad '14)

Provided that $N, p \to \infty$ s.t. $N/p \to c \in (o, \infty)$, we have $\forall z \in \mathbb{C}_+$,

$$\lim_{N\to\infty} \left| S_{\mathsf{B}_N}(z) - \mathbb{E} \big(S_{\mathsf{H}_N}(z) \big) \right| = 0 \text{ a.s.}$$

Corollary

If $N, p \to \infty$ s.t. $N/p \to c \in (0, \infty)$ and if there exists μ such that for any continuous and bounded function $f : \mathbb{R} \to \mathbb{R}$,

$$\mathbb{E}\int f d\mu_{\mathbf{H}_N} \xrightarrow[N \to +\infty]{} \int f d\mu,$$

then

$$\mu_{\mathbf{B}_N} \xrightarrow[N \to \infty]{w} \mu \quad a.s.$$

Theorem (B., Merlevède and Peligrad '14)

Let $N, p \to \infty$ t.q. $N/p \to c \in (0, \infty)$. Assume that

 $\sum_{k,\ell\in\mathbb{Z}} |\mathrm{Cov}(X_{0,0},X_{k,\ell})| < \infty$

Then,

$$\mu_{\mathbf{B}_N} \xrightarrow[N \to \infty]{w} \mu \quad a.s.$$

whose Stieltjes transform $S := S_{\mu}(z)$ verifies: $\forall z \in \mathbb{C}^+$ $S(z) = \int_0^1 h(x, z) dx$

where h(x, z) is a solution of the equation

$$h(x,z) = \left(-z + \int_0^1 \frac{f(x,s)}{1+c \int_0^1 f(u,s)h(u,z)du} ds\right)^{-1}$$

avec

$$f(x,y) = \sum_{k,\ell \in \mathbb{Z}} \operatorname{Cov}(X_{0,0}, X_{k,\ell}) e^{-2\pi i (kx+\ell y)}$$

Corollary (B., Merlevède, Peligrad '14)

Let $(a_{i,j})_{(i,j)\in\mathbb{Z}^2}$ be a double indexed sequence of numbers such that $\sum_{i,j\in\mathbb{Z}} |a_{i,j}| < \infty$. Let

$$X_{k,\ell} = \sum_{i,j\in\mathbb{Z}} a_{i,j}\xi_{k+i,\ell+j}.$$

The result follows with

$$f(x,y) = \mathbb{E}(\xi_{0,0}^2) \sum_{k,\ell \in \mathbb{Z}} \sum_{i,j \in \mathbb{Z}} a_{i,j} a_{k+i,\ell+j} e^{-2\pi(kx+\ell y)}$$

Corollary (B., Merlevède, Peligrad '14)

Let $(a_{i,j})_{(i,j)\in\mathbb{Z}^2}$ be a double indexed sequence of numbers such that $\sum_{i,j\in\mathbb{Z}} |a_{i,j}| < \infty$. Let

$$X_{k,\ell} = \sum_{i,j\in\mathbb{Z}} a_{i,j}\xi_{k+i,\ell+j}.$$

The result follows with

$$f(x,y) = \mathbb{E}(\xi_{0,0}^2) \sum_{k,\ell \in \mathbb{Z}} \sum_{i,j \in \mathbb{Z}} a_{i,j} a_{k+i,\ell+j} e^{-2\pi(kx+\ell y)}$$

We generalize the results of Hachem et al. '05, Yao '12 and Pan et al. '13.

Other Applications

Other possible applications:

- functions of linear processes,
- Volterra type processes,
- ► ARCH models, ... etc

Thank you for your attention!